# Resampled Branching Particle Filters[∗](#page-0-0)

Michael A. Kouritzin ,

Department of Mathematical and Statistical Sciences University of Alberta, Edmonton, AB T6G 2G1 Canada e-mail: [michaelk@ualberta.ca](mailto:michaelk@ualberta.ca)

Abstract: A large class of discrete-time branching particle filters with Bayesian model selection capabilities and effective resampling is introduced in algorithmic form, shown empirically to outperform the popular bootstrap algorithm and analyzed mathematically. The particles interact weakly in the resampling procedure. The weighted particle filter, which has no resampling, and the fully-resampled branching particle filter are included in the class as extreme points. Each particle filter in the class is coupled to a McKean-Vlasov particle system, corresponding to a reduced, unimplementable particle filter, for which Marcinkiewicz strong laws of large numbers (Mllns) and the central limit theorem (clt) can be written down. Coupling arguments are used to show the reduced system can be used to predict performance of the particle filter and to transfer the Mllns to the original weakly-interacting particle filter. This clt is also shown transferable when extra particles are used.

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#### 1. Introduction

Nonlinear filtering deals with determining the distribution of the current state of a non-observable, random, dynamic signal X given the history of a distorted, corrupted partial observation process Y living on the same probability space  $(\Omega, \mathcal{F}, P)$  as X. Bayesian model selection, sometimes done while filtering, deals with determining which of a class of signal models  $\{X^{(i)}\}_{i\in I}$  best fits the observed values of Y by pairwise Bayes' factor comparison. For many practical problems each potential signal is a time-homogeneous discrete-time Markov process  $\{X_n, n = 0, 1, 2, ...\}$ , living on some complete, separable metric space  $(E, \rho)$ , with initial distribution  $\pi_0$  and transition probability kernel K. The observation process takes the form  $(Y_0 = 0 \text{ and})$  $Y_n = h(X_{n-1}) + V_n$  for  $n \in \mathbb{N}$ , where  ${V_n}_{n=1}^{\infty}$  are independent random vectors with common strictly positive, bounded density g that are independent of  $X$ , and the sensor function h is a measurable mapping from E to  $\mathbb{R}^d$ . Then, the objective of filtering (with respect to any given signal model X) is to compute the conditional probabilities  $\pi_n(A) = P(X_n \in A | \mathcal{F}_n^Y)$ ,  $n = 1, 2, ...,$  for all Borel sets A or, equivalently, the conditional expectations  $\pi_n(f) = E^P(f(X_n) | \mathcal{F}_n^Y)$  for all bounded, measurable functions  $f: E \to \mathbb{R}$ , where  $\mathcal{F}_n^Y \triangleq \sigma\{Y_l, l = 1, ..., n\}$  is the information obtained from the back observations. On the other hand, the objective of Bayes' factor model selection is to compare the ratio  $B_n^{12} = \frac{E^Q[L_n(Y|X^{(1)})|\mathcal{F}_n^Y]}{E^Q[L_n(Y|X^{(2)})|\mathcal{F}_n^Y]}$  $\frac{E^{\bullet}[L_n(Y|X^{(2)})|\mathcal{F}_n]}{E^Q[L_n(Y|X^{(2)})|\mathcal{F}_n^Y]}$  of marginal likelihoods between potential signal models  $X^{(1)}$  and  $X^{(2)}$  with respect to some reference probability measure Q. (The metric space E, initial distribution  $\pi_0$ , transition probability K and sensor function h can all depend upon the potential sigmal  $X^{(i)}$  as long as the observation noise  $\{V_n\}$  is the same.)

<span id="page-0-0"></span>Suppose without loss of generality that  $\Omega = (E \times \mathbb{R}^d)^{\infty}$  and  $\mathcal{F} = \mathcal{B}((E \times \mathbb{R}^d)^{\infty})$  until later extended. Moreover, suppose hereafter  $\mathcal{F}_{-1}^{\xi} \triangleq \{\emptyset, \Omega\}$ ,  $\mathcal{F}_{n}^{\xi} \triangleq \sigma\{\xi_{l}^{k}, k \in \mathcal{K}, l \leq n\}$  when  $n \in \mathbb{N}_{0}$  and  $\mathcal{F}_{\infty}^{\xi} \triangleq \sigma\{\xi_{l}^{k}, k \in \mathcal{K}, l \leq n\}$ 

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 $\mathcal{K}, l < \infty$ } for random variables  $\{\xi_n^k, k \in \mathcal{K}, n \in \{0, 1, ...\}\}$  on  $(\Omega, \mathcal{F})$ . (This is consistent with  $\mathcal{F}_n^Y$  defined above if  $K$  has one element.) Unnormalized filters transfer the information contained in the observations to a likelihood process by measure change. In this method, a reference probability measure  $Q$  is introduced under which the signal, observation process  $\{(X_n, Y_{n+1}), n = 0, 1, ...\}$  has the same distribution as the signal, noise process  $\{(X_n, V_{n+1}), n = 0, 1, ...\}$  does under P. Hence, the observations are i.i.d. random vectors with strictly positive bounded density  $g$  and are independent of  $X$  under measure  $Q$ . All the observation information is absorbed into the likelihood process  $\{L_n, n = 1, 2, ...\}$  transforming Q back to P, which in our case has the form

$$
\frac{dP}{dQ}\Big|_{\mathcal{F}_{\infty}^X \vee \mathcal{F}_n^Y} = L_n = \prod_{j=1}^n \alpha_j(X_{j-1}), \text{ with } \alpha_j(x) = \frac{g(Y_j - h(x))}{g(Y_j)},\tag{1.1}
$$

<span id="page-1-0"></span>so  $L_n = \alpha_n(X_{n-1})L_{n-1}$  and  $L_0 = 1$ . The following (well-known) discrete Girsanov's theorem constructs the real probability  $P$  from the fictitious one  $Q$ .

**Theorem 1.1.** Suppose under probability Q that  $\{X_n, n = 0, 1, ...\}$  and  $\{Y_n, n = 1, 2, ...\}$  are independent processes with on  $(\Omega, \mathcal{F})$ , the  $\{Y_n\}$  are i.i.d. with strictly-positive, bounded density g on  $\mathbb{R}^d$  and  $V_n \cong$  $Y_n - h(X_{n-1})$  for all  $n = 1, 2, ...$  Then, there exists a probability measure P such that [\(1.1\)](#page-1-0) holds,  $\{V_n, n = 1\}$ 1, 2, ...} are i.i.d. on  $(\Omega, \mathcal{F}, P)$  with density g and  $\{X_n\}$  is independent of  $\{V_n\}$  with the same law as on  $(\Omega, \mathcal{F}, Q)$ .

Filtering and model selection can be done simultaneously by using the *unnormalized filters* 

<span id="page-1-1"></span>
$$
\sigma_n(f) = E^Q \left( L_n f \left( X_n \right) \big| \mathcal{F}_n^Y \right) \tag{1.2}
$$

so  $\sigma_0 = \pi_0$ , as  $L_0 = 1$  and  $\mathcal{F}_0^Y = \{\emptyset, \Omega\}$ . Then,  $\pi_n(f) = \frac{\sigma_n(f)}{\sigma_n(1)}$  by Bayes rule and  $B_n^{12} = \frac{\sigma_n^{(1)}(1)}{\sigma_n^{(2)}(1)}$  $\frac{\sigma_n^{\gamma}(1)}{\sigma_n^{(2)}(1)},$  where  $\sigma_n^{(i)}(f) = E^Q\left($  $L_n^{(i)} f\left(X_n^{(i)}\right) \bigg| \mathcal{F}^Y_n$  $\setminus$ with  $L_n^{(i)} = \prod^n$  $\prod_{j=1}^{n} \alpha_j(X_{j-1}^{(i)})$  is the unnormalized filter for signal model

 $X^{(i)}$ . Therefore, we can combine Bayesian model selection and filtering (for each potential signal) by constructing approximations (denoted  $S_n^N$  and  $\mathbb{S}_n^N$  below) to the unnormalized filter for each candidate signal model as done in Kouritzin and Zeng [\[17](#page-51-0)],[\[18\]](#page-51-1) and Kouritzin [\[15\]](#page-51-2).

Nowadays, particle filters are utilized in a wide variety of applications in as diverse areas as econometrics, defense and clickstream analysis. The original (resampled) interacting particle filters have been intensely studied (see e.g. Cappe, Godsill and Moulines [\[3\]](#page-50-0) for an overview and historical account). However, Del Moral, Kouritzin and Miclo [\[7\]](#page-50-1) show that the performance of a particle filter depends heavily upon the resampling used and little theory is known about optimal resampling. Furthermore, these particle filters approximate the actual filter  $\pi_n$  and hence are not amenable to model selection (without storing prior filter estimates). On the other hand, the weighted particle filter (largely credited to Handschin [\[12](#page-50-2)] as well as Handschin and Mayne [\[13\]](#page-51-3) and studied in Kurtz and Xiong [\[19\]](#page-51-4), [\[20\]](#page-51-5)) approximates the unnormalized particle filter  $\sigma_n$ , is the most basic particle filter with model selection capabilities and is embarassingly computer parallelizable. More generally, branching particle filters, like those introduced by Crisan and Lyons [\[5](#page-50-3)], can have model selection capabilities, effective resampling and be highly parallelizable. Herein, we introduce and analyze a class of exchangeable branching particle filters with resampling that avoids the weighted-particle-filter particle spread problems yet still has model selection capabilities. It includes the weighted particle filter as the extreme zero-resampling case and a model selection variation of the better algorithm in Del Moral, Kouritzin and Miclo [\[7\]](#page-50-1) as the extreme fully-resampled case.

One might be tempted to think that the asymptotic properties of practical particle filters are already known. However, this is untrue: Most of the analysis (see e.g. Del Moral and Miclo [\[8](#page-50-4)]) has been done on a somewhat-impractical particle filter that resamples every particle at every observation leading to poor performance as noted in Del Moral, Kouritzin and Miclo [\[7\]](#page-50-1). To combat this over-resampling problem, many researchers brought in importance sampling and delayed bulk resampling methods (see e.g. Del Moral, Doucet and Jasra [\[6](#page-50-5)]) and studied the resulting algorithms. However, there are fewer studies of the seemingly-practical partially-resampled algorithms (see e.g. Ballantyne, Chan and Kouritzin [\[1\]](#page-50-6)), where decisions are made at a particle-by-particle basis with the aim of only removing the poor particles and splitting the best particles (in an unbiased manner). These algorithms are apparently difficult to analyse. However, Del Moral, Kouritzin and Miclo [\[7](#page-50-1)] studied such an algorithm and established a historical strong law of large number and a central limit theorem. Their results show superior performance over prior algorithms but rely on a strong bounded-and-strictly-positive assumption on a 'Feynman-Kac' process  $\{U_n, n \in \mathbb{N}\},$  which limits use in non-linear filtering. Separately, Kouritzin and Sun [\[16\]](#page-51-6) obtain  $L_2$ -rates of convergence for a partially-resampled algorithm. However, no other results were attained and their results are in a specific setting. Our present work sets up a framework for studying resampled branching particle filters.

Our algorithm is given in the next section and compared with the most famous and popular particle filter, the *bootstrap algorithm*. In particular, it is illustrated that our algorithm is faster and more accurate at tracking than the bootstrap algorithm. This comparison is followed by our mathematical notation in Section 3. To state our results, we let  $\mathbb{S}_n^N(f)$  be our branching particle approximation to the unnormalized filter  $\sigma_n(f)$ . Our main result, Theorem [5.1](#page-21-0) in Section 5, states that, for almost all observation paths,  $\mathbb{S}_n^N(f)$ satisfies the Mllns (with all possible rates) and the normalized difference  $\sqrt{N}(\mathbb{S}_n^N(f) - \sigma_n(f))$  satisfies the clt (with variance characterized by the resampling employed). Taken together these results say the same polynomial rates of almost-sure convergence in number of particles N hold for our resampled branching particle filters as for other particle filters (like the weighted) even when no extra particles are used. Moreover, under the extra particle condition  $\frac{N}{m_N} \to 0$ , the random weak particle interactions in our algorithm average out enough to characterize the optimal convergence with a clt. To obtain these results, we couple our algorithm to a reduced particle system, introduced in Section 4, which is unimplementable but mathematically simpler. Conceptually, our partially-resampled particle filter is a weakly-interacting particle system and the reduced system is a more-tractable McKean-Vlasov-type limit (with average weight  $\mathbb{A}_n$ replaced by  $\sigma_n(1)$ , which can be used to predict performance of the partially-resampled particle filter. We also introduce tracking systems in Section 6, which run as weighted filters but indicate where the resampled and reduced filters would resample (at least initially). These tracking systems are introduced for purely analytical reasons to help us divide the resampled and weighted particle filters into comparable pieces. They also have to be coupled to the resampled and reduced particle filters. The actual coupling and its ramifications are contained in Section 7. The first appendix contains the derivation of the clt variance for the McKean-Vlasov and partially-resampled filter. The second appendix contains a technical total-mass ergodic theorem for the partially-resampled filter using the coupling.

# 2. New Algorithm and Numeric Comparison

# 2.1. Bootstrap Algorithm

The bootstrap particle filter algorithm was introduced in 1993 by Gordon, Salmond and Smith [\[11\]](#page-50-7). It is one of the big breakthroughs in big data sequential estimation and its convergence properties have been

thoroughly studied in e.g. Del Moral and Miclo [\[8](#page-50-4)]. It overcomes the increasing variance weight problem of the weighted filter pointed out in Doucet, Godsill and Andrieu [\[10](#page-50-8)]. However, it has its limitations in terms of model selection, parallelizability, performance and speed. For clarity, we first summarize the bootstrap algorithm:

**Initialize:**  $\left\{\mathbb{X}_{0}^{k}\right\}_{k=1}^{N}$  are independent initial particle samples of  $\pi_{0}$ ,  $V_{N+1} = 1$ 

## **Repeat:** for  $n = 0, 1, 2, \dots$  do

- 1. Weight by Observation:  $\widehat{\mathbb{L}}_n^k = \alpha_{n+1} (\mathbb{X}_n^k)$  for  $k = 1, 2, ..., N$
- 2. Normalize Weight:  $w_{n+1}^k = \frac{\widehat{\mathbb{L}}_n^k}{\widehat{\mathbb{L}}_n}$  for  $k = 1, 2, ..., N$ , where  $\widehat{\mathbb{L}}_n = \sum_{i=1}^N$  $\sum_{i=1}^{\infty} \widehat{\mathbb{L}}_n^i$
- 3. Evolve Independently:

$$
P^{Y}(\widehat{\mathbb{X}}_{n+1}^{k} \in \Gamma_{k} \ \forall \ k | \mathcal{F}_{n}^{\mathbb{X}}) = \prod_{k=1}^{N} K(\mathbb{X}_{n}^{k}, \Gamma_{k}) \text{ for all } \Gamma_{k}
$$

- 4. Estimate  $\pi_{n+1}$  by:  $\mathbb{P}_{n+1}^{N} = \sum^{N}$  $k=1$  $w_{n+1}^k \delta_{\widehat{\mathbb{X}}^k_{n+1}}.$
- 5. Resample:  $p_i = \sum_{k=1}^i w_{n+1}^k$  for  $i = 1, ..., N, j = N 1$

**Repeat:** for  $k = N, N - 1, ..., 2, 1$  do

- Draw [0, 1]-uniform  $U_k$  and set  $V_k = U_k^{\frac{1}{k}} V_{k+1}$
- While  $V_k \leq p_i$  set  $j = j 1$
- Set  $\mathbb{X}_{n+1}^k \triangleq \widehat{\mathbb{X}}_{n+1}^{j+1}$

Remark 2.1. We extract our estimate before resampling to avoid excess noise. Our actual code represents our attempt to make the algorithm as efficient as reasonably possible and is available upon request.

This algorithm is  $O(N)$  in terms of operations per particle. In particular, we utilized a clever idea credited to Carpenter, Clifford and Fearnhead [\[4](#page-50-9)] to keep the resampling to  $O(N)$ .  $(V_1, ..., V_N)$  has the joint distribution of the order statistics for  $\{U_k\}_{k=1}^N$ .

There are variations to the evolution and resampling steps (see e.g. DelMoral, Kouritzin and Miclo [\[7](#page-50-1)]; Douc, Cappé and Moulines  $[9]$ ) that can be better in certain instances. But rather than considering the speed-performance tradeoffs of these variations, we chose to present the standard algorithm in the best possible light. Doucet, Godsill and Andrieu [\[10\]](#page-50-8) give another variation that alternates between the weighted and bootstrap algorithms depending upon how many effective particles there are. This later variant definitely shows the tradeoff between introducing resampling noise into the system and coping with continual weight variance increase. However, we argue that it is better for performance to make the resampling decisions on a particle-by-particle basis. This would also avoid the two separate time problem: a fast time when there is no resampling and a slow one when there is. Some real-time applications are not conducive to sudden switches to slow times.

#### 2.2. New Branching Filter - Implementable Version

We define the following branching Markov process  $\{S_n^N, n = 0, 1, ...\}$  approximation to unnormalized filter  $\{\sigma_n, n = 0, 1, ...\}$  in terms of the observations as follows:

**Initialize:**  $\left\{\mathbb{X}_{0}^{k}\right\}_{k=1}^{N}$  are independent samples of  $\pi_{0}$ ,  $\mathbb{N}_{0} = N$ ,  $\mathbb{N}_{n} = 0$  for all  $n \in \mathbb{N}$  and  $\mathbb{L}_{0}^{k} = 1$  for  $k = 1, ..., N$ .

**Repeat:** for  $n = 0, 1, 2, ...$  do

- 1. Weight by Observation:  $\widehat{\mathbb{L}}_{n+1}^k = \alpha_{n+1} \left( \mathbb{X}_n^k \right) \mathbb{L}_n^k$  for  $k = 1, 2, ..., N_n$
- 2. Evolve Independently:

$$
P^{Y}(\widehat{\mathbb{X}}_{n+1}^{k} \in \Gamma_{k} \ \forall \ k | \mathcal{F}_{n}^{\mathbb{X}} \vee \mathcal{F}_{\infty}^{\mathbb{U}}) = \prod_{k=1}^{\mathbb{N}_{n}} K(\mathbb{X}_{n}^{k}, \Gamma_{k}) \ \forall \Gamma_{k}
$$

- 3. Estimate  $\sigma_{n+1}$  by:  $\mathbb{S}_{n+1}^N = \frac{1}{N}$ N  $\sum_{n}$  $k=1$  $\widehat{\mathbb{L}}_{n+1}^k \delta_{\widehat{\mathbb{X}}^k_{n+1}}.$ 4. Average Weight:  $\mathbb{A}_{n+1} = \mathbb{S}_{n+1}^N(1)$
- **Repeat (5-6):** for  $k = 1, 2, ..., N_n$  do
- 5. Resampled Case: If  $\widehat{\mathbb{L}}_{n+1}^k \notin (a_n \mathbb{A}_{n+1}, b_n \mathbb{A}_{n+1})$  then
	- (a) Offspring Number:  $\mathbb{N}^k_{n+1}$  =  $\frac{\widehat{\mathbb{L}}_{n+1}^k}{\Lambda}$  $\frac{\widehat{\mathbb{L}}_{n+1}^k}{\mathbb{A}_{n+1}}$  +  $\rho_n^k$ , with  $\rho_n^k$  a  $\begin{cases} \widehat{\mathbb{L}}_{n+1}^k = \\ \widehat{\mathbb{A}}_{n+1} \end{cases}$  $\frac{\widehat{\mathbb{L}}_{n+1}^k}{\Lambda}$  $\left(\frac{\widehat{\mathbb{L}}_{n+1}^k}{\mathbb{A}_{n+1}}\right)$ -Bernoulli independent of everything
	- (b) Resample:  $\mathbb{L}_{n+1}^{\mathbb{N}_{n+1}+j} = \mathbb{A}_{n+1}, \mathbb{X}_{n+1}^{\mathbb{N}_{n+1}+j} = \widehat{\mathbb{X}}_{n+1}^k$  for  $j = 1, ..., \mathbb{N}_{n+1}^k$
	- (c) Add Offspring Number:  $\mathbb{N}_{n+1} = \mathbb{N}_{n+1} + \mathbb{N}_{n+1}^k$
- 6. Non-resample Case: If  $\widehat{\mathbb{L}}_{n+1}^k \in (a_n \mathbb{A}_{n+1}, b_n \mathbb{A}_{n+1})$  then  $\mathbb{N}_{n+1} = \mathbb{N}_{n+1} + 1, \, \mathbb{L}_{n+1}^{\mathbb{N}_{n+1}} = \widehat{\mathbb{L}}_{n+1}^k, \, \mathbb{X}_{n+1}^{\mathbb{N}_{n+1}} = \widehat{\mathbb{X}}_{n+1}^k$

Remark 2.2. We extract our estimate before resampling to avoid excess noise. Our actual code represents our attempt to make the algorithm as efficient as reasonably possible and is available upon request.

After establishing the appropriate bounds on  $\mathbb{N}_{n+1}^k$  in the sequel, we can easily see that this algorithm is also  $O(N)$ . Indeed, a careful comparison of this algorithm to the prior bootstrap one leads us to the believe that the constant implied in the  $O(N)$  notation for the branching algorithm may be smaller than that for the bootstrap, especially when the Resampled Case does not occur too often. We will establish this fact experimentally below. Since  $\sigma_n$  is estimated both model selection and filtering can be done simultaneously.

We are not the first to use branching particle filters for tracking. Indeed, we were inspired by the works of Crisan and Lyons [\[5](#page-50-3)] and Ballantyne, Chan and Kouritzin [\[1](#page-50-6)]. However, our algorithm differs from the ones presented in those papers and our goals are also different.

#### 2.3. Numeric Comparison

The bootstrap algorithm has two inherent disadvantages over our branching particle filter:

- 1. Model selection is not readily available. Since the bootstrap algorithm works with the (normalized) filter versus the unnormalized filter model selection via filtering is not immediately available. Rather one must first convert to the unnormalized filters, which requires storing prior filter values as well as extra computations.
- 2. Effective parallelization is difficult (see e.g. [\[21](#page-51-7)]).

These are reason enough to choose the branching algorithm over the bootstrap. However, we can also consider both algorithms on a single-processor implementation, purely tracking problem (which is what the bootstrap was designed for) and focus on two questions:

- 1. Does our branching filter perform better and require less computation time than the bootstrap algorithm, even for tracking?
- 2. Within the branching algorithm, how much resampling should we use? In other words, how should  $a_n$  and  $b_n$  be chosen for best performance?

We applied bootstrap algorithm and branching algorithm to the following simple model

$$
X_n = 0.95X_{n-1} + 0.3W_n
$$
  
\n
$$
Y_n = X_{n-1} + V_n,
$$

where  $X_0$ ,  $W_n$  and  $V_n$  are independent with standard Cauchy distribution.

For simplicity, we define  $(a_n, b_n) = (1/r, r)$ , where  $r \in [1, \infty]$ , and refer to r as the resampling parameter. All particles will resample when  $r = 1$ , which we call complete resampling. No particle will resample when  $r = \infty$ , which means we have the weighted particle filter.

The results for computation time are showed in Table [1.](#page-5-0) It shows that branching is significantly faster

<span id="page-5-0"></span>

| Number Particles N   | 100   | 400   | 2000  | 10000 | 50000 |  |  |  |
|----------------------|-------|-------|-------|-------|-------|--|--|--|
| Bootstrap            | 0.004 | 0.017 | 0.078 | 0.467 | 2.513 |  |  |  |
| Branching $(r=1)$    | 0.003 | 0.016 | 0.066 | 0.413 | 1.913 |  |  |  |
| Branching $(r=2.25)$ | 0.003 | 0.015 | 0.063 | 0.343 | 1.612 |  |  |  |
| Weighted             | 0.001 | 0.006 | 0.032 | 0.231 | 1.133 |  |  |  |
| TABLE                |       |       |       |       |       |  |  |  |

Execution Time in Seconds

than the bootstrap algorithm. Moreover, the branching algorithm becomes faster as one resamples less i.e. r increases. It is worth noting that there is no explicit particle control in the branching algorithms. However, there is a mild implicit control built in. As a consequence, the times and performance are reasonably consistent from sample path to sample path.

To compare the tracking performance of the branching filter with different r's and the bootstrap filter, we define the residual as

$$
residual = \sqrt{\frac{1}{n} \sum_{k=1}^{n} (\pi_k^N(f) - f(X_k))^2},
$$

where  $\pi_k^N(f)$  is the normalized filter approximation at time instant k and N initial particles,  $X_k$  is the real signal and  $f$  is a bounded function defined as

$$
f(x) = \begin{cases} 30 & : x > 30 \\ x & : -30 \le x \le 30 \\ -30 & : x < -30 \end{cases}
$$



<span id="page-6-0"></span>Fig 1: Average Residual versus Resampling Amount

(For the branching filters, we obtain  $\pi_k^N(f)$  by dividing  $\sigma_k^N(f)$  by  $\sigma_k^N(1)$ .) We let r range from 1 to 10 in increments of 0.25, use N initial particles and  $n = 50$  time steps. The comparison between the bootstrap and branching algorithm is presented in Figure [1](#page-6-0) for 3000 runs with  $N = 400$  initial particles. It shows the residual for the bootstrap, which does not depend upon  $r$ , and branching particle filter with different r. From the graph, it is obvious that the residual for bootstrap is far greater than that for branching. Next, we remove the bootstrap from this figure to determine the best amount of resampling within the branching algorithm. The best branching particle filter should not only be able to kill the bad particles but also avoid excess resampling which introduces more noise. The best  $r$  for resampling is around 2.25 for this problem with  $N = 400$  as shown in Figure [2.](#page-7-0) Table [2](#page-9-0) summarizes the performance and establishes that all branching particle filters are basically at the optimal filter by  $N = 50,000$  particles. However, the bootstrap is still a way off. In fact, it is quite amazing that the performance of the bootstrap with 10, 000 particles is significantly worse than the branching filters with only 400 initial particles for this problem.

Next, we show how the residual changes versus time on two randomly chosen outcomes. The results for bootstrap and branching with  $r = 2.25$  are presented in Figures [3](#page-8-0) and [4.](#page-8-1) At the beginning, the two algorithms produce the same estimate because they have the same initial particles. Clearly, the bootstrap deviates further and longer from the signal than the branching does. This is probably due to too many particles resampling to a favorite site that happens (due to randomness) to be quite wrong. The branching



<span id="page-7-0"></span>Fig 2: Average Residual versus Resampling Amount



<span id="page-8-0"></span>Fig 3: Typical Residual versus Time - Outcome 1



<span id="page-8-1"></span>Fig 4: Typical Residual versus Time - Outcome 2

| Number Particles N   | 100      | 400     | 2000    | 10000    | 50000   |  |  |  |
|----------------------|----------|---------|---------|----------|---------|--|--|--|
| Bootstrap            | 8.53602  | 7.87591 | 6.99739 | 5.32543  | 4.91826 |  |  |  |
| Branching $(r=1)$    | 5.49346  | 4.96183 | 4.50097 | 4.504418 | 4.49376 |  |  |  |
| Branching $(r=2.25)$ | 5.428414 | 4.91768 | 4.64787 | 4.539548 | 4.4812  |  |  |  |
| Weighted             | 5.789023 | 5.13474 | 4.74356 | 4.68098  | 4.52188 |  |  |  |
| TABLE 2              |          |         |         |          |         |  |  |  |

<span id="page-9-0"></span>Average Residual

algorithm is much more cautious, especially for  $r > 1$ . The deviations in these figures also help explain why the overall residual for bootstrap is much greater than branching as shown in Figure [1.](#page-6-0)

## 2.4. New Branching Filter - Analyzable Version

To analyze our branching particle filter, we re-introduce it in new notation, using two initial particle types:  $N \in \mathbb{N}$  filter particles and  $m_N - N \in \mathbb{N}$  extra particles. The purpose of the extra particles is to allow enough asymptotic independence for the central limit theorem (clt) to hold. (They are not necessary for the Mllns to hold.) We define the following branching Markov process  $\{\mathbb{S}_n^N, n = 0, 1, ...\}$  approximation to  $\{\sigma_n, n = 0, 1, ...\}$  in terms of the observations as follows:

 $\left\{\mathbb{X}^{k,1}_0\right\}_{k=1}^{m_N}$  $\sum_{k=1}^{m_N}$  are independent (initial particle) samples from  $\pi_0$ ,  $\left\{\mathbb{V}_n^{k,i}\right\}_{n,i,k=1}^{\infty,\infty,m_N}$  are zero-mean i.i.d. random variables, and  $\mathbb{N}_0^k = 1$ ,  $\mathbb{L}_0^{k,1} = 1$  for  $k = 1, ..., m_N$ . To handle possible degeneracy, we also preset  $\mathbb{N}_n^{k,i} = 0$  for all  $i, k, n \in \mathbb{N}$ .

#### **Repeat:** for  $n = 0, 1, 2, ...$  do

- 1. Weight by Observation:  $\widehat{\mathbb{L}}_n^{k,i} = \alpha_{n+1} \left( \mathbb{X}_n^{k,i} \right) \mathbb{L}_n^{k,i}$  for  $i = 1, 2, ..., N_n^k$ ,  $k = 1, 2, ..., m_N$
- 2. Average Weight:  $\mathbb{A}_{n+1} = \frac{1}{m_N} \sum_{i=1}^{m_N}$  $k=1$  $\sum_{n=1}^{\mathbb{N}_{n}^{k}}$  $\sum_{i=1}^n \widehat{\mathbb{L}}_n^{k,i}$

**Repeat (3-5):** for  $k = 1, 2, ..., m_N$  do

**Repeat (3-5):** for  $i = 1, 2, ..., N_n^k$  do

- 3. Resampled Case: If  $\hat{\mathbb{L}}_n^{k,i} + \mathbb{V}_{n+1}^{k,i} \notin (a_n \mathbb{A}_{n+1}, b_n \mathbb{A}_{n+1})$  then
	- (a) Offspring Number:  $\mathbb{N}_{n+1}^{k,i}$  =  $\left| \frac{\widehat{\mathbb{L}}^{k,i}_{n}}{\mathbb{A}_{n+1}} \right| + \rho^{k,i}_{n},$  with  $\rho^{k,i}_{n}$  a  $\begin{cases} \frac{\widehat{\mathbb{L}}^{k,i}_n}{\mathbb{A}_{n+1}} \, - \end{cases}$  $\left\lfloor \frac{\widehat{\mathbb{L}}^{k,i}_n}{\mathbb{A}_{n+1}} \right\rfloor$ -Bernoulli independent of everything
	- (b) Resampled Weight:  $\overline{\mathbb{L}}_n^{k,i} = \mathbb{A}_{n+1}$
- 4. Non-resample Case: If  $\widehat{\mathbb{L}}_n^{k,i} + \mathbb{V}_{n+1}^{k,i} \in (a_n \mathbb{A}_{n+1}, b_n \mathbb{A}_{n+1})$  then  $\overline{\mathbb{L}}_n^{k,i} = \widehat{\mathbb{L}}_n^{k,i}, \, \mathbb{N}_{n+1}^{k,i} = 1$ 5. Combine:  $\widehat{\mathbb{X}}_n^{k,j} \triangleq \mathbb{X}_n^{k,i}, \mathbb{L}_{n+1}^{k,j} \triangleq \overline{\mathbb{L}}_n^{k,i}$  for

$$
j \! \in \! \left\{ \mathbb{N}_{n+1}^{k,1} + \cdots + \mathbb{N}_{n+1}^{k,i-1} + 1, ..., \mathbb{N}_{n+1}^{k,1} + \cdots + \mathbb{N}_{n+1}^{k,i} \right\}
$$

6. Evolve Independently:

$$
P^Y(\mathbb{X}_{n+1}^{k,j}\in\Gamma_{k,j}\,\,\forall\,\,k,j|\mathcal{F}_n^{\mathbb{X}}\vee\mathcal{F}_{\infty}^{\mathbb{U},\mathbb{V}})=\prod_{k=1}^{m_N}\prod_{j=1}^{\mathbb{N}_{n+1}^{k}}K(\widehat{\mathbb{X}}_n^{k,j},\Gamma_{k,j})
$$

.

for all  $\Gamma_{k,j}$ , where  $\mathbb{N}_{n+1}^k = \mathbb{N}_{n+1}^{k,1} + \cdots + \mathbb{N}_{n+1}^{k,\mathbb{N}_n^k}$ 

7. Estimate 
$$
\sigma_{n+1}
$$
 by:  $\mathbb{S}_{n+1}^N = \frac{1}{N} \sum_{k=1}^N \sum_{j=1}^{\mathbb{N}_{n+1}^k} \mathbb{L}_{n+1}^{k,j} \delta_{\mathbb{X}_{n+1}^{k,j}}$ 

**Remark 2.3.** (1) weights particles by their odds of producing the last observation.  $(3-5)$  resample the particles without bias, killing unlikely particles and duplicating likely ones while keeping the expected number of particles and total mass of all the particles constant. Parameter  $a_n, b_n$  in  $(3,4)$  control the amount of resampling.  $a_n = -\infty, b_n = \infty$  turns off resampling and results in the weighted particle system.  $a_n =$  $b_n$  ensure complete resampling and gives an unnormalized version of the better algorithm in Del Moral, Kouritzin and Miclo [\[7](#page-50-1)].

**Remark 2.4.** The  $\{V_n^{k,i}\}$  are required for analytical reasons. They provide enough smoothness that we can compare this resampled branching particle filter to a reduced McKean-Vlasov particle system. Without these V's the resampling events would be abrupt in the weight values.

<span id="page-10-1"></span>Remark 2.5. The algorithm can fail. During resampling, there is a possibility of immediately killing all particles if max  $j \leq N_{n-1}^k, k \leq m_N$  $m_N \widehat{\mathbb L}_{n}^{k,j}$  $\sum_{k=1}^{m_N}\sum_{i=1}^{\mathbb{N}_n^k}\widehat{\mathbb{L}}_n^{k,j}$ < 1. Ironically, this can only happen if there are more particles

than at start. However, it may still be possible to degenerate immediately to one particle when  $\sum_{k=1}^{m_N} \mathbb{N}_n^k \leq m_N$ .

Conversely, it is not possible to increase by more than  $m_N - 1$  particles in one step. The weight variation is a big concern:  $\mathbb{L}_n^{k,j}$  can become very uneven as  $m_N$  increases. Some regularity results are required to ensure that there are enough effective particles and moment bounds to justify the anticipation of the clt as  $m_N \rightarrow \infty$ .

<span id="page-10-0"></span>To rationalize the use of  $m_N - N$  extra particles, we quote the clt (see Weber [\[22\]](#page-51-8)) for triangular sequences of exchangeable random variables:

**Theorem 2.1.** Suppose  $\{X_{N,j} : j = 1, ..., m_N\}$  are exchangeable random variables for all  $N = 2, 3, ...$  and:

(i) 
$$
\frac{N}{m_N} \to 0
$$
, (ii)  $NE[X_{N,1}^2] \to 1$ , (iii)  $\sum_{j=1}^N X_{N,j}^2 \to^P 1$ , (iv)  $N^2E[X_{N,1}X_{N,2}] \to 0$ , and (v)  $\max_{j \le N} |X_{N,j}| \to^P 0$ .  
Then,  $\sum_{j=1}^N X_{N,j} \xrightarrow{\mathcal{D}} \mathcal{N}(0,1)$ .

Notice  $m_N - N$  extra random variables are required for the desired central limit theorem. Moreover, when using our resampled branching particle filter in practice, you can take  $m<sub>N</sub>$  to be something like  $m_N = N(1 + \log \log N)$  (for large enough N) so you may not add many extra particles until N is very large. Finally, the Mlln rates of convergence hold even for  $m_N = N$  so the extra particles are really only for characterizing performance.

#### 3. Notation, Unnormalized Filter, Weighted Approximation

Since  $Q(X_{n+1} \in A | \mathcal{F}_n^X) = K(X_n, A)$ , one has  $E^Q[f(X_n) | \mathcal{F}_{n-1}^X] = E^P[f(X_n) | \mathcal{F}_{n-1}^X] = Kf(X_{n-1})$ . For any finite measure  $\mu$  and integrable function  $f$ , we define

$$
\mu f = \int_{E} f(x) \mu(dx), \quad K^{n}(y, dx) = \int_{E} K^{n-1}(z, dx) K(y, dz)
$$

$$
\mu K^{n}(dx) = \int_{E} K^{n}(z, dx) \mu(dz) \text{ and } K^{n} f(x) = \int_{E} f(z) K^{n}(x, dz)
$$

for all  $n = 2, 3, ...$  with  $K^1 = K$ . Now, recall E is a Polish space and let  $B(E)$ ,  $B(E)_+$ ,  $C(E)_{++}$ ,  $\overline{C}(E)$ and  $\overline{C}(E)_+$  be the bounded measurable, non-negative bounded, strictly-positive continuous, continuous bounded, and non-negative continuous bounded functions respectively and define  $|f|_{\infty} = \sup_{x \in E} |f(x)|$ . Clearly,  $Kf \in B(E)_+$  if  $f \in B(E)_+$ . We use the extended Vinogradov symbol (introduced in [\[14](#page-51-9)]): Suppose  $a(n, m)$ ,  $b(n, m)$  are expressions depending upon two sets of variables  $n, m$ . Then,  $a(n, m) \ll b(n, m)$  means there exists a  $c_m > 0$ , depending only on m, such that  $a(n, m) \leq c_m b(n, m)$  for all n,m.

<span id="page-11-0"></span>Lastly, we define the operators  $A_n$  and  $A_{i,n}$  as

$$
A_n f(x) = \begin{cases} \alpha_n(x) K f(x), & n \in \mathbb{N} \\ f(x), & n = 0 \end{cases} \text{ and } (3.1)
$$

$$
A_{i,n}f(x) = \begin{cases} A_i (A_{i+1} \cdots (A_n f))(x), & i \le n \\ f(x), & i = n+1 \end{cases}
$$
 (3.2)

Then,  $\sigma_0 = \pi_0$  and, using [\(1.1,](#page-1-0)[1.2\)](#page-1-1), we have the following recursion for  $\sigma_n$ :

<span id="page-11-1"></span>
$$
\sigma_n(f) = \sigma_{n-1}(A_n f) \ \forall n = 1, 2, ..., \tag{3.3}
$$

Applying this recursion repeatedly, we have that

$$
\sigma_n\left(f\right) = \pi_0\left(A_{1,n}f\right). \tag{3.4}
$$

Bayes' rule implies that  $\pi_n(f) = \frac{\sigma_{n-1}(A_n f)}{\sigma_{n-1}(A_n)} = \frac{\pi_0(A_{1,n}f)}{\pi_0(A_{1,n}f)}$ .

Weighted particle filters approximate the unnormalized filter without resampling. The conditional expectation  $\sigma_n(f) = E^Q[L_n f(X_n)] \mathcal{F}_n^Y]$  with respect to reference probability Q is replaced with an independent sample average to obtain

$$
\sigma_n^N\left(f\right) = \frac{1}{N} \sum_{k=1}^N L_n^k f\left(\mathfrak{X}_n^k\right),\tag{3.5}
$$

our weighted-particle estimator of  $\sigma_n(f)$ , where the *particles*  $\{\mathfrak{X}^k\}_{k=1}^\infty$  are independent  $(\pi_0, K)$ -Markov processes that are independent of Y and the *weights* satisfy  $L_n^k = \prod_{i=1}^{n-k}$  $\prod_{j=1}^{\ldots} \alpha_j (\mathfrak{X}_{j-1}^k).$ 

In the sequel, we will fix an observation path, set  $Q^Y(\cdot) = Q(\cdot | \mathcal{F}_{\infty}^Y)$  and let  $E^Y[Z]$  denote expectation with respect to  $Q^Y$ .

#### 4. Reduced McKean-Vlasov Particle System

The problem with the weighted particle system is, due to randomness, most particles do not behave like the signal so their weights become small compared to the weights of very few good particles. This results in a particle filter that effectively consists of an average over only a very small portion of the particles. This problem manifests itself theoretically in the large expected variance of the central limit theorem and practically in the need to use a huge number of particles in most applications. Indeed, the weighted particle filter might not work regardless of the number of particles. To combate these problems, one introduces resampling. Initially, we pretend herein that we have access to the actual unnormalized filter total mass  $\{\sigma_n(1), n = 0, 1, 2, ...\}$  and consider an unimplementable reduced system of McKean-Vlasov type. In particular, we use the algorithm given in Section 2 with  $\mathbb{A}_n$  replaced with  $\sigma_n(1)$ . To facilitate analysis, we make explicit reference to the random variables that drive the particle system. Suppose we have enlarged  $(\Omega, \mathcal{F}, Q)$  to support the following random variables:

- 1.  $\{\chi^k\}_{k=1}^{\infty}$  are independent samples from  $\pi_0$ ,
- 2.  $\{\mathcal{Z}_n^{k,i,x}: n, k, i \in \mathbb{N}, x \in E\}$  are independent with  $\mathcal{Z}_n^{k,i,x}$  having distribution  $K(x, \cdot),$
- 3.  $\{\mathcal{U}_n^{k,i}: n, k, i \in \mathbb{N}\}\$ are independent and Uniform $[0, 1],$
- 4.  $\{\mathcal{V}_n^{k,i} : n, k, i \in \mathbb{N}\}$  are zero mean, i.i.d. with common pdf  $f_V$ ,

which are mutually independent and independent of  $X, Y$ . The actual pdf  $f_V$  does not matter for this section but has to be bounded in the next section. k is used to denote the first ancestor of each particle. Then, our reduced particle filter will be the average of  $N$  *i.i.d.* weighted branching Markov processes  $\{\mathcal{B}_n^k, n=0,1,...\}$  each starting from an independent sample  $\delta_{\chi^k}$ . All particles evolve independently of each other only interacting with  $\{\sigma_n(1)\}\)$ , which is deterministic with respect to  $Q^Y$ . At any time, many of the  $\mathcal{B}^k$  may have died out while others have branched into multiple particles. For clarity, the particles at time n (if any) that are offspring of the initial particle  $\chi^k$  will be denoted  $\{\mathcal{X}_n^{k,i}\}_{i=1}^{N_n^k}$  and the weight of such a particle after resampling will be denoted  $\mathcal{L}_n^{k,i}$ . Then, the branching Markov process corresponding to the  $k<sup>th</sup>$  original particle and the complete filter estimate will be

<span id="page-12-0"></span>
$$
\mathcal{B}_n^k = \sum_{i=1}^{\mathcal{N}_n^k} \mathcal{L}_n^{k,i} \delta_{\mathcal{X}_n^{k,i}} \text{ and } \mathcal{S}_n^N = \frac{1}{N} \sum_{k=1}^N \mathcal{B}_n^k \tag{4.1}
$$

respectively. We define the branching Markov processes  $\{\mathcal{B}^k\}$  as follows:

**Initialize:** 
$$
\mathcal{X}_0^{k,1} = \chi^k
$$
,  $\mathcal{N}_0^k \mathcal{L}_0^{k,1} = 1 \ \forall k = 1, ..., m_N$ ;  $\mathcal{N}_n^{k,i} = 0 \ \forall i, k, n \in \mathbb{N}$ .

**Repeat:** for  $n = 0, 1, 2, ...$  do

**Repeat (1-6):** for  $k = 1, 2, ..., m_N$  do

**Repeat (1-5):** for  $i = 1, 2, ..., \mathcal{N}_n^k$  do

1. Weight:

<span id="page-12-1"></span>
$$
\hat{\mathcal{L}}_n^{k,i} = \alpha_{n+1}(\mathcal{X}_n^{k,i})\mathcal{L}_n^{k,i}
$$
\n(4.2)

<span id="page-13-4"></span>2. Resample Case: If  $\hat{\mathcal{L}}_n^{k,i} + \mathcal{V}_{n+1}^{k,i} \notin (a_n \sigma_{n+1}(1), b_n \sigma_{n+1}(1))$  then

$$
\mathcal{N}_{n+1}^{k,i} = \left\lfloor \frac{\widehat{\mathcal{L}}_n^{k,i}}{\sigma_{n+1}(1)} \right\rfloor + 1_{\mathcal{U}_{n+1}^{k,i} + \left\lfloor \frac{\widehat{\mathcal{L}}_n^{k,i}}{\sigma_{n+1}(1)} \right\rfloor \le \frac{\widehat{\mathcal{L}}_n^{k,i}}{\sigma_{n+1}(1)}} \overline{\mathcal{L}}_n^{k,i} = \sigma_{n+1}(1) \tag{4.3}
$$

3. Non-resample Case: If  $\hat{\mathcal{L}}_n^{k,i} + \mathcal{V}_{n+1}^{k,i} \in (a_n \sigma_{n+1}(1), b_n \sigma_{n+1}(1))$  then  $\overline{\mathcal{L}}_n^{k,i} = \widehat{\mathcal{L}}_n^{k,i}, \, \mathcal{N}_{n+1}^{k,i} = 1$ 

4. Combine:  $\widehat{\mathcal{X}}_n^{k,j} \triangleq \mathcal{X}_n^{k,i}, \mathcal{L}_{n+1}^{k,j} \triangleq \overline{\mathcal{L}}_n^{k,i}$  for  $j \in \left\{ \right.$  $\overline{\mathcal{N}}_{n+1}^{k,i-1}$  + 1, ...,  $\overline{\mathcal{N}}_{n+1}^{k,i}$  }, where

<span id="page-13-1"></span>
$$
\overline{\mathcal{N}}_{n+1}^{k,i-1} = \sum_{j=1}^{i-1} \mathcal{N}_{n+1}^{k,j}
$$
\n(4.4)

5. Evolve Independently:  $\mathcal{X}_{n+1}^{k,j} = \mathcal{Z}_{n+1}^{k,j,\hat{\mathcal{X}}_n^{k,j}}$  for  $j \in \left\{ \right.$  $\overline{\mathcal{N}}^{k,i-1}_{n+1} + 1, ..., \overline{\mathcal{N}}^{k,i}_{n+1}$ 6. Estimate:  $\mathcal{B}_{n+1}^k = \sum_{n=1}^{\mathcal{N}_{n+1}^k}$  $\sum_{j=1}^{n+1} \mathcal{L}_{n+1}^{k,j} \delta_{\mathcal{X}_{n+1}^{k,j}}$ , where  $\mathcal{N}_{n+1}^k = \mathcal{N}_{n+1}^{k,1} + \cdots + \mathcal{N}_{n+1}^{k,\mathcal{N}_{n}^k}$ .

**Remark 4.1.** This reduced filter can plunge into a zero particle trap if max  $j \leq \mathcal{N}_{n-1}^k, k \leq m_N$  $\frac{\widehat{\mathcal{L}}_n^{k,j}}{\sigma_{n+1}(1)} < 1$ . The

weights can also become very uneven. We defined an extra  $m_N - N$  particles that were independent of the other particles and not used in the estimate. This was entirely for comparison with the resampled system (given in the introduction), where the extra particles are required to establish the central limit theorem.

**Remark 4.2.** To handle the index change in Step 5, we use the parent operators

$$
p_{n+1}(j) = i \quad such \, that \, j \in \{\overline{\mathcal{N}}_{n+1}^{k,i-1} + 1, ..., \overline{\mathcal{N}}_{n+1}^{k,i}\}.
$$
\n
$$
(4.5)
$$

<span id="page-13-2"></span>This i is unique.  $p_{n+1}$  is defined explicitly in a slightly different context in [\(7.45\)](#page-30-0) to follow.

After Step (4), we have  $\mathcal{N}_{n+1}^{k,i}$  particles at location  $\mathcal{X}_n^{k,i}$  each with weight  $\overline{\mathcal{L}}_n^{k,i}$ . Hence, the expected weight at location  $\mathcal{X}_n^{k,i}$  after possible resampling satisfies:

$$
E^{Y}\left[\overline{\mathcal{L}}_{n}^{k,i}\mathcal{N}_{n+1}^{k,i} \left|\mathcal{F}_{n}^{\mathcal{U}\mathcal{X}}\vee\mathcal{F}_{n+1}^{\mathcal{V}}\right.\right] = \widehat{\mathcal{L}}_{n}^{k,i} \quad \forall i = 1, 2, ..., \mathcal{N}_{n}^{k},\tag{4.6}
$$

<span id="page-13-3"></span><span id="page-13-0"></span>which is the weight in (1) prior to resampling, so the system is *unbiased*. However, we need to go further and establish a martingale property. First, averaging over the  $\mathcal{U}_n^{k,i}$ , one has

$$
E^{Y}\left[\frac{\overline{\mathcal{N}}_{n}^{k,i}}{j=\overline{\mathcal{N}}_{n}^{k,i-1}+1}f(\mathcal{X}_{n}^{k,j})\bigg|\mathcal{F}_{n-1}^{\mathcal{U}^{k,i}}\vee\mathcal{F}_{n}^{\mathcal{V}\mathcal{X}}\right]
$$
\n
$$
=E^{Y}\left[\sum_{j=\overline{\mathcal{N}}_{n}^{k,i-1}+1}^{\widehat{\mathcal{N}}_{n}^{k,i}}f(\mathcal{X}_{n}^{k,j})+\left|\frac{\widehat{\mathcal{L}}_{n-1}^{k,i}}{\overline{\mathcal{L}}_{n-1}^{k,i}}-\left|\frac{\widehat{\mathcal{L}}_{n-1}^{k,i}}{\overline{\mathcal{L}}_{n-1}^{k,i}}\right|\bigg|f(\mathcal{X}_{n}^{k,\widehat{\mathcal{N}}_{n}^{k,i+1}})\bigg|\mathcal{F}_{n-1}^{\mathcal{U}^{k,i}}\vee\mathcal{F}_{n}^{\mathcal{V}\mathcal{X}}\right],
$$
\n(4.7)

where  $\widehat{\mathcal{N}}_n^{k,i} = \overline{\mathcal{N}}_n^{k,i-1}$  +  $\left| \frac{\widehat{\mathcal{L}}_{n-1}^{k,i}}{\frac{1}{n} \widehat{k}^{i}} \right|$  $\overline{\mathcal{L}}^{k,i}_{n-1}$  $\overline{\phantom{a}}$ and  $\mathcal{F}_{n-1}^{\mathcal{U}^{k,i}} = \sigma\{\mathcal{U}_m^{l,j} : m \leq n, (l,j,m) \neq (k,i,n)\}\.$  (Notice [\(4.7\)](#page-13-0) holds whether we resample or not.) Using [\(4.7\)](#page-13-0) plus the facts  $\mathcal{N}_{n-1}^k \in \mathcal{F}_{n-1}^{UVX}$  and  $(\mathcal{L}_n^{k,j}, E^Y[f(\mathcal{X}_n^{k,j})|\mathcal{F}_{n-1}^X \vee$  $\mathcal{F}_n^{U\mathcal{V}} \vee$ ]) =  $(\overline{\mathcal{L}}_{n-1}^{k,i})$  $_{n-1}^{k,i}, Kf(\mathcal{X}_{n-1}^{k,i}))$  for  $j \in \left\{$  $\overline{\mathcal{N}}_n^{k,i-1}+1,...,\overline{\mathcal{N}}_n^{k,i}$  $\binom{k,i}{n}$ , one finds by  $(4.1, 4.4, 3.1)$  that

<span id="page-14-0"></span>
$$
E^{Y}[\mathcal{B}_{n}^{k}(f)|\mathcal{F}_{n-1}^{UVX}] = E^{Y} \left[ \sum_{j=1}^{N_{n}^{k}} \mathcal{L}_{n}^{k,j} f(\mathcal{X}_{n}^{k,j}) \middle| \mathcal{F}_{n-1}^{UVX} \right]
$$
\n
$$
= \sum_{i=1}^{N_{n-1}^{k}} E^{Y} \left[ \sum_{j=N_{n}^{k,i-1}+1}^{\overline{N}_{n}^{k,i}} \mathcal{L}_{n}^{k,j} f(\mathcal{X}_{n}^{k,j}) \middle| \mathcal{F}_{n-1}^{UVX} \right]
$$
\n
$$
= \sum_{i=1}^{N_{n-1}^{k}} E^{Y} \left[ \frac{\hat{\mathcal{L}}_{n-1}^{k,i}}{\overline{\mathcal{L}}_{n-1}^{k,i}} \overline{\mathcal{L}}_{n-1}^{k,i} K f(\mathcal{X}_{n-1}^{k,i}) \middle| \mathcal{F}_{n-1}^{UVX} \right]
$$
\n
$$
= \sum_{i=1}^{N_{n-1}^{k}} \alpha_{n} (\mathcal{X}_{n-1}^{k,i}) \mathcal{L}_{n-1}^{k,i} K f(\mathcal{X}_{n-1}^{k,i})
$$
\n
$$
= \mathcal{B}_{n-1}^{k}(A_{n}f) \text{ subject to } \mathcal{B}_{0}^{k}(f) = f(\chi^{k}).
$$
\n(4.8)

(One can check this equation in the two cases:  $\mathcal{N}_{n-1}^k = 0$  and  $\mathcal{N}_{n-1}^k \neq 0$ .) Using [\(4.8\)](#page-14-0) recursively, one finds by [\(3.2](#page-11-0)[,3.4\)](#page-11-1) that

$$
E^{Y}[B_{n}^{k}(f)] = E^{Y}[A_{1,n}f(\chi^{k})] = \sigma_{n}(f)
$$
\n(4.9)

<span id="page-14-2"></span>so by [\(4.8,](#page-14-0)[4.9\)](#page-14-1)

<span id="page-14-1"></span>
$$
\mathcal{B}_n^k(f) -\sigma_n(f) = M_n^{\mathcal{B}^k}(f), \quad \text{where}
$$
\n(4.10)

$$
M_n^{\mathcal{B}^k}(f) = \sum_{l=0}^n \left[ \mathcal{B}_l^k (A_{l+1,n} f) - E^Y \left[ \mathcal{B}_l^k (A_{l+1,n} f) \left| \mathcal{F}_{l-1}^{\mathcal{U} \mathcal{Y} \mathcal{X}} \right] \right] \right]. \tag{4.11}
$$

 ${M_n^{\mathcal{B}^k}(f)}$ ,  $n = 0, 1, ...$  is a zero-mean  ${\{\mathcal{F}_n^{\mathcal{U} \mathcal{Y} \mathcal{X}}\}}_{n=0}^{\infty}$ -martingale with respect to  $Q^Y$ . Averaging over the initial ancestrial branches  $k$ , one finds by  $(4.1, 4.8, 4.9, 4.10, 4.11)$  $(4.1, 4.8, 4.9, 4.10, 4.11)$  $(4.1, 4.8, 4.9, 4.10, 4.11)$  that

$$
E^{Y}[S_n^N(f)|\mathcal{F}_{n-1}^{UVX}] = S_{n-1}^N(A_nf) \text{ subject to } S_0^N(f) = \frac{1}{N} \sum_{k=1}^N f(\chi^k)
$$
(4.12)

$$
E^{Y}[\mathcal{S}_n^N(f)] = \sigma_n(f) \tag{4.13}
$$

$$
\mathcal{S}_n^N(f) = \sigma_n(f) + \mathcal{M}_n^N(f) \tag{4.14}
$$

<span id="page-14-3"></span>with

$$
\mathcal{M}_{n}^{N}(f) = \frac{1}{N} \sum_{k=1}^{N} M_{n}^{\mathcal{B}^{k}}(f)
$$
\n
$$
= \sum_{l=0}^{n} \left[ \mathcal{S}_{l}^{N}(A_{l+1,n}f) - E^{Y} \left[ \mathcal{S}_{l}^{N}(A_{l+1,n}f) \left| \mathcal{F}_{l-1}^{U\mathcal{V}\mathcal{X}} \right| \right] \right].
$$
\n(4.15)

Now, we define the  $\mathcal{F}_{\infty}^Y$ -measurable random variance

$$
\gamma_n^P(f) = E^Y[|M_n^{\mathcal{B}^1}(f)|^2].\tag{4.16}
$$

Recall  $\sigma_n(f)$ ,  $\alpha_n$  from  $(1.2)$ , $(1.1)$  respectively. To find an expression for the variance  $\gamma_n^P(f)$  of this reduced system and the resampled system to follow, we define the resampling function:

$$
r(x) = x - \lfloor x \rfloor - (x - \lfloor x \rfloor)^2,\tag{4.17}
$$

<span id="page-15-3"></span>which is an artifact of resampling and is clearly bounded by  $\frac{1}{4}$ . Now, let

<span id="page-15-2"></span>
$$
\nu_n(l) = \int 1_{(a_{n-1}\sigma_n(1), b_{n-1}\sigma_n(1))}(s) f_V(l-s) ds, \ \overline{\nu}_n(l) = 1 - \nu_n(l). \tag{4.18}
$$

<span id="page-15-5"></span>For notational simplicity, we recall  $\sigma_0(1) = \pi_0(1) = 1$  and define

$$
\alpha_{i,m}(x_i,...,x_{m-1}) = \alpha_m(x_{m-1}) \cdots \alpha_{i+2}(x_{i+1}) \alpha_{i+1}(x_i) \sigma_i(1) \tag{4.19}
$$

$$
\nu_{i,m}(x_i,...,x_{m-1}) = \nu_m(\alpha_{i,m}(x_i,...,x_{m-1})) \cdots \nu_{i+1}(\alpha_{i,i+1}(x_i)) \tag{4.20}
$$

$$
\overline{\nu}_{i,m}(x_i,...,x_{m-1}) = \overline{\nu}_m(\alpha_{i,m}(x_i,...,x_{m-1}))
$$
\n(4.21)

so  $\alpha_{i,i}(x) = \sigma_i(1)$  and  $\nu_{i,i}(x) = 1$ . The following proposition gives the clt variance for both the reduced McKean-Vlasov particle system and, as will be shown later, the partially-resampled particle filter in terms of the resampling used. The proof is necessarily technical, and hence delayed until Appendix 1.

<span id="page-15-1"></span>**Proposition 4.1.** Let h be bounded and 
$$
\sum_{\substack{i_1 < \dots < i_j \\ j < l}} \text{ denote the sum over } 1 \leq i_1 < \dots < i_j < l \text{ and } 0 \leq j < l
$$

<span id="page-15-0"></span> $l \leq n$ . Then,

$$
\gamma_n^P(f) = \pi_0((A_{1,n}f)^2) - (\pi_0(A_{1,n}f))^2
$$
\n
$$
+ \sum_{\substack{i_1 < \dots < i_j \\ j < l}} \sigma_l(1)\pi_0[A_{1,l-1}\{A_l(A_{l+1,n}f)^2 - \alpha_l(KA_{l+1,n}f)^2\} \overline{\nu}_{i_j,l} \nu_{i_j,l-1} \overline{\nu}_{i_1,i_2,\dots,i_j}]
$$
\n
$$
+ \sum_{\substack{i_1 < \dots < i_j \\ j < l}} \pi_0[A_{1,l-1}\alpha_{i_j,l}\{A_l(A_{l+1,n}f)^2 - \alpha_l(KA_{l+1,n}f)^2\} \nu_{i_j,l} \overline{\nu}_{i_1,i_2,\dots,i_j}]
$$
\n
$$
+ \sum_{\substack{i_1 < \dots < i_j \\ j < l}} \sigma_l^2(1)\pi_0[A_{1,l-1}\frac{\overline{\nu}_{i_j,l}}{\alpha_{i_j,l-1}}r\left(\frac{\alpha_{i_j,l}}{\sigma_l(1)}\right)(KA_{l+1,n}f)^2 \nu_{i_j,l-1} \overline{\nu}_{i_1,i_2,\dots,i_j}]
$$
\n(4.22)

for all  $f \in B(E)_+$ , where

$$
\overline{\nu}_{i_1, i_2, \dots, i_j} \triangleq \overline{\nu}_{i_{j-1}, i_j} \cdots \nu_{i_1, i_2 - 1} \overline{\nu}_{0, i_1} \nu_{0, i_1 - 1}
$$
\n(4.23)

<span id="page-15-4"></span> $A_{1,m}$  is defined in [\(3.2\)](#page-11-0) and operator  $A_i$  applies to the last argument of  $A_{i+1,m}\phi_m(x_0, x_1, ..., x_{i-1}, x_i)$ .

**Remark 4.3.** The first term on the right of  $(4.22)$  represents the error variance of introducing an independent particle system. The remaining terms incorporate the resampling scheme used. To understand this formula, we can think of  $j \in \{0, 1, ..., l-1\}$  as a number of resampling events up to  $l-1$  and  $i_1, i_2, ..., i_j$ as possible resample times up to  $l - 1$  so the system would run without resampling between these times.  $\nu_{i_j,l-1}$  $\overline{\nu}_{i_{j-1},i_j}$   $\cdots$   $\nu_{i_1,i_2-1}$   $\overline{\nu}_{0,i_1}\nu_{0,i_1-1}$  is then the joint probability that these are the resample times. In particular,  $\nu_{0,i_1-1}$  is the probability of not resampling before  $i_1$  and  $\overline{\nu}_{0,i_1}$  is the conditional probability of resampling at i<sub>1</sub> given no prior resampling. Under our conditions, each  $\sigma_l$  is a finite measure and  $\frac{\overline{v}_{i_j,l}}{\alpha_{i_j,l-1}},$  $\alpha_l$ ,  $A_{l,n}f$  are bounded for each fixed  $Y_1, ..., Y_n$ ,  $f \in B(E)_+$  so  $\gamma_n^P(f)$  is an R-valued random variable.

<span id="page-16-0"></span>To facilitate the discussion to follow, we break the final two terms of [\(4.22\)](#page-15-0) into the cases of resampling at time  $l - 1$  and not, which yields:

$$
\gamma_{n}^{P}(f) = \pi_{0}((A_{1,n}f)^{2}) - (\pi_{0}(A_{1,n}f))^{2}
$$
\n
$$
+ \sum_{l=1}^{n} \sigma_{l}(1) \sum_{j=0}^{l-1} \sum_{1 \leq i_{1} < \dots < i_{j} < l} \pi_{0}[A_{1,l-1}\{f_{l,n}\}\overline{\nu}_{i_{j},l}\nu_{i_{j},l-1}\overline{\nu}_{i_{1},i_{2},\dots,i_{j}}]
$$
\n
$$
+ \sum_{l=2}^{n} \sum_{j=0}^{l-2} \sum_{1 \leq i_{1} < \dots < i_{j} < l-1} \pi_{0}[A_{1,l-1}\alpha_{i_{j},l}\{f_{l,n}\}\nu_{i_{j},l}\overline{\nu}_{i_{1},i_{2},\dots,i_{j}}]
$$
\n
$$
+ \sum_{l=1}^{n} \sigma_{l-1}(1) \sum_{j=1}^{l-1} \sum_{1 \leq i_{1} < \dots < i_{j} = l-1} \pi_{0}[A_{1,l-1}\alpha_{l}\{f_{l,n}\}\nu_{i_{j},l}\overline{\nu}_{i_{1},i_{2},\dots,i_{j}}]
$$
\n
$$
+ \sum_{l=2}^{n} \sigma_{l}^{2}(1) \sum_{j=0}^{l-2} \sum_{1 \leq i_{1} < \dots < i_{j} < l-1} \pi_{0}[A_{1,l-1}\frac{\overline{\nu}_{i_{j},l}}{\alpha_{i_{j},l-1}}r\left(\frac{\alpha_{i_{j},l}}{\sigma_{l}(1)}\right)f^{l,n}\nu_{i_{j},l-1}\overline{\nu}_{i_{1},i_{2},\dots,i_{j}}]
$$
\n
$$
+ \sum_{l=1}^{n} \sigma_{l}^{2}(1) \sum_{j=1}^{l-1} \sum_{1 \leq i_{1} < \dots < i_{j} < l-1} \pi_{0}[A_{1,l-1}\frac{\overline{\nu}_{i_{j},l}}{\sigma_{l-1}(1)}r\left(\frac{\alpha_{l}\sigma_{l-1}(1)}{\sigma_{l}(1)}\right)f^{l,n}\overline{\nu}_{i_{1},i_{2},\dots,i_{j}}]
$$

<span id="page-16-2"></span>for all  $f \in B(E)_+$ , where  $f_{l,n} = A_l (A_{l+1,n} f)^2 - \alpha_l (K A_{l+1,n} f)^2$  and  $f^{l,n} = (K A_{l+1,n} f)^2$ .

<span id="page-16-1"></span>**Remark 4.4.** Notice, there are no  $j = 0$  cases in the fourth and sixth terms of  $(4.24)$ . For the second, third and fifth terms, the multiple sum over the i's degenerates to just one item,

$$
\sigma_l(1)\pi_0[A_{1,l-1}\left\{A_l(A_{l+1,n}f)^2 - \alpha_l(KA_{l+1,n}f)^2\right\}\overline{\nu}_{0,l}\nu_{0,l-1}],
$$
\n(4.25)

$$
\pi_0[A_{1,l-1}\alpha_{0,l}\left\{A_l(A_{l+1,n}f)^2-\alpha_l(KA_{l+1,n}f)^2\right\}\nu_{0,l}] \text{ and } (4.26)
$$

$$
\sigma_l^2(1)\pi_0[A_{1,l-1}\frac{1}{\alpha_{0,l-1}}r\left(\frac{\alpha_{0,l}}{\sigma_l(1)}\right)(KA_{l+1,n}f)^2\overline{\nu}_{0,l}\nu_{0,l-1}]
$$
\n(4.27)

respectively, when j = 0. Furthermore, in the non-resampled case where  $a_i = -\infty$  and  $b_i = \infty$  so  $\nu_i = 1$ , we have this  $j = 0$  case only but also we do not resample at time l either so terms  $(4.25)$  and  $(4.27)$  also disappear. Then, we can incorporate the  $\alpha_i$  into the operators by letting

$$
A_j^{(2)}f(x) = \begin{cases} \alpha_j^2(x)Kf(x) & j = 1, 2, ... \\ f(x) & j = 0 \end{cases} and \qquad (4.28)
$$

$$
A_{i,n}^{(2)}f = \begin{cases} A_i^{(2)}(A_{i+1}^{(2)} \cdots (A_n^{(2)}f)) & \forall i \le n \\ f & i = n+1 \end{cases},
$$
\n(4.29)

<span id="page-17-2"></span>and note that  $\nu_{0,l} = 1$  in this non-resampled case. Hence, the non-resampled case variance is

$$
\gamma_n^W(f) = \pi_0((A_{1,n}f)^2) - (\pi_0(A_{1,n}f))^2
$$
  
+ 
$$
\sum_{l=1}^n \pi_0 A_{1,l-1}^{(2)} \left[ A_l^{(2)} (A_{l+1,n}f)^2 - (A_{l,n}f)^2 \right] \ \forall f \in B(E)_+,
$$
\n(4.30)

which is the variance for the weighted particle filter.

<span id="page-17-1"></span><span id="page-17-0"></span>**Remark 4.5.** Full resampling occurs if all  $a_i = b_i$  so  $\overline{v}_i = 1$  so only the  $j = l - 1$  terms remain. The multiple sums over the *i*'s in the second, fourth and sixth terms of  $(4.24)$  reduce to

$$
\sigma_l(1)\pi_0[A_{1,l-1}\left\{A_l(A_{l+1,n}f)^2 - \alpha_l(KA_{l+1,n}f)^2\right\}\overline{\nu}_{l-1,l}],
$$
\n(4.31)

$$
\sigma_{l-1}(1)\pi_0[A_{1,l-1}\alpha_l\left\{A_l(A_{l+1,n}f)^2-\alpha_l(KA_{l+1,n}f)^2\right\}\nu_{l-1,l}],\tag{4.32}
$$

$$
\sigma_l^2(1)\pi_0\left[\frac{A_{1,l-1}}{\sigma_{l-1}(1)}r\left(\frac{\alpha_l\sigma_{l-1}(1)}{\sigma_l(1)}\right)(KA_{l+1,n}f)^2\overline{\nu}_{l-1,l}\right]
$$
\n(4.33)

<span id="page-17-3"></span>respectively since  $\overline{\nu}_{l-2,l-1} \cdots \overline{\nu}_{0,1} = 1$  in this case. However, [\(4.32\)](#page-17-0) also vanishes since  $\nu_{l-1,l} = 0$ . Therefore, the variance of the fully-resampled McKean-Vlasov system is by  $(1.1)$ ,  $(1.2)$  and  $(3.4)$ 

$$
\gamma_n^R(f) = \pi_0((A_{1,n}f)^2) - (\pi_0(A_{1,n}f))^2
$$
\n
$$
+ \sum_{l=1}^n \sigma_{l-1}(\alpha_l)\sigma_{l-1} (A_l(A_{l+1,n}f)^2 - \alpha_l(KA_{l+1,n}f)^2)
$$
\n
$$
+ \sum_{l=1}^n \sigma_{l-1}(\alpha_l)\sigma_{l-1} \left(\frac{\sigma_l(1)}{\sigma_{l-1}(1)}r\left(\frac{\alpha_l\sigma_{l-1}(1)}{\sigma_l(1)}\right)(KA_{l+1,n}f)^2\right)
$$
\n(4.34)

for all  $f \in B(E)_{+}$ . Comparing  $\gamma^{W}$  and the non-remainder part of  $\gamma^{R}$  (i.e. ignoring the last term of  $\gamma^{R}$ ), we see that the main difference is that the former uses  $A^{(2)}$  while the later uses A, so the function  $\alpha_l$  is not squared in  $\gamma^R$ . Roughly speaking, this means that the errors are not compounded to the same degrees.

Remark 4.6. By the above expressions and the proof (in the first appendix), we see that there is no need for h to be bounded in either the non-resampled (i.e. weighted) or fully-resampled case.

<span id="page-17-4"></span>This leads us to our main results of this section, which are laws of large numbers, rates of  $L^p$ -convergence and a quenched central limit theorem.

**Theorem 4.1.** Let h be bounded and g be positive and continuous. Then, for  $Q$ -a.a. Y, the reduced particle system satisfies:

 $\textbf{slln:}\;\mathcal{S}_n^N \Rightarrow \sigma_n\;\text{ (i.e. weak convergence)}\;\;a.s.\;[Q^Y];$ **Mlln:**  $\left| \mathcal{S}_{n}^{N}(f) - \sigma_{n}(f) \right|$  $\ll N^{-\beta}$  a.s.  $[Q^Y]$  for all  $f \in \overline{C}(E)_+$ ,  $0 \leq \beta < \frac{1}{2}$ ; L<sup>2</sup>-rates:  $E^Y \left| \mathcal{S}_n^N(f) - \sigma_n(f) \right|^2 = \frac{\gamma_n^P(f)}{N}$  for all  $f \in \overline{C}(E)_+$ ;  $L^p$ -rates:  $E^Y \left| \mathcal{S}_n^N(f) - \sigma_n(f) \right|$  $p \ll N^{-\frac{p}{2}}$  for all  $f \in \overline{C}(E)_+$ ,  $p \ge 1$ ; clt:  $\sqrt{N} \left( \mathcal{S}_n^N(f) - \sigma_n(f) \right) \Rightarrow \mathcal{N} \left( 0, \sqrt{\gamma_n^P(f)} \right)$  for all  $f \in \overline{C}(E)_+$ .

Proof.  $S_n^N(f) - \sigma_n(f) = \frac{1}{N} \sum_{n=1}^N$  $k=1$  $M_n^{\mathcal{B}^k}(f)$  is an average of i.i.d. random variables (see [\(4.15\)](#page-14-3)) so the theorem follows by  $(4.22)$ , the classical laws of large numbers,  $L^p$  bounds and central limit theorem. Note: 1)  $M_n^{\mathcal{B}^k}(f)$  is bounded for fixed  $Y_1, ..., Y_n$  by the following Lemma. 2)  $\mathcal{S}_n^N(f_i) \to \sigma_n(f_i)$  a.s.  $[Q^Y]$  for all i implies  $S_n^N \Rightarrow \sigma_n$  a.s.  $[Q^Y]$ , where

$$
\{f_i\}_{i=1}^{\infty} = \left\{ \prod_{j=1}^{l} (1 - \rho(\cdot, x_j)) \vee 0 : l \in \{0, 1, 2, \ldots\}, x_j \in \{y_k\}_{k=1}^{\infty} \right\},
$$
\n(4.35)

for some dense collection  $\{y_k\} \subset E$ . (See Blount and Kouritzin [\[2\]](#page-50-11) and note the product over zero functions is taken to be the constant function 1.)  $\Box$ 

The boundedness of  $M_n^{\mathcal{B}^k}(f)$ , required above follows from  $(4.11, 4.8, 4.1)$  $(4.11, 4.8, 4.1)$  $(4.11, 4.8, 4.1)$  $(4.11, 4.8, 4.1)$  and the following lemma.

<span id="page-18-2"></span>**Lemma 4.1.** Suppose h is bounded while g is positive and continuous. Then, there is a function  $C_n$ :  $\mathbb{R}^{dn} \to (0,\infty)$  such that the reduced system particle numbers and weights satisfy:

$$
\mathcal{N}_l^k, \max_{i \in \{1, ..., \mathcal{N}_l^k\}} \mathcal{L}_l^{k, i} \le C_n(Y_1, ..., Y_n) \quad \forall k \in \{1, ..., m_N\}, l \in \{0, ..., n\} \text{ on } \Omega.
$$

*Proof.* Let  $W_l^{k,i} = \alpha_l(\mathcal{X}_{l-1}^{k,i})$  with  $\alpha_l$  defined in [\(1.1\)](#page-1-0). Since

$$
0 < \inf_{x \in E} \frac{g(Y_l - h(x))}{g(Y_l)} < \sup_{x \in E} \frac{g(Y_l - h(x))}{g(Y_l)} < \infty
$$

<span id="page-18-0"></span>and  $\sigma_l$  is a positive finite measure for each  $l \in \mathbb{N}$ , there is a  $C = C(Y_1, ..., Y_n) > 1$  such that

$$
\frac{1}{C} \le \sigma_l(1), \mathcal{W}_l^{k,i} \le C
$$
\n
$$
\forall i = 1, ..., \mathcal{N}_{l-1}^k; l = 1, ..., n; k = 1, ..., m_N; N = 1, 2, ....
$$
\n(4.36)

<span id="page-18-1"></span>Now, recall from the reduced system algorithm (given above) that

$$
\mathcal{L}_{l+1}^{k,j} \le \sigma_{l+1}(1) \vee \mathcal{W}_{l+1}^{k,p_{l+1}(j)} \mathcal{L}_l^{k,p_{l+1}(j)} \tag{4.37}
$$

$$
\mathcal{N}_{l+1}^{k} = \sum_{i_{l}=1}^{\mathcal{N}_{l}^{k}} \mathcal{N}_{l+1}^{k,i_{l}} \leq \sum_{i_{1}=1}^{\mathcal{N}_{1}^{k}} \sum_{i_{2}=\overline{\mathcal{N}}_{2}^{k,i_{1}-1}+1}^{\overline{\mathcal{N}}_{2}^{k,i_{1}}-1} \cdots \sum_{i_{l}=\overline{\mathcal{N}}_{l}^{k,i_{l-1}-1}+1}^{\overline{\mathcal{N}}_{l}^{k,i_{l-1}-1}} \left[ \frac{\mathcal{L}_{l}^{k,i_{l}} \mathcal{W}_{l+1}^{k,i_{l}}}{\sigma_{l+1}(1)} + 1 \right] \tag{4.38}
$$

for  $j = 1, ..., N_{l+1}^k; k = 1, 2, ..., m_N$ , where the parent operator p is defined in [\(4.5\)](#page-13-2). Now, the stated bounds follow from  $(4.36, 4.37, 4.38)$ , the fact  $\mathcal{N}_0^k = \mathcal{L}_0^{k,1} = 1$  and induction.  $\Box$ 

 $\gamma_n^P(f)$  is  $\gamma_n^W(f)$  or  $\gamma_n^R(f)$  when there is no resampling or full resampling respectively, where  $\gamma_n^W(f)$ ,  $\gamma_n^R(f)$  are defined in Remarks [4.4,](#page-16-2) [4.5.](#page-17-1) h need not be bounded in these two cases.

Bounded regularity for the resampled system will not be so easy to come by but is handled in the next section.

#### 5. Resampled Particle System

The reduced system uses  $\sigma_n(1)$ , which is usually unrepresentable on a finite computer, so we use the particle filter algorithm in the introduction, expressed now in terms of random variables  $\{\chi^k\}$ ,  $\{\mathbb{U}_n^{k,i}\}$ ,  $\{\mathbb{V}_n^{k,i}\}$  and  $\{\mathbb{Z}_n^{k,i,x}\}$  analogous to those of the previous section. Particles can now interact weakly through an average weight process  $\{A_n^{m_N}, n = 0, 1, ...\}$ . However, we still break up the system by the first ancestor of each particle so our resampled particle filter will be the average of  $N$  exchangeable branching Markov processes  $\{\mathbb{B}_n^k, n=0,1,...\}$ , each starting from an independent sample  $\delta_{\chi^k}$ . For clarity, the particles at time n that are offspring from the original particle  $\chi^k$  will be denoted  $\{X_n^{k,i}\}_{i=1}^{N_n^k}$  and the weight of such a particle after resampling will be denoted  $\mathbb{L}_{n}^{k,i}$ . Then, the branching Markov process corresponding to this original particle and the complete (partially) resampled particle filter are:

$$
\mathbb{B}_n^k = \sum_{i=1}^{N_n^k} \mathbb{L}_n^{k,i} \delta_{\mathbb{X}_n^{k,i}} \text{ and } \mathbb{S}_n^N = \frac{1}{N} \sum_{k=1}^N \mathbb{B}_n^k.
$$
 (5.1)

<span id="page-19-1"></span>The branching Markov processes  $\{\mathbb{B}_n^k\}$  are defined by:

 $\textbf{Initialize: } \mathbb{X}_{0}^{k,1} = \chi^{k}, \, \mathbb{N}_{0}^{k} = 1 = \mathbb{L}_{0}^{k,1} \,\, \forall k = 1,2,...,m_{N}. \,\, \mathbb{N}_{n}^{k,i} = 0 \forall i,k,n \in \mathbb{N}.$ **Repeat:** for  $n = 0, 1, 2, ...$  do

- 1. Weight:  $\widehat{\mathbb{L}}_n^{k,i} = \alpha_{n+1} \left( \mathbb{X}_n^{k,i} \right) \mathbb{L}_n^{k,i}$  for  $i = 1, 2, ..., \mathbb{N}_n^k$ ,  $k = 1, 2, ..., m_{\mathbb{N}}$
- 2. Average Weight:

<span id="page-19-0"></span>
$$
\mathbb{A}_{n+1} = \frac{1}{m_N} \sum_{k=1}^{m_N} \sum_{i=1}^{N_n^k} \widehat{\mathbb{L}}_n^{k,i}
$$
\n(5.2)

**Repeat (3-7):** for  $k = 1, 2, ..., m_N$  do

- **Repeat (3-6):** for  $i = 1, 2, ..., N_n^k$  do
- 3. Resampled Case: If  $\hat{\mathbb{L}}_n^{k,i} + \mathbb{V}_{n+1}^{k,i} \notin (a_n \mathbb{A}_{n+1}, b_n \mathbb{A}_{n+1})$  then
	- (a) Offspring Numbers:  $\mathbb{N}_{n+1}^{k,i} =$  $\left\lfloor \frac{\widehat{\mathbb{L}}_n^{k,i}}{\mathbb{A}_{n+1}} \right\rfloor + 1_{\mathbb{U}_{n+1}^{k,i} + \left\lceil \frac{\widehat{\mathbb{L}}_n^{k,i}}{\mathbb{A}_{n+1}} \right\rceil \leq \frac{\widehat{\mathbb{L}}_n^{k,i}}{\mathbb{A}_{n+1}} }$ ,
	- (b) Resampled Weight:  $\overline{\mathbb{L}}_n^{k,i} = \mathbb{A}_{n+1}$
- 4. Non-resample Case: If  $\widehat{\mathbb{L}}_n^{k,i} + \mathbb{V}_{n+1}^{k,i} \in (a_n \mathbb{A}_{n+1}, b_n \mathbb{A}_{n+1})$  then  $\overline{\mathbb{L}}_n^{k,i} = \widehat{\mathbb{L}}_n^{k,i}, \, \mathbb{N}_{n+1}^{k,i} = 1$

5. Combine:  $\widehat{\mathbb{X}}_n^{k,j} \triangleq \mathbb{X}_n^{k,i}, \mathbb{L}_{n+1}^{k,j} \triangleq \overline{\mathbb{L}}_n^{k,i}$  for  $j \in \left\{ \overline{\mathbb{N}}_{n+1}^{k,i-1} + 1, ..., \overline{\mathbb{N}}_{n+1}^{k,i} \right\}$ , where

$$
\overline{\mathbb{N}}_{n+1}^{k,i} = \sum_{m=1}^{i} \mathbb{N}_{n+1}^{k,m}.
$$
\n(5.3)

6. Evolve Independently:  $\mathbb{X}_{n+1}^{k,j} = \mathbb{Z}_{n+1}^{k,j,\mathbb{X}_n^{k,j}}$  for all  $j \in \left\{ \overline{\mathbb{N}}_{n+1}^{k,i-1} + 1, ..., \overline{\mathbb{N}}_{n+1}^{k,i} \right\}$ 

7. Estimate: 
$$
\mathbb{B}_{n+1}^k = \sum_{j=1}^{N_{n+1}^k} \mathbb{L}_{n+1}^{k,j} \delta_{\mathbb{X}_{n+1}^{k,j}}
$$
, where  $\mathbb{N}_{n+1}^k = \mathbb{N}_{n+1}^{k,1} + \cdots + \mathbb{N}_{n+1}^{k,\mathbb{N}_n^k}$ .

Remark 5.1. For the index change in Step 5, we re-use the parent operator

$$
p_{n+1}(j) = i \quad such \, that \, j \in \{\overline{N}_{n+1}^{k,i-1} + 1, ..., \overline{N}_{n+1}^{k,i}\},\tag{5.4}
$$

defined now in terms of  $\overline{N}_{n+1}^{k,i}$  instead of  $\overline{N}_{n+1}^{k,i}$ . The context will make it clear for which system  $p_n$  is operating on.

<span id="page-20-2"></span>Remark 5.2. The distinguishing feature between the resampled and reduced particle filters is the resampling events. The resample sets for these systems are respectively

$$
\mathbb{H}_m^{k,i} = \left\{ \frac{\widehat{\mathbb{L}}_{m-1}^{k,i} + \mathbb{V}_m^{k,i}}{\mathbb{A}_m^{m_N}} \notin (a_{m-1}, b_{m-1}) \right\},
$$
\n(5.5)

$$
\mathcal{H}_m^{k,i} = \left\{ \frac{\hat{\mathcal{L}}_{m-1}^{k,i} + \mathcal{V}_m^{k,i}}{\sigma_m(1)} \notin (a_{m-1}, b_{m-1}) \right\}.
$$
\n(5.6)

The expected effective weight of resampled filter particle  $\mathbb{X}_n^i$  after resampling is:

$$
E^Y \left[ \overline{\mathbb{L}}_n^{k,i} \mathbb{N}_{n+1}^{k,i} \middle| \mathcal{F}_n^{\mathbb{U}\mathbb{X}} \vee \mathcal{F}_{n+1}^{\mathbb{V}} \right] = \widehat{\mathbb{L}}_n^{k,i},
$$

which is the weight before resampling so the system is *unbiased*. Moreover, noting  $\mathbb{L}_{n}^{k,i} \in \mathcal{F}_{n-1}^{\mathbb{UVX}}$ , one finds as in  $(4.7-4.8)$  $(4.7-4.8)$  that

$$
E^{Y}[\mathbb{B}_{n}^{k}(f) | \mathcal{F}_{n-1}^{\text{UVX}}] = \mathbb{B}_{n-1}^{k}(A_{n}f) \text{ subject to } \mathbb{B}_{0}^{k}(f) = f(\chi^{k}).
$$
\n(5.7)

<span id="page-20-0"></span>Using  $(5.7)$  recursively with  $(3.2)$  and  $(3.4)$ , one finds that

$$
E^{Y}[\mathbb{B}_{n}^{k}(f)] = \sigma_{n}(f) \text{ and } \mathbb{B}_{n}^{k}(f) - \sigma_{n}(f) = M_{n}^{\mathbb{B}^{k}}(f), \qquad (5.8)
$$

<span id="page-20-1"></span>with

$$
M_n^{\mathbb{B}^k}(f) = \sum_{l=0}^n \left[ \mathbb{B}_l^k (A_{l+1,n}f) - E^Y \left[ \mathbb{B}_l^k (A_{l+1,n}f) \, \middle| \mathcal{F}_{l-1}^{\mathbb{UV}X} \right] \right]
$$
\n
$$
= \sum_{l=0}^n \left[ \mathbb{B}_l^k (A_{l+1,n}f) - \mathbb{B}_{l-1}^k (A_{l,n}f) \right] \text{ if } \mathbb{B}_{-1}^k = \pi_0.
$$
\n(5.9)

Hence,  $E^{Y}[M_{n}^{\mathbb{B}^{k}}(f)] = 0$  by [\(5.9\)](#page-20-1). Moreover,  $\{M_{n}^{\mathbb{B}^{k}}(f), n = 0, 1, ...\}$  is a  $\{\mathcal{F}_{n}^{\mathbb{UVX}}\}$ -martingale with respect to  $Q<sup>Y</sup>$ . Averaging over the first N ancestrial branches, one finds that

$$
E^{Y}[\mathbb{S}_{n}^{N}(f)|\mathcal{F}_{n-1}^{\mathbb{UVX}}] = \mathbb{S}_{n-1}^{N}(A_{n}f) \text{ subject to } \mathbb{S}_{0}^{N}(f) = \frac{1}{N} \sum_{k=1}^{N} f(\chi^{k})
$$
(5.10)

$$
E^Y[\mathbb{S}_n^N(f)] = \sigma_n(f) \tag{5.11}
$$

$$
\mathbb{S}_n^N(f) = \sigma_n(f) + \mathbb{M}_n^N(f) \tag{5.12}
$$

with

$$
\mathbb{M}_{n}^{N}(f) = \frac{1}{N} \sum_{k=1}^{N} M_{n}^{\mathbb{B}^{k}}
$$
\n
$$
= \sum_{l=0}^{n} \left[ \mathbb{S}_{l}^{N} \left( A_{l+1,n} f \right) - E^{Y} \left[ \mathbb{S}_{l}^{N} \left( A_{l+1,n} f \right) \middle| \mathcal{F}_{l-1}^{\mathbb{U} \mathbb{V} \mathbb{X}} \right] \right].
$$
\n(5.13)

This leads to our main result, laws of large numbers and a quenched clt for our resampled particle filter.

<span id="page-21-0"></span>**Theorem 5.1.** Suppose  $m_N \geq N$ ; h and  $f_V$  are bounded; and g is strictly positive and continuous. Then, for any  $n \in \mathbb{N}$  and Q-almost all Y, the resampled particle filter satisfies:

slln:  $\mathbb{S}_n^N \Rightarrow \sigma_n$  (i.e. weak convergence) a.s.  $[Q^Y]$ ; **Mlln:**  $\left| \mathbb{S}_{n}^{N}(f) - \sigma_{n}(f) \right|$  $\stackrel{N}{\ll} N^{-\beta}$  a.s.  $[Q^Y]\,\forall f \in \overline{C}(E)_+, 0 \leq \beta < \frac{1}{2}$ ; clt:  $\sqrt{N} (\mathbb{S}_n^N(f) - \sigma_n(f)) \Rightarrow \mathcal{N}\left(0, \sqrt{\gamma_n^P(f)}\right) \forall f \in \overline{C}(E)_+ \text{ if } \frac{N}{m_N} \to 0.$ 

**Remark 5.3.** 1) This clt requires exactly the same "extra particle" condition  $\frac{N}{m_N} \to 0$  as the clt for exchangeable random variables in Theorem [2.1.](#page-10-0) 2)  $\gamma_n^P(f) = \gamma_n^W(f)$ , given in [\(4.30\)](#page-17-2), when there is no resampling and  $\gamma_n^P(f) = \gamma_n^R(f)$ , given in [\(4.34\)](#page-17-3), when there is full resampling.

We use the following theorem to prove Theorem [5.1.](#page-21-0)

<span id="page-21-1"></span>**Theorem 5.2.** Suppose  $\rho \in [0,1], N_0 \in \mathbb{N}, m_N \ge N + N^{\rho} - 1$  for all  $N \ge N_0$  and  $\{\psi_{N,k}\}_{k=1}^{m_N}$  are exchangeable random variables such that: i)  $N^{1-\rho}E[\psi_{N,1}^2] \to 0$ , and ii)  $NE[\psi_{N,1}\psi_{N,2}] \to 0$ . Then, √ 1  $\frac{1}{N}$  $\sum_{l=1}^{N}$  $\sum_{k=1} \psi_{N,k} \rightarrow^P 0.$ 

*Proof.* Define  $\mathcal{F}_{N,i} = \sigma$  $\sqrt{ }$  $\psi_{N,1},...,\psi_{N,i}, \sum^{m_N}$  $\left\{\sum_{j=i+1}^{m_N} \psi_{N,j}\right\}$  and let  $\Theta_{N,i} = \psi_{N,i} - E\left[\psi_{N,m_N}|\mathcal{F}_{N,i-1}\right]$ . Then, using the exchangeability, one has that

$$
\lim_{N \to \infty} E \left| \frac{1}{\sqrt{N}} \sum_{i=1}^{N} \Theta_{N,i} \right|^2 = \lim_{N \to \infty} \frac{1}{N} \sum_{i=1}^{N} E \left[ \Theta_{N,i}^2 \right]
$$
\n
$$
= \lim_{N \to \infty} \frac{1}{N} \sum_{i=1}^{N} E \left[ \psi_{N,i}^2 \right]
$$
\n
$$
- \lim_{N \to \infty} \frac{1}{N} \sum_{i=1}^{N} E \left[ E^2 \left[ \psi_{N,i} | \mathcal{F}_{N,i-1} \right] \right]
$$
\n
$$
\leq \lim_{N \to \infty} E \left[ \psi_{N,1}^2 \right] = 0 \text{ by i}.
$$
\n(5.14)

By exhangeability, linearity and the definition of  $\mathcal{F}_{N,i}$ , we find that

$$
E\left[\psi_{N,m_N}|\mathcal{F}_{N,i-1}\right] = \frac{1}{m_N - i + 1} \sum_{j=i}^{m_N} E\left[\psi_{N,j}|\mathcal{F}_{N,i-1}\right]
$$
\n
$$
= (m_N - i + 1)^{-1} \sum_{j=i}^{m_N} \psi_{N,j}.
$$
\n(5.15)

Therefore, it follows by Jensen's inequality that

$$
\lim_{N \to \infty} E \left| \frac{1}{\sqrt{N}} \sum_{i=1}^{N} E \left[ \psi_{N,m_N} | \mathcal{F}_{N,i-1} \right] \right|
$$
\n
$$
\leq \lim_{N \to \infty} \sum_{i=1}^{N} \sqrt{\frac{\sum_{j=i}^{m_N} E \left[ \psi_{N,j}^2 \right] + \sum_{j \neq k=i}^{m_N} E \left[ \psi_{N,j} \psi_{N,k} \right]}{N \left( m_N - i + 1 \right)^2}}
$$
\n
$$
\leq \lim_{N \to \infty} \sqrt{\frac{N}{m_N - N + 1} E \left[ \psi_{N,1}^2 \right] + N E \left[ \psi_{N,1} \psi_{N,2} \right]} = 0
$$
\n(5.16)

by i) and ii).  $\square$ 

As noted in Remark [2.5,](#page-10-1) our resampled filter can degenerate to few particles or grossly uneven weights. The following one step bounds, used to prove Theorem [5.1,](#page-21-0) ensure the risk of such system irregularity decreases exponentially in the initial number of particles.

<span id="page-22-1"></span>**Theorem 5.3.** Suppose  $n \in \mathbb{N}$ ;  $\{m_N\}_{N=1}^{\infty}$  satisfies  $m_1 \geq 2$ ,  $m_N \nearrow \infty$ ;  $h \in B(\mathbb{R}^d)$ ;  $f_V \in B(\mathbb{R})$ ; and  $g \in C_{++}(\mathbb{R}^d)$ . Then, there are  $\epsilon_n > 0$ ,  $C_n > 1$  and  $\mathbb{D}_n^N \in \sigma$   $\Big\{ \sum_{n=1}^{N} \frac{1}{n^N} \Big\}$  $k=1$  $\mathbb{N}_l^k$ ,  $l \leq n$ ) such that  $\mathbb{D}_{n+1}^N \subset \mathbb{D}_n^N$  for all  $n = 0, 1, 2...; Q^Y(\mathbb{D}_n^N) \ge 1 - 2ne^{-\epsilon_n m_N}$  for  $N \ge 1$ ; and

$$
\mathbb{N}_{l}^{k}, \max_{i \in \{1, ..., \mathbb{N}_{l}^{k}\}} \mathbb{L}_{l}^{k, i}, \mathbb{A}_{l}^{m_{N}} \leq C_{n} \ \forall k \in \{1, ..., m_{N}\}, l \in \{0, ..., n\} \ on \ \mathbb{D}_{n-1}^{N}.
$$

**Remark 5.4.** This result says that the algorithms are well behaved for at least one step on  $\mathbb{D}_n^N$ , which allows comparsion of the resampled and reduced filters on  $\mathbb{D}_{n}^{N}$ .

*Proof.* Initial Setup: Let  $\mathbb{W}_l^{k,i} = \alpha_l(\mathbb{X}_{l-1}^{k,i})$ . Since

$$
0 < \inf_{x \in E} \frac{g(Y_l - h(x))}{g(Y_l)}, \, \sup_{x \in E} \frac{g(Y_l - h(x))}{g(Y_l)} < \infty
$$

<span id="page-22-0"></span>there is a  $C = C(Y_1, ..., Y_n) > 1$  such that

$$
\frac{1}{C} \le \mathbb{W}_l^{k,i} \le C \ \forall 1 \le i \le N_{l-1}^k; 1 \le l \le n; 1 \le k \le m_N; N \ge 1. \tag{5.17}
$$

<span id="page-23-1"></span>For  $l \geq 1$ , we define  $v_C(l), \tau_C(l), \mathbb{D}_l^N$  recursively by

$$
v_C(l) = Cv_C(l-1)\tau_C(l-1), \text{ subject to } v_C(0) = 1,
$$
\n(5.18)

$$
\tau_C(l) = 2\tau_C(l-1)(C\nu_C(l)\nu_C(l-1)+1) \text{ subject to } \tau_C(0) = 1,
$$
\n(5.19)

$$
\mathbb{D}_l^N = \left\{ \frac{1}{\tau_C(l)} \le \frac{1}{m_N} \sum_{k=1}^{m_N} \mathbb{N}_l^k \le \tau_C(l) \right\} \cap \mathbb{D}_{l-1}^N \text{ subject to } \mathbb{D}_0^N = \Omega. \tag{5.20}
$$

Clearly,  $\mathbb{D}_{l}^{N} \in \mathcal{F}_{l}^{\mathbb{X} \mathbb{U} \mathbb{V}}$ . Now, recall from [\(5.2\)](#page-19-0) and the resampled algorithm that

$$
\mathbb{A}_{l+1}^{m_N} = \frac{1}{m_N} \sum_{k=1}^{m_N} \sum_{i=1}^{N_l^k} \mathbb{W}_{l+1}^{k,i} \mathbb{L}_l^{k,i}
$$
\n(5.21)

$$
\mathbb{A}_{l+1}^{m_N} \wedge \mathbb{W}_{l+1}^{k, p_{l+1}(j)} \mathbb{L}_l^{k, p_{l+1}(j)} \le \mathbb{L}_{l+1}^{k, j} \le \mathbb{A}_{l+1}^{m_N} \vee \mathbb{W}_{l+1}^{k, p_{l+1}(j)} \mathbb{L}_l^{k, p_{l+1}(j)} \tag{5.22}
$$

$$
\mathbb{N}_{l+1}^{k} = \sum_{i_{l}=1}^{\mathbb{N}_{l}^{k}} \mathbb{N}_{l+1}^{k, i_{l}} \leq \sum_{i_{1}=1}^{\mathbb{N}_{l}^{k}} \sum_{i_{2}=\mathbb{N}_{2}^{k, i_{1}-1}+1}^{\overline{\mathbb{N}}_{2}^{k, i_{1}}} \cdots \sum_{i_{l}=\mathbb{N}_{l}^{k, i_{l}-1}-1}^{\overline{\mathbb{N}}_{l}^{k, i_{l}-1}} \left[ \frac{\mathbb{L}_{l}^{k, i_{l}} \mathbb{W}_{l+1}^{k, i_{l}}}{\mathbb{A}_{l+1}^{m_{N}}} + 1 \right] \tag{5.23}
$$

<span id="page-23-0"></span>for  $j = 1, ..., N_{l+1}^k; k = 1, 2, ..., m_N$ . These imply that

$$
\frac{1}{\nu_C(l+1)} \le \mathbb{A}_{l+1}^{m_N} \le \nu_C(l+1)
$$
\n(5.24)

$$
\frac{1}{v_C(l+1)} \le \mathbb{L}_{l+1}^{k,i} \le v_C(l+1) \ \forall k \in \{1, 2, ..., m_N\}, \ i \in \{1, ..., \mathbb{N}_{l+1}^k\}
$$
 (5.25)

$$
\frac{1}{C\upsilon_C(l)} \le \mathbb{W}_{l+1}^{k,i} \mathbb{L}_l^{k,i} \le C\upsilon_C(l) \ \forall k \in \{1, 2, ..., m_N\}, \ i \in \{1, ..., \mathbb{N}_{l+1}^k\}
$$
\n(5.26)

$$
\mathbb{N}_{l+1}^k \le \prod_{i=0}^l (v_C(i+1)v_C(i)C + 1) \stackrel{\circ}{=} M_C(l+1) \ \forall k \in \{1, 2, ..., m_N\}
$$
\n(5.27)

on  $\mathbb{D}_l^N$  for all  $l = 0, 1, 2..., n$  by induction and  $(5.17)$ .

<span id="page-23-2"></span>**Base Case:**  $\{ \mathbb{N}_1^k \}$  are bounded by  $M_C(1)$  (since  $\mathbb{D}_0^N = \Omega$ ) and conditionally independent so Hoeffding's inequality applies to find

$$
Q^{Y}\left(\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\left[\mathbb{N}_{1}^{k}-\left[\frac{\mathbb{W}_{1}^{k,1}}{\mathbb{A}_{1}^{m_{N}}}\mathbf{1}_{\mathbb{H}_{1}^{k,1}}+1_{(\mathbb{H}_{1}^{k,1})^{C}}\right]\right]\right|>t\left|\mathcal{F}_{0}^{X}\vee\mathcal{F}_{1}^{V}\right)
$$
\n
$$
\leq 2\exp\left(-\frac{2m_{N}t^{2}}{M_{C}^{2}(1)}\right) \text{ a.s.},\tag{5.28}
$$

where resample set  $\mathbb{H}_1^{k,i}$  is defined in [\(5.5\)](#page-20-2). Next, by [\(5.17\)](#page-22-0), [\(5.24\)](#page-23-0), [\(5.18\)](#page-23-1), [\(5.19\)](#page-23-1) and [\(5.28\)](#page-23-2)

$$
Q^{Y}\left(\left\{\frac{1}{\tau_{C}(1)} \leq \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \mathbb{N}_{1}^{k} \leq \tau_{C}(1)\right\}\right)
$$
\n
$$
\geq Q^{Y}\left(\left\{\frac{2}{\tau_{C}(1)} \leq \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \left[\frac{\mathbb{W}_{1}^{k,1}}{\mathbb{A}_{1}^{m_{N}}} 1_{\mathbb{H}_{1}^{k,1}} + 1_{(\mathbb{H}_{1}^{k,1})^{C}}\right] \leq \frac{\tau_{C}(1)}{2}\right\})
$$
\n
$$
- E^{Y}\left[Q^{Y}\left(\left|\sum_{k=1}^{m_{N}} \left[\mathbb{N}_{1}^{k} - \left[\frac{\mathbb{W}_{1}^{k,1}}{\mathbb{A}_{1}^{m_{N}}} 1_{\mathbb{H}_{1}^{k,1}} + 1_{(\mathbb{H}_{1}^{k,1})^{C}}\right]\right]\right| > \frac{m_{N}}{\tau_{C}(1)}\left|\mathcal{F}_{0}^{\mathbb{X}} \vee \mathcal{F}_{1}^{\mathbb{V}}\right|\right]
$$
\n
$$
\geq 1 - 2 \exp\left(-\frac{2m_{N}}{M_{C}^{2}(1)\tau_{C}^{2}(1)}\right).
$$
\n(5.29)

# <span id="page-24-1"></span>Inductive Step: Suppose that

$$
Q^Y(\mathbb{D}_l^N) \ge 1 - 2l \exp\left(-\frac{2m_N}{M_C^2(l)\tau_C^2(l)}\right),\tag{5.30}
$$

which is true when  $l = 1$ , and let

$$
\rho_l^{k,i} = \frac{\mathbb{W}_{l+1}^{k,i} \mathbb{L}_l^{k,i}}{\mathbb{A}_{l+1}^{m_{N}}} 1_{\mathbb{H}_{l+1}^{k,i}} + 1_{(\mathbb{H}_{l+1}^{k,i})^C}.
$$
\n(5.31)

<span id="page-24-0"></span>Then, it follows by  $(5.26)$ ,  $(5.24)$   $(5.19)$  and  $(5.18)$  that

$$
Q^{Y}\left(\left\{\frac{1}{\tau_{C}(l+1)} \leq \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \mathbb{N}_{l+1}^{k} \leq \tau_{C}(l+1) \right\} \cap \mathbb{D}_{l}^{N}\right)
$$
\n
$$
\geq Q^{Y}\left(\left\{\frac{2}{\tau_{C}(l+1)} \leq \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \sum_{i=1}^{N_{l}^{k}} \rho_{l}^{k,i} \leq \frac{\tau_{C}(l+1)}{2} \right\} \cap \mathbb{D}_{l}^{N}\right)
$$
\n
$$
- Q^{Y}\left(\left\{\left|\frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \left[\mathbb{N}_{l+1}^{k} - \sum_{i=1}^{N_{l}^{k}} \rho_{l}^{k,i}\right] \right| > \frac{1}{\tau_{C}(l+1)}\right\} \cap \mathbb{D}_{l}^{N}\right)
$$
\n
$$
\geq Q^{Y}(\mathbb{D}_{l}^{N}) - Q^{Y}\left(\left\{\left|\frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \left[\mathbb{N}_{l+1}^{k} - \sum_{i=1}^{N_{l}^{k}} \rho_{l}^{k,i}\right] \right| > \frac{1}{\tau_{C}(l+1)}\right\} \cap \mathbb{D}_{l}^{N}\right).
$$
\n(5.32)

<span id="page-24-2"></span>However, we have by the independence of the  $U$ 's,  $(5.27)$  and Hoeffding's inequality that

$$
Q^{Y}\left(\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\left[\mathbb{N}_{l+1}^{k}-\sum_{i=1}^{\mathbb{N}_{l}^{k}}\rho_{l}^{k,i}\right]\right|>t\left|\mathcal{F}_{\infty}^{\mathbb{X},\mathbb{V}}\vee\mathcal{F}_{l}^{\mathbb{U}}\right)\right)
$$
\n
$$
\leq 2\exp\left(-\frac{2m_{N}t^{2}}{M_{C}^{2}(l+1)}\right) \quad \text{on } \mathbb{D}_{l}^{N},\tag{5.33}
$$

# M. A. Kouritzin/Branching Particle Filters 26

where resample set  $\mathbb{H}_l^{k,i}$  is defined in [\(5.5\)](#page-20-2), so by [\(5.32\)](#page-24-0), [\(5.30\)](#page-24-1) and [\(5.33\)](#page-24-2) with  $t = \frac{1}{\tau_C(l+1)}$ 

$$
Q^{Y}\left(\left\{\tau_{C}(l+1) \leq \frac{1}{m_{N}}\sum_{k=1}^{m_{N}} \mathbb{N}_{l+1}^{k} \leq \tau_{C}(l+1)\right\} \cap \mathbb{D}_{l}^{N}\right)
$$
\n
$$
\geq 1 - 2(l+1)\exp\left(-\frac{2m_{N}}{M_{C}^{2}(l+1)\tau_{C}^{2}(l+1)}\right).
$$
\n(5.34)

**Conclusion:** The result follows by induction,  $(5.20)$ ,  $(5.24)$  and  $(5.25)$ .  $\Box$ 

The proof of Theorem [5.1](#page-21-0) also relies on a coupling of our systems as well as tracking systems that run as weighted particle filters but signal the resampling events for the resampled and reduced particle filters.

#### 6. Tracking Systems

For analytical reasons, we define tracking systems corresponding to the resampled and reduced systems. These systems do not resample but do track where resampling would occur (at least initially). They are used in Appendix 2 to establish the "closeness" of the resampled and weighted filter total masses. However, they are introduced now in order that we can couple these tracking systems with the resampled and reduced systems on the same probability space.

The reduced tracking system is defined as follows:

**Initialize:**  $\underline{\mathcal{X}}_0^k = \chi^k$  and  $\underline{\mathcal{L}}_0^k = 1$  for  $k = 1, 2, ..., m_N$ ;

**Repeat:** for  $n = 0, 1, 2, ...$  do

For  $k = 1, 2, ..., m_N$  do:

$$
\underline{\hat{\mathcal{L}}}_n^k = \alpha_{n+1}(\underline{\mathcal{X}}_n^k)\underline{\mathcal{L}}_n^k
$$
\n
$$
\begin{pmatrix}\n\sigma & (1) & \widehat{\mathcal{C}}^k & (1)^k \sigma & (1)^k \sigma & (1)^k\n\end{pmatrix}
$$
\n(6.1)

$$
\underline{\mathcal{L}}_{n+1}^k = \begin{cases} \sigma_{n+1}(1), & \underline{\widehat{\mathcal{L}}}_n^k + \mathcal{V}_{n+1}^{k,1} \notin (a_n \sigma_{n+1}(1), b_n \sigma_{n+1}(1)) \\ \underline{\widehat{\mathcal{L}}}_n^k, & \underline{\widehat{\mathcal{L}}}_n^k + \mathcal{V}_{n+1}^{k,1} \in (a_n \sigma_{n+1}(1), b_n \sigma_{n+1}(1)) \end{cases} (6.2)
$$

$$
\underline{\mathcal{X}}_{n+1}^k = \mathcal{Z}_{n+1}^{k,1,\underline{\mathcal{X}}_n^k} \tag{6.3}
$$

while the *reduced tracking* system is:

**Initialize:**  $\underline{\mathbb{X}}_0^k = \chi^k$  and  $\underline{\mathbb{L}}_0^k = 1$  for  $k = 1, 2, ..., m_N$ ;

**Repeat:** for  $n = 0, 1, 2, ...$  do

For  $k = 1, 2, ..., m_N$  do:

$$
\underline{\widehat{\mathbb{L}}}_n^k = \alpha_{n+1}(\underline{\mathbb{X}}_n^k) \underline{\mathbb{L}}_n^k
$$
\n
$$
(6.4)
$$

$$
\underline{\mathbb{L}}_{n+1}^k = \begin{cases} \mathbb{A}_{n+1}, & \underline{\widehat{\mathbb{L}}}_n^k + \mathbb{V}_{n+1}^{k,1} \notin (a_n \mathbb{A}_{n+1}, b_n \mathbb{A}_{n+1}) \\ \underline{\widehat{\mathbb{L}}}_n^k, & \underline{\widehat{\mathbb{L}}}_n^k + \mathbb{V}_{n+1}^{k,1} \in (a_n \mathbb{A}_{n+1}, b_n \mathbb{A}_{n+1}) \end{cases} (6.5)
$$

$$
\underline{\mathbb{X}}_{n+1}^k = \mathbb{Z}_{n+1}^{k,1,\underline{\mathbb{X}}_n^k}.
$$
\n(6.6)

In the above algorithms,  $\{\mathbb{V}_n^{k,1}; n, k = 1, 2, ...\}$  and  $\{\mathbb{Z}_n^{k,1,x}; n, k = 1, 2, ..., x \in E\}$  are the random variables used in the resampled system while  $\{\mathcal{V}_n^{k,1}; n, k = 1, 2, ...\}$  and  $\{\mathcal{Z}_n^{k,1,x}; n, k = 1, 2, ..., x \in E\}$  are the random

variables used in the reduced system.  $\{A_n, n = 0, 1, 2, ...\}$  is also from the resampled system. Hence, the resampled and reduced tracking systems have been defined on the same probability space as the resampled and reduced particle filters respectively.

One would never implement these tracking systems. Roughly speaking, they run as weighted filters but indicate (at least initially) where resampling for the reduced and resampled particle filter would have taken place. Their importance is solely to ease the analysis by facilitating a break up of the weighted and resampled particle filters over certain resampling events. In particular, the resample sets of the tracking systems:

<span id="page-26-1"></span>
$$
\underline{\mathbb{H}}_m^k = \left\{ \frac{\underline{\widehat{\mathbb{L}}}_{m-1}^k + \mathbb{V}_m^{k,1}}{\mathbb{A}_m^{m_N}} \notin (a_{m-1}, b_{m-1}) \right\},
$$
\n(6.7)

$$
\underline{\mathcal{H}}_m^k = \left\{ \frac{\underline{\hat{\mathcal{L}}}_{m-1}^k + \mathcal{V}_m^{k,1}}{\sigma_m(1)} \notin (a_{m-1}, b_{m-1}) \right\} \tag{6.8}
$$

are important to break up the weighted and reduced particle systems into comparable pieces once we have coupled all systems together on the same probability space.

# 7. Coupling

To obtain "nearness" estimates between the resampled, tracking and reduced filters, we couple them through an infinite particle system. Suppose  $\mathbb{N}^0 = \{\emptyset\}$ ,  $\mathbb{M} = \bigcup_{n=0}^{\infty} \mathbb{N}^n$ ,  $|\kappa| = n$  if multi-index  $\kappa \in \mathbb{N}^n$  and we enlarge  $(\Omega, \mathcal{F}, Q)$  to support the following random variables:

- 1.  $\{\chi^k\}_{k=1}^{\infty}$  are independent samples from  $\pi_0$ , 2.  $\left\{Z_{\kappa}^{k;x}: \ \kappa \in \bigcup_{n=1}^{\infty} \mathbb{N}^n, k \in \mathbb{N}; x \in E\right\}$  are independent, distribution  $K(x, \cdot)$ ,
- 3.  $\{U^k_{\kappa}: \ \kappa \in \bigcup_{n=1}^{\infty} \mathbb{N}^n, k \in \mathbb{N}\}$  are independent and Uniform[0, 1],
- 4.  ${V_{\kappa}^k : \kappa \in \bigcup_{n=1}^{\infty} \mathbb{N}^n, k \in \mathbb{N}}$  are independent and zero mean with common pdf  $f_V$ ,

<span id="page-26-0"></span>which are mutually independent and independent of X, Y. Then, at time n, there is a particle  $X_{\kappa}^{k}$  corresponding to each initial particle k and multi-index  $\kappa$  with  $|\kappa| = n$  that satisfies:

$$
X_{\emptyset}^{k} = \chi^{k}, \ X_{(\kappa,i)}^{k} = Z_{(\kappa,i)}^{k;X_{\kappa}^{k}} \quad \forall \kappa \in \mathbb{M}; k, i \in \mathbb{N}.
$$
\n
$$
(7.1)
$$

<span id="page-27-1"></span> $(\mathbb{N}_{\kappa}^k, \mathbb{L}_{\kappa}^k)_{\kappa \in \mathbb{N}, k \in \mathbb{N}}, (\mathcal{N}_{\kappa}^k, \mathcal{L}_{\kappa}^k)_{\kappa \in \mathbb{M}, k \in \mathbb{N}}, (\mathbb{L}_{\kappa}^k)_{\kappa \in \mathbb{M}, k \in \mathbb{N}}$  and  $(\underline{\mathcal{L}}_{\kappa}^k)_{\kappa \in \mathbb{M}, k \in \mathbb{N}}$  then extend the notion of offspring numbers and likelihood for the finite systems to the infinite system, where

$$
(\mathbb{N}_{\emptyset}^{k}, \mathbb{L}_{\emptyset}^{k}) = (\mathcal{N}_{\emptyset}^{k}, \mathcal{L}_{\emptyset}^{k}) = (1, 1), \underbrace{\mathbb{L}_{\emptyset}^{k}} = \underline{\mathcal{L}}_{\emptyset}^{k} = 1
$$
\n
$$
(\mathcal{T} \cup \mathcal{T} \cup \mathcal{T}^{k}) \cup (\mathcal{T} \cup \mathcal{T}^{k}) \cup
$$

$$
(\mathbb{N}_{(\kappa,i)}^k, \mathbb{L}_{(\kappa,i)}^k) = \begin{cases} \left( \mathbb{L} U_{(\kappa,i)}^k, \mathbb{A}_{n+1} \right), & \frac{\mathbb{L}_{\kappa}^k + V_{(\kappa,i)}^k}{\mathbb{A}_{n+1}} \notin (a_n, b_n), i \le \mathbb{N}_{\kappa}^k\\ (1, \mathbb{L}_{\kappa}^k), & \frac{\mathbb{L}_{\kappa}^k + V_{(\kappa,i)}^k}{\mathbb{A}_{n+1}} \in (a_n, b_n), i \le \mathbb{N}_{\kappa}^k\\ (0, 0), & i > \mathbb{N}_{\kappa}^k \end{cases} \tag{7.3}
$$

$$
(\mathcal{N}_{(\kappa,i)}^k, \mathcal{L}_{(\kappa,i)}^k) = \begin{cases} (\mathcal{L}U_{(\kappa,i)}^k, \sigma_{n+1}(1)), & \frac{\widehat{\mathcal{L}}_{\kappa+1}^k(\kappa, i)}{\sigma_{n+1}(1)} \notin (a_n, b_n), i \le \mathcal{N}_{\kappa}^k\\ (1, \widehat{\mathcal{L}}_{\kappa}^k), & \frac{\widehat{\mathcal{L}}_{\kappa+1}^k(\kappa, i)}{\sigma_{n+1}(1)} \in (a_n, b_n), i \le \mathcal{N}_{\kappa}^k\\ (0, 0), & i > \mathcal{N}_{\kappa}^k \end{cases}
$$
(7.4)

$$
\underline{\mathbb{L}}_{(\kappa,i)}^{k} = \begin{cases}\n\mathbb{A}_{n+1}, & \frac{\underline{\widehat{\mathbb{L}}}_{\kappa}^{k} + V_{(\kappa,i)}^{k}}{\mathbb{A}_{n+1}} \notin (a_{n}, b_{n}), i = 1 \\
\frac{\underline{\mathbb{L}}_{\kappa}^{k}}{\mathbb{L}_{\kappa}}, & \frac{\underline{\mathbb{L}}_{\kappa} + V_{(\kappa,i)}^{k}}{\mathbb{A}_{n+1}} \in (a_{n}, b_{n}), i = 1 \\
0, & i > 1\n\end{cases} (7.5)
$$

$$
\underline{\mathcal{L}}_{(\kappa,i)}^k = \begin{cases}\n\sigma_{n+1}(1), & \frac{\underline{\widehat{\mathcal{L}}}_{\kappa}^k + V_{(\kappa,i)}^k}{\sigma_{n+1}(1)} \notin (a_n, b_n), i = 1 \\
\frac{\underline{\mathcal{L}}}_{(\kappa,i)}^k & \frac{\underline{\widehat{\mathcal{L}}}_{\kappa}^k + V_{(\kappa,i)}^k}{\sigma_{n+1}(1)} \in (a_n, b_n), i = 1 \\
0, & i > 1\n\end{cases}
$$
\n(7.6)

<span id="page-27-0"></span>for all  $k, i \in \mathbb{N}, |\kappa| = n, n = 0, 1, 2, ...$  Here,

$$
\mathbb{L}U_{(\kappa,i)}^k = \left\lfloor \frac{\widehat{\mathbb{L}}_{\kappa}^k}{\mathbb{A}_{n+1}} \right\rfloor + 1_{U_{(\kappa,i)}^k + \left\lfloor \frac{\widehat{\mathbb{L}}_{\kappa}^k}{\mathbb{A}_{n+1}} \right\rfloor \le \frac{\widehat{\mathbb{L}}_{\kappa}^k}{\mathbb{A}_{n+1}}} \tag{7.7}
$$

$$
\mathcal{L}U_{(\kappa,i)}^k = \left[\frac{\widehat{\mathcal{L}}_\kappa^k}{\sigma_{n+1}(1)}\right] + 1_{U_{(\kappa,i)}^k + \left[\frac{\widehat{\mathcal{L}}_\kappa^k}{\sigma_{n+1}(1)}\right] \le \frac{\widehat{\mathcal{L}}_\kappa^k}{\sigma_{n+1}(1)}}\tag{7.8}
$$

$$
\mathbb{A}_{n+1} = \frac{1}{m_N} \sum_{k=1}^{m_N} \sum_{\kappa: |\kappa| = n} \widehat{\mathbb{L}}_{\kappa}^k \text{ for } n = 0, 1, \dots; \mathbb{A}_0 = 1; \tag{7.9}
$$

$$
\widehat{\mathbb{L}}_{\kappa}^{k} = \alpha_{|\kappa|+1}(X_{\kappa}^{k})\mathbb{L}_{\kappa}^{k}, \quad \widehat{\mathcal{L}}_{\kappa}^{k} = \alpha_{|\kappa|+1}(X_{\kappa}^{k})\mathcal{L}_{\kappa}^{k},\tag{7.10}
$$

$$
\underline{\widehat{\mathbb{L}}}_{\kappa}^{k} = \alpha_{|\kappa|+1}(X_{\kappa}^{k})\underline{\mathbb{L}}_{\kappa}^{k} \text{ and } \underline{\widehat{\mathcal{L}}}_{\kappa}^{k} = \alpha_{|\kappa|+1}(X_{\kappa}^{k})\underline{\mathcal{L}}_{\kappa}^{k} \text{ for } \kappa \in \mathbb{M}, k \in \mathbb{N}.
$$
 (7.11)

Next, we introduce a partial order on M:  $\kappa \prec \hat{\kappa}$  if  $|\kappa| = |\hat{\kappa}|$  and  $\min\{i : \kappa_i < \hat{\kappa}_i\} < \min\{i : \hat{\kappa}_i < \kappa_i\}$ . To make room for live particles from all finite systems, we let

$$
N_{\kappa}^{k} = \mathbb{N}_{\kappa}^{k} \vee \mathcal{N}_{\kappa}^{k} \vee 1 \quad \forall k \in \mathbb{N}, \kappa \in \mathbb{M}
$$
\n
$$
(7.12)
$$

and define the subset of *alive multi-indices*  $\mathbb{M}^A$  by  $\kappa \in \mathbb{M}^A$  if  $\kappa \in \mathbb{M}$  and either

<span id="page-27-2"></span>
$$
\kappa = \emptyset \text{ or } \kappa = (\kappa_1, ..., \kappa_n) \text{ with } \kappa_l \in \{1, ..., N^k_{(\kappa_1, ..., \kappa_{l-1})}\} \ \forall l = 1, ..., n,
$$
\n(7.13)

<span id="page-28-0"></span>so particles  $X_{(\kappa_1,\ldots,\kappa_n)}^k$  with  $(n \geq 1 \text{ and})$  some  $\kappa_l > N_{(\kappa_1,\ldots,\kappa_{l-1})}^k$  are not in any finite system. To recover the finite systems, we drop explicit reference to the ancesteral chain and set:

$$
X_n^{k,j} = X_\kappa^k, \ U_n^{k,j} = U_\kappa^k, \ V_n^{k,j} = V_\kappa^k, \ Z_n^{k,j,x} = Z_\kappa^{k,x}, \tag{7.14}
$$

$$
\mathcal{K}_n^{k,j} = \mathcal{L}_\kappa^k, \ \mathbb{K}_n^{k,j} = \mathbb{L}_\kappa^k, \ \hat{\mathcal{K}}_n^{k,j} = \hat{\mathcal{L}}_\kappa^k, \ \hat{\mathbb{K}}_n^{k,j} = \hat{\mathbb{L}}_\kappa^k,\tag{7.15}
$$

$$
N_n^{k,j} = N_\kappa^k, \overline{N}_l^{k,i} = \sum_{m=1}^i N_l^{k,m} \text{ and } N_l^k = \sum_{m=1}^{N_{l-1}^k} N_l^{k,m} \text{ with } N_0^k = 1,
$$
 (7.16)

where  $\kappa$  is the unique alive multi-index such that  $|\kappa| = n$  and

$$
j = \eta(\kappa) \stackrel{\circ}{=} \# \{ \widehat{\kappa} \in \mathbb{M}^A : \widehat{\kappa} \prec \kappa \} + 1. \tag{7.17}
$$

(Many  $K, K$  could be zero.) For the tracking systems, we define

$$
\underline{\mathcal{K}}_n^k = \underline{\mathcal{L}}_\kappa^k, \underline{\mathbb{K}}_n^k = \underline{\mathbb{L}}_\kappa^k, \ \underline{\hat{\mathcal{K}}}_n^k = \underline{\hat{\mathcal{L}}}_\kappa^k, \ \underline{\hat{\mathbb{K}}}_n^k = \underline{\hat{\mathbb{L}}}_\kappa^k,\tag{7.18}
$$

<span id="page-28-3"></span>for  $\kappa = (1, 1, ..., 1)$  with  $|\kappa| = n$ . Now, it follows by  $(7.9)$ ,  $(7.3)$ ,  $(7.12)$ ,  $(7.10)$ ,  $(7.1)$  and  $(7.14-7.16)$  that

<span id="page-28-2"></span>
$$
\mathbb{A}_{n+1} = \frac{1}{m_N} \sum_{k=1}^{m_N} \sum_{\kappa: |\kappa| = n} \widehat{\mathbb{L}}_{\kappa}^k = \frac{1}{m_N} \sum_{k=1}^{m_N} \sum_{j=1}^{N_n^k} \widehat{\mathbb{K}}_{n}^{k,j} \text{ and } (7.19)
$$

$$
X_n^{k,j} = Z_n^{k,j,X_{n-1}^{k,i}} \text{ for } j \in \{\overline{N}_n^{k,i-1} + 1, ..., \overline{N}_n^{k,i}\}.
$$
 (7.20)

<span id="page-28-1"></span>For convenience, let  $\mathbb{I}_n^k = \{i : \mathbb{K}_n^{k,i} \neq 0\}$  and  $\mathcal{I}_n^k = \{i : \mathcal{K}_n^{k,i} \neq 0\}$  be the resampled and reduced particles at time *n* that started from the  $k^{\text{th}}$  initial particle and  $|\mathbb{I}_n^k|$  denote the cardinality of  $\mathbb{I}_n^k$ . Redefine the resample and non-resample sets (previously defined in [\(5.5,5.6,](#page-20-2)[6.7,](#page-26-1) [6.8\)](#page-26-1))

$$
\mathbb{R}_{m}^{k,i} = \left\{ \frac{\widehat{\mathbb{K}}_{m-1}^{k,i} + V_{m}^{k,i}}{\mathbb{A}_{m}^{m_{N}}} \notin (a_{m-1}, b_{m-1}), i \in \mathbb{I}_{m-1}^{k} \right\},
$$
\n(7.21)

$$
\mathbb{S}_m^{k,i} = \{ i \in \mathbb{I}_{m-1}^k \} \setminus \mathbb{R}_m^{k,i},\tag{7.22}
$$

$$
\mathcal{R}_{m}^{k,i} = \left\{ \frac{\hat{\mathcal{K}}_{m-1}^{k,i} + V_{m}^{k,i}}{\sigma_{m}(1)} \notin (a_{m-1}, b_{m-1}), i \in \mathcal{I}_{m-1}^{k} \right\},
$$
\n(7.23)

$$
\mathcal{S}_m^{k,i} = \{ i \in \mathcal{I}_{m-1}^k \} \setminus \mathcal{R}_m^{k,i},
$$
\n
$$
\left( \widehat{\mathfrak{S}}_m^{k} \right)_{k=1}^k \tag{7.24}
$$

$$
\underline{\mathbb{R}}_m^k = \left\{ \frac{\underline{\hat{\mathbb{K}}}_{m-1}^k + V_m^{k,1}}{\mathbb{A}_m^{m_N}} \notin (a_{m-1}, b_{m-1}) \right\},\tag{7.25}
$$

$$
\underline{\mathcal{R}}_m^k = \left\{ \frac{\underline{\hat{\mathcal{K}}}_{m-1}^k + V_m^{k,1}}{\sigma_m(1)} \notin (a_{m-1}, b_{m-1}) \right\}.
$$
\n(7.26)

The following combinations of resample and non-resample events will be useful in comparing our resampled particle filter total mass to the weighted total mass in Appendix 2:

$$
\mathbb{R}\mathbb{S}\mathbb{I}_{l,n}^{k,i_{l-1},i_{l},...,i_{n}} = \mathbb{R}_{l}^{k,i_{l-1}} \cap \mathbb{S}_{l+1}^{k,i_{l}} \cap \cdots \cap \mathbb{S}_{n}^{k,i_{n-1}} \cap \{i_{n} \in \mathbb{I}_{n}^{k}\},
$$
\n(7.27)

$$
\mathcal{RST}_{l,n}^{k,i_{l-1},i_l,...,i_n} = \mathcal{R}_l^{k,i_{l-1}} \cap \mathcal{S}_{l+1}^{k,i_l} \cap \dots \cap \mathcal{S}_n^{k,i_{n-1}} \cap \{i_n \in \mathcal{I}_n^k\}.
$$
 (7.28)

Our coupling of the finite systems on common probability space  $(\Omega, \mathcal{F}, Q^Y)$  is complete. We use this coupling to transfer the bounds of Theorem [5.3](#page-22-1) to the infinite particle system, to prove Theorem [5.1](#page-21-0) and to ease notation about [\(9.14\)](#page-46-0) of Appendix 2. For these uses, we need the following result.

<span id="page-29-3"></span><span id="page-29-0"></span>**Theorem 7.1.** Suppose  $\mathbb{B}_n^k$ ,  $\mathcal{B}_n^k$  are the resampled, reduced Markov branching processes defined in [\(5.1\)](#page-19-1),  $(4.1)$  and  $\mathbb{N}_{n}^{k}$ ,  $\mathcal{N}_{n}^{k}$  are the corresponding particle numbers. Then,

$$
\left\{ \mathbb{N}_{n}^{k}, \mathbb{B}_{n}^{k} \right\}_{1 \leq k \leq m_{N}, n \in \mathbb{N}_{0}, N \in \mathbb{N}} \stackrel{D}{=} \left\{ \left| \mathbb{I}_{n}^{k} \right|, \sum_{i=1}^{N_{n}^{k}} \mathbb{K}_{n}^{k, i} \delta_{X_{n}^{k, i}} \right\}_{1 \leq k \leq m_{N}, n \in \mathbb{N}_{0}, N \in \mathbb{N}} \tag{7.29}
$$

$$
\left\{\mathcal{N}_{n}^{k},\mathcal{B}_{n}^{k}\right\}_{1\leq k\leq m_{N},n\in\mathbb{N}_{0},N\in\mathbb{N}}\equiv\left\{\left|\mathcal{I}_{n}^{k}\right|,\sum_{i=1}^{N_{n}^{k}}\mathcal{K}_{n}^{k,i}\delta_{X_{n}^{k,i}}\right\}_{1\leq k\leq m_{N},n\in\mathbb{N}_{0},N\in\mathbb{N}}\tag{7.30}
$$

<span id="page-29-1"></span>and

$$
\sum_{i_{l-1}=1}^{N_{l-1}^{k}} \sum_{i_{l}=\overline{N}_{l}^{k,i_{l-1}}}^{N_{l}^{k,i_{l-1}}} W_{l+1,n+1}^{k,i_{l},...,i_{n}} \sigma_{l}(1) 1_{\mathcal{RST}_{l,n}^{k,i_{l-1},i_{l},...,i_{n}}} \left\}_{l,k,n,N} \sum_{l,k,n,N} \sum_{i_{l-1}=1}^{N_{l-1}^{k}} \sum_{i_{l-1}=1}^{N_{l}^{k,i_{l-1}-1}^{k,i_{l-1}}} \mathcal{W}_{l+1,n+1}^{k,i_{l},...,i_{n}} \sigma_{l}(1) 1_{\mathcal{H}_{l}^{k,i_{l-1}}} (\mathcal{H}_{l+1}^{k,i_{l}})^{C} \cdots (\mathcal{H}_{n}^{k,i_{n-1}})^{C} \right\}_{l,k,n,N}
$$
\n(7.31)

where

$$
W_{l+1,n+1}^{k,i_l,\dots,i_n} = W_{n+1}^{k,i_n} \cdots W_{l+2}^{k,i_{l+1}} W_{l+1}^{k,i_l}
$$
\n
$$
(7.32)
$$

$$
\mathcal{W}_{l+1,n+1}^{k,i_l,\dots,i_n} = \mathcal{W}_{n+1}^{k,i_n} \cdots \mathcal{W}_{l+2}^{k,i_{l+1}} \mathcal{W}_{l+1}^{k,i_l}
$$
\n(7.33)

<span id="page-29-4"></span>
$$
W_l^{k,i} = \alpha_l(X_{l-1}^{k,i}) \text{ and } \mathcal{W}_l^{k,i} = \alpha_l(X_{l-1}^{k,i}).
$$
\n(7.34)

<span id="page-29-2"></span>Moreover, there are  $\epsilon_n, C_n > 0$  and  $\mathbb{D}_n^N \in \sigma \bigg\{ \sum_{n=1}^{m_N} \mathcal{O}(\epsilon_n^N) \bigg\}$  $k=1$  $|\mathbb{I}_l^k|, l \leq n$  $\mathcal{L}$ , such that  $\mathbb{D}_{n+1}^N \subset \mathbb{D}_n^N$ ,

$$
Q^Y(\mathbb{D}_n^N) \ge 1 - 2ne^{-\epsilon_n m_N} \tag{7.35}
$$

$$
\max_{i \in \{1,\ldots,N_l^k\}} \mathbb{K}_l^{k,i} \vee |\mathbb{I}_l^k| \vee \mathbb{A}_l \le C_n \ \forall 1 \le k \le m_N; 0 \le l \le n \ on \ \mathbb{D}_{n-1}^N,\tag{7.36}
$$

$$
\max_{i \in \{1, \dots, N^k_l\}} \mathcal{K}^{k,i}_l \vee |\mathcal{I}^k_l| \le C_n \ \forall 1 \le k \le m_N; 0 \le l \le n \ on \ \Omega \tag{7.37}
$$

for all  $n = 0, 1, 2...$ 

**Note:** For notational simplicity, we take  $\mathbb{D}_{-1}^{N} = \Omega$  in the sequel.

*Proof.* Suppose (temporarily) the alive multi-indices  $\mathbb{M}^A$  were  $\kappa \in \mathbb{M}^A$  if  $\kappa \in \mathbb{M}$  and either

$$
\kappa = \emptyset \text{ or } \kappa = (\kappa_1, ..., \kappa_n) \text{ with } \kappa_l \in \{1, ..., \mathbb{N}^k_{(\kappa_1, ..., \kappa_{l-1})}\} \ \forall l = 1, ..., n,
$$
\n(7.38)

replacing  $N_{(\kappa_1,\ldots,\kappa_{l-1})}^k$  with  $\mathbb{N}_{(\kappa_1,\ldots,\kappa_{l-1})}^k$ , and we defined

$$
\mathbb{X}_n^{k,j} = X_\kappa^k, \mathbb{U}_n^{k,j} = U_\kappa^k, \mathbb{V}_n^{k,j} = V_\kappa^k,
$$
\n(7.39)

$$
\mathbb{Z}_{n}^{k,j,x} = Z_{\kappa}^{k;x}, \ \mathbb{L}_{n}^{k,j} = \mathbb{L}_{\kappa}^{k}, \ \hat{\mathbb{L}}_{n}^{k,j} = \hat{\mathbb{L}}_{\kappa}^{k}, \tag{7.40}
$$

<span id="page-30-1"></span>
$$
\mathbb{N}_n^{k,j} = \mathbb{N}_\kappa^k, \ \overline{\mathbb{N}}_l^{k,i} = \sum_{m=1}^i \mathbb{N}_l^{k,m} \text{ and } \mathbb{N}_l^k = \sum_{m=1}^{N_{l-1}^*} \mathbb{N}_l^{k,m} \text{ with } \mathbb{N}_0^k = 1,
$$
 (7.41)

where  $\kappa$  is the unique alive multi-index such that  $|\kappa| = n$  and

$$
j = \eta(\kappa) \stackrel{\circ}{=} \# \{ \widehat{\kappa} \in \mathbb{M}^A : \widehat{\kappa} \prec \kappa \} + 1. \tag{7.42}
$$

Then, the resampled particle system algorithm is recovered by [\(7.1](#page-26-0)[-7.3\)](#page-27-1), [\(7.9\)](#page-27-0), [\(7.10\)](#page-27-0) with these definitions. Moreover, the process distribution of the resampled estimates  $\mathbb{B}_n^k$  and particle numbers  $\mathbb{N}_n^k$  do not change if we select from the (independent)  $\{Z_{\kappa}^{k,x}\}\$ ,  $\{U_{\kappa}^{k}\}$  and  $\{V_{\kappa}^{k}\}$  differently nor if we add in zero weights and zero offspring numbers. Therefore, examining the equations [\(7.1](#page-26-0)[-7.42\)](#page-30-1) and concentrating on this resampled particle algorithm, we find

$$
\left\{ \mathbb{N}_{n}^{k}, \mathbb{B}_{n}^{k} \right\}_{k,n,N} \stackrel{D}{=} \left\{ \left| \mathbb{I}_{n}^{k} \right|, \sum_{i=1}^{N_{n}^{k}} \mathbb{K}_{n}^{k,i} \delta_{X_{n}^{k,i}} \right\}_{k,n,N} . \tag{7.43}
$$

 $(7.30,7.31)$  $(7.30,7.31)$  are handled similarly.  $(7.35-7.37)$  now follow from Lemma [4.1,](#page-18-2) Theorem [5.3](#page-22-1) and  $(7.29,7.30)$ .  $\Box$ For notational convenience, we define the (exchangeable random) signed measures  ${B_n^{N,k}}_{k=1}^m$  and the parent operators (with respect to  $\kappa$  and  $\eta$  defined in [\(7.42\)](#page-30-1)) by:

<span id="page-30-0"></span>
$$
B_n^{N,k} = 1_{\mathbb{D}_{n-1}^N} \sum_{i=1}^{N_n^k} K_n^{k,i} \delta_{X_n^{k,i}} \text{ with } K_n^{k,i} = \mathbb{K}_n^{k,i} - \mathcal{K}_n^{k,i}
$$
 (7.44)

$$
p_l(i) = \eta(\kappa_1, ..., \kappa_{l-1})
$$
 when  $i = \eta(\kappa_1, ..., \kappa_l)$  (7.45)

<span id="page-30-3"></span>
$$
p_{l,m}(i) = \begin{cases} p_l(\cdots p_{m-1}(p_m(i))) & \text{for } l \le m \\ i & \text{for } l > m \end{cases}
$$
 (7.46)

(so  $i \in {\{\overline{N}_l^{k, p_l(i) - 1} + 1, ..., \overline{N}_l^{k, p_l(i)}}$  $\{u_i^{\alpha,\mu_i(\nu)}\}$ . Finally, by the argument in  $(5.17)$  there is a  $c > 1$  so that

$$
W_l^{k,i} \le c \,\forall i, k, l \in \mathbb{N}.\tag{7.47}
$$

Now that we have redefined the algorithms on the same (infinite particle system and) probability space  $(\Omega, \mathcal{F}, Q^Y)$  (for each fixed Y), we can compare their particle systems.

<span id="page-30-2"></span>**Theorem 7.2.** Suppose  $p \in \mathbb{N}$  as well as the conditions and setting of Theorem [7.1](#page-29-3) with all algorithms

<span id="page-31-0"></span>defined on  $(\Omega, \mathcal{F}, Q^Y)$ . Then, there are  $C_n = C_n^{p,Y} > 0$  such that

$$
E^{Y}\left[|\mathbb{A}_{n}^{m_{N}} - \sigma_{n}(1)|^{p}1_{\mathbb{D}_{n-1}^{N}}\right] \leq C_{n}m_{N}^{-\frac{p}{2}},
$$
\n
$$
\left[|\mathbb{I}_{n} - m_{N}\right] \qquad \left| \begin{array}{c} p \\ 1 \end{array} \right] \tag{7.48}
$$

$$
E^{Y} \left[ \left| \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \sum_{i \in \mathbb{I}_{n}^{k}} |\mathbb{K}_{n}^{k,i} - \mathcal{K}_{n}^{k,i}| \right|^{2} \mathbf{1}_{\mathbb{D}_{n-1}^{N}} \right] \leq C_{n} m_{N}^{-\frac{p}{2}} \tag{7.49}
$$

$$
E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{n}^{k}}|\widehat{\mathbb{K}}_{n}^{k,i}-\widehat{\mathcal{K}}_{n}^{k,i}|\right|^{p}1_{\mathbb{D}_{n-1}^{N}}\right] \leq C_{n}m_{N}^{-\frac{p}{2}} \quad \text{and} \tag{7.50}
$$

$$
E^{Y} \left[ \left| \frac{1}{m_N} \sum_{k=1}^{m_N} \left| \widehat{\underline{\mathbb{K}}}^k_n - \widehat{\underline{\mathbb{K}}}^k_n \right| \right|^p 1_{\mathbb{D}_{n-1}^N} \right] \leq C_n m_N^{-\frac{p}{2}} \tag{7.51}
$$

for all  $m_N = p + 1, p + 2, ...$  and  $n = 1, 2...,$  where  $\mathbb{D}_{n-1}^N$  is as in Theorem [7.1.](#page-29-3)

<span id="page-31-1"></span>As these are bounds in N, we highlighted the previously-suppressed N-dependence in  $\mathbb{A}_n^{m_N}$ . The following lemma is used (with induction) to prove [\(7.50\)](#page-31-0) implies [\(7.48\)](#page-31-0) in Theorem [7.2.](#page-30-2)

# **Lemma 7.1.** Suppose  $n \in \mathbb{N}_0$  and  $E^Y\left[|\mathbb{A}_l^{m_N} - \sigma_l(1)|^p 1_{\mathbb{D}_{l-1}^N}\right]$  $\Big] \stackrel{N}{\ll} m_N^{-\frac{p}{2}}$  for all  $l \leq n$ . Then,

$$
E^{Y} [\vert A_{n+1}^{m_N} - \sigma_{n+1}(1) \vert^{p} 1_{\mathbb{D}_{n}^{N}}] \ll \sum_{j=1}^{N} E^{Y} \left[ \left| \frac{1}{m_N} \sum_{k=1}^{m_N} \vert \hat{\underline{K}}_{j}^{k} - \hat{\underline{\mathbb{K}}}_{j}^{k} \vert \right|^{p} 1_{\mathbb{D}_{j-1}^{N}} \right]
$$
  
+ 
$$
m_N^{-\frac{p}{2}} + \sum_{j=1}^{n-1} E^{Y} \left[ \left| \frac{1}{m_N} \sum_{k=1}^{m_N} \sum_{i \in \mathbb{I}_{j}^{k} \cup \mathcal{I}_{j}^{k}} \vert \hat{\mathbb{K}}_{j}^{k,i} - \hat{\mathbb{K}}_{j}^{k,i} \vert \right|^{p} 1_{\mathbb{D}_{j-1}^{N}} \right].
$$

<span id="page-31-2"></span>The proof of this lemma is involved and hence delayed to Appendix 2. **Proof of Theorem [7.2.](#page-30-2) Set Up:** Using the independence of the  $V$ 's, letting

$$
\mathcal{G}_k^l = \sigma\{V_m^{j,i} : i, j \in \mathbb{N}, m \neq l\} \vee \sigma\{V_l^{j,i} : j \leq k, i \in \mathbb{N}\} \vee \mathcal{F}_\infty^{UZ},\tag{7.52}
$$

<span id="page-32-1"></span>noting the boundedness of  $f_V$  and considering  $-\infty \le a_l \le b_l \le \infty$ , one has by [\(7.21,7.23,7.25\)](#page-28-1) that

$$
E^{Y}\left[\sum_{i\in I_{l-1}^{k}\cup I_{\mathbb{R}^{k,i}}\cup I_{\mathbb{R}^{k,i}}^{k,i}}|\mathcal{G}_{k-1}^{l}\right]
$$
(7.53)  
\n
$$
\leq \sum_{i\in I_{l-1}^{k}\cup I_{\mathbb{R}^{1}}\cup I_{\mathbb{R}^{1}}^{k,i}}\left|\left|\int_{a_{l-1}\mathbb{A}_{l}^{m_{N}}-\widetilde{\mathbb{R}}_{l-1}^{k,i}}^{a_{l-1}\sigma_{l}(1)-\widetilde{K}_{l-1}^{k,i}}f_{V}(v)dv\right| + \left|\int_{b_{l-1}\mathbb{A}_{l}^{m_{N}}-\widetilde{\mathbb{R}}_{l-1}^{k,i}}^{b_{l-1}\sigma_{l}(1)-\widetilde{K}_{l-1}^{k,i}}f_{V}(v)dv\right|\right]
$$
  
\n
$$
\ll 1_{(\mathbb{D}_{l-1}^{N})C} + |\mathbb{A}_{l}^{m_{N}}-\sigma_{l}(1)|1_{\mathbb{D}_{l-1}^{N}} + \sum_{i\in I_{l-1}^{k}\cup I_{\mathbb{R}^{k,i}}^{k,i}}|\widehat{\mathcal{K}}_{l-1}^{k,i}-\widetilde{\mathbb{R}}_{l-1}^{k,i}|1_{\mathbb{D}_{l-1}^{N}}
$$
  
\n
$$
E^{Y}\left[\sum_{i\in I_{l-1}^{k}\cup I_{\mathbb{R}^{1}}^{k,i}\cup I_{\mathbb{R}^{k,i}}^{k,i}}|g_{k-1}^{l}\right]
$$
(7.54)  
\n
$$
\ll 1_{(\mathbb{D}_{l-1}^{N})C} + |\mathbb{A}_{l}^{m_{N}}-\sigma_{l}(1)|1_{\mathbb{D}_{l-1}^{N}} + \sum_{i\in I_{l-1}^{k}\cup I_{l-1}^{k}}|\widehat{\mathcal{K}}_{l-1}^{k,i}-\widehat{\mathbb{R}}_{l-1}^{k,i}|1_{\mathbb{D}_{l-1}^{N}}
$$
  
\n
$$
E^{Y}\left[1_{\mathbb{E}_{l}^{k}\triangle\mathcal{R}_{l}^{k}}|\mathcal{G}_{0}^{l}\right]
$$
(7.55)  
\n
$$
\ll 1_{(\mathbb{D}_{l-1}^{N})C} +
$$

$$
\overset{N}{\ll} 1_{(\mathbb{D}_{l-1}^N)^C} + |\mathbb{A}_l^{m_N} - \sigma_l(1)| 1_{\mathbb{D}_{l-1}^N} + |\hat{\underline{\mathcal{K}}}^k_{l-1} - \hat{\underline{\mathbb{K}}}^k_{l-1}| 1_{\mathbb{D}_{l-1}^N}
$$

<span id="page-32-0"></span>almost surely for all  $l = 1, 2, ...$  Now, set  $S_l^{k,j} = \mathbb{S}_l^{k,j} \cap S_l^{k,j}$ , with  $\mathbb{S}_l^{k,j}$ ,  $S_l^{k,j}$  defined in [\(7.22,7.24\)](#page-28-1). If  $i \in \mathbb{I}_{n}^{k} \cup \mathcal{I}_{n}^{k}$ , then either  $i \in \mathbb{I}_{n}^{k} \triangle \mathcal{I}_{n}^{k}$  so there is a time  $l \geq 1$  when only one algorithm ancestor was resampled or  $i \in \mathbb{I}_n^k \mathcal{I}_n^k$  so the resampled and reduced particles have the same ancestorial chains. Hence, by [\(7.45,7.46\)](#page-30-0)

$$
|\mathbb{K}_{n}^{k,i} - \mathcal{K}_{n}^{k,i}| \leq \sum_{l=1}^{n} \left| \left\{ |\mathbb{K}_{n}^{k,i}| + |\mathcal{K}_{n}^{k,i}| \right\} 1_{\mathbb{R}_{l}^{k,p_{l,n}(i)} \Delta \mathcal{R}_{l}^{k,p_{l,n}(i)}} + \right. \tag{7.57}
$$
\n
$$
1_{S_{n}^{k,p_{n,n}(i)} S_{n-1}^{k,p_{n-1,n}(i)} \cdots S_{l+1}^{k,p_{l+1,n}(i)}} \left| \prod_{j=l+1}^{n} W_{j}^{k,p_{j,n}(i)} [\mathbb{A}_{l}^{m_{N}} - \sigma_{l}(1)] \right| 1_{\mathbb{R}_{l}^{k,p_{l,n}(i)} \mathcal{R}_{l}^{k,p_{l,n}(i)}} \right|,
$$
\n
$$
|\widehat{\mathbb{K}}_{n}^{k,i} - \widehat{\mathcal{K}}_{n}^{k,i}| \leq \sum_{l=1}^{n} \left| \left\{ |\widehat{\mathbb{K}}_{n}^{k,i}| + |\widehat{\mathcal{K}}_{n}^{k,i}| \right\} 1_{\mathbb{R}_{l}^{k,p_{l,n}(i)} \Delta \mathcal{R}_{l}^{k,p_{l,n}(i)}} + \right. \tag{7.58}
$$
\n
$$
1_{S_{n}^{k,p_{n,n}(i)} S_{n-1}^{k,p_{n-1,n}(i)} \cdots S_{l+1}^{k,p_{l+1,n}(i)}} \left| \prod_{j=l+1}^{n+1} W_{j}^{k,p_{j,n}(i)} [\mathbb{A}_{l}^{m_{N}} - \sigma_{l}(1)] \right| 1_{\mathbb{R}_{l}^{k,p_{l,n}(i)} \mathcal{R}_{l}^{k,p_{l,n}(i)}}.
$$

<span id="page-33-1"></span>For the tracking systems, we let  $\underline{S}_l^k = (\underline{\mathbb{R}}_l^k)^C \cap (\underline{\mathcal{R}}_l^k)^C$  and find by  $(7.5, 7.6, 7.11, 7.18)$  that

$$
|\underline{\widehat{\mathbb{K}}}_n^k - \underline{\widehat{\mathcal{K}}}_n^k| \leq \sum_{l=1}^n \left\{ |\underline{\widehat{\mathbb{K}}}_n^k| + |\underline{\widehat{\mathcal{K}}}_n^k| \right\} 1_{\underline{\mathbb{R}}_l^k \triangle \underline{\mathcal{R}}_l^k} + \sum_{l=1}^n 1_{\underline{S}_n^k \underline{S}_{n-1}^k \cdots \underline{S}_{l+1}^k} \left| \prod_{j=l+1}^{n+1} W_j^{k,1} [\mathbb{A}_l^{m_N} - \sigma_l(1)] \right| 1_{\underline{\mathbb{R}}_l^k \underline{\mathcal{R}}_l^k}.
$$
\n(7.59)

<span id="page-33-2"></span>**Base Case:** Clearly, [\(7.48-7.51\)](#page-31-0) hold with  $n = 0$  and  $C_0 = 0$ , even though this trivial case is not claimed in the theorem statement. Inductive Step: Suppose

$$
E^{Y} \left[ |\mathbb{A}_{l}^{m_{N}} - \sigma_{l}(1)|^{p} 1_{\mathbb{D}_{l-1}^{N}} \right] \ll m_{N}^{-\frac{p}{2}} \tag{7.60}
$$

$$
E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{l}^{k}\cup\mathcal{I}_{l}^{k}}|\mathbb{K}_{l}^{k,i}-\mathcal{K}_{l}^{k,i}|\right|^{p}1_{\mathbb{D}_{l-1}^{N}}\right] \stackrel{N}{\ll}m_{N}^{-\frac{p}{2}}\tag{7.61}
$$

$$
E^{Y} \left[ \left| \frac{1}{m_N} \sum_{k=1}^{m_N} \sum_{i \in \mathbb{I}_l^k \cup \mathcal{I}_l^k} |\widehat{\mathbb{K}}_l^{k,i} - \widehat{\mathcal{K}}_l^{k,i}| \right|^p 1_{\mathbb{D}_{l-1}^N} \right] \stackrel{N}{\ll} m_N^{-\frac{p}{2}} \tag{7.62}
$$

and

$$
E^{Y} \left[ \left| \frac{1}{m_N} \sum_{k=1}^{m_N} \left| \widehat{\underline{\mathbb{K}}}^k_l - \widehat{\underline{\mathbb{K}}}^k_l \right| \right|^p 1_{\mathbb{D}_{l-1}^N} \right] \stackrel{N}{\ll} m_N^{-\frac{p}{2}} \tag{7.63}
$$

hold for all  $l \leq n$ , which are true when  $n = 0$ . Then, it follows by Lemma [7.1](#page-31-1) that

<span id="page-33-0"></span>
$$
E^{Y} \left[ |\mathbb{A}_{l}^{m_{N}} - \sigma_{l}(1)|^{p} 1_{\mathbb{D}_{l-1}^{N}} \right] \ll m_{N}^{-\frac{p}{2}} \quad \forall \ l \leq n+1. \tag{7.64}
$$

<span id="page-34-0"></span>Recalling [\(7.58\)](#page-32-0), noting [\(7.36\)](#page-29-2), [\(7.37\)](#page-29-2), [\(7.47\)](#page-30-3), [\(7.64\)](#page-33-0) and using exchangeability, one finds that

$$
E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{n+1}^{k}}|\widehat{\mathbf{K}}_{n+1}^{k,i}-\widehat{\mathbf{K}}_{n+1}^{k,i}|\right|^{p}\mathbf{1}_{\mathbb{D}_{n}^{N}}\right]
$$
\n
$$
\ll E^{Y}\left[\sum_{j=1}^{n+1}\left\{|A_{j}-\sigma_{j}(1)|+\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{n+1}^{k}}\mathbf{1}_{\mathbb{R}_{j}^{k,p_{j,n+1}(i)}\Delta\mathcal{R}_{j}^{k,p_{j,n+1}(i)}}\right\|^{p}\mathbf{1}_{\mathbb{D}_{n}^{N}}\right]
$$
\n
$$
\ll E^{Y}\left[\left|\sum_{j=1}^{n+1}\left\{|B_{j}-\sigma_{j}(1)|^{p}\mathbf{1}_{\mathbb{D}_{j-1}^{N}}\right\|+E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{n+1}^{k}}\mathbf{1}_{\mathbb{R}_{j}^{k,i}\Delta\mathcal{R}_{j}^{k,i}}\right|^{p}\mathbf{1}_{\mathbb{D}_{j-1}^{N}}\right]\right\}
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}}+\sum_{k_{1}\neq k_{2}\neq\cdots\neq k_{q}}\sum_{j=1}^{n+1}\frac{E^{Y}\left[\sum_{i_{1}}\mathbf{1}_{\mathbb{R}\mathcal{R}_{j}^{k_{1},i_{1}}}\sum_{i_{2}}\mathbf{1}_{\mathbb{R}\mathcal{R}_{j}^{k_{2},i_{2}}}\cdots\sum_{i_{q}}\mathbf{1}_{\mathbb{R}\mathcal{R}_{j}^{k_{q},i_{q}}\mathbb{1}_{\mathbb{D}_{j-1}^{N}}\right]}{\sum_{q\leq p}^{N}}\right]
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}}+\sum_{q=1}^{p}\sum_{j=1}^{n+1}\frac{E^{Y}\left[\sum_{i_{1}\in\mathbb{I}_{j-1}^{n}}\mathbf{1}_{\mathbb{R}\mathcal{R}_{j}^{1,i_{1}}}\sum_{i_{2}\in\mathbb{I
$$

where

$$
\mathbb{I}\mathcal{I}_{j-1}^q = \mathbb{I}_{j-1}^q \cup \mathcal{I}_{j-1}^q \text{ and } \mathbb{R}\mathcal{R}_j^{k,i} = \mathbb{R}_j^{k,i} \triangle \mathcal{R}_j^{k,i}.
$$
 (7.66)

<span id="page-34-1"></span>In exactly the same way, we also get from [\(7.57\)](#page-32-0), [\(7.36,7.37\)](#page-29-2), [\(7.47\)](#page-30-3) and [\(7.64\)](#page-33-0)

$$
E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{n+1}^{k}}|\mathbb{K}_{n+1}^{k,i}-\mathcal{K}_{n+1}^{k,i}|\right|^{p}1_{\mathbb{D}_{n}^{N}}\right]
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}}+\sum_{q=1}^{p}\sum_{j=1}^{n+1}\frac{E^{Y}\left[\sum_{i_{1}\in\mathbb{I}_{j-1}^{T}}1_{\mathbb{R}\mathcal{R}_{j}^{1,i_{1}}}\sum_{i_{2}\in\mathbb{I}_{j-1}^{T}}1_{\mathbb{R}\mathcal{R}_{j}^{2,i_{2}}}\cdots\sum_{i_{q}\in\mathbb{I}_{j-1}^{q}}1_{\mathbb{R}\mathcal{R}_{j}^{q,i_{q}}}\mathbb{1}_{\mathbb{D}_{j-1}^{N}}\right]
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}}+\sum_{q=1}^{p}\sum_{j=1}^{n+1}\frac{E^{Y}\left[\sum_{i_{1}\in\mathbb{I}_{j-1}^{T}}1_{\mathbb{R}\mathcal{R}_{j}^{1,i_{1}}}\sum_{i_{2}\in\mathbb{I}_{j-1}^{T}}1_{\mathbb{R}\mathcal{R}_{j}^{2,i_{2}}}\cdots\sum_{i_{q}\in\mathbb{I}_{j-1}^{q}}1_{\mathbb{R}\mathcal{R}_{j}^{q,i_{q}}}\mathbb{1}_{\mathbb{D}_{j-1}^{N}}\right]
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}}+\sum_{q=1}^{p}\sum_{j=1}^{n+1}\frac{E^{Y}\left[\sum_{i_{1}\in\mathbb{I}_{j-1}^{T}}1_{\mathbb{R}\mathcal{R}_{j}^{1,i_{1}}}\sum_{i_{2}\in\mathbb{I}_{j-1}^{T}}1_{\mathbb{R}\mathcal{R}_{j}^{2,i_{2}}}\cdots\sum_{i_{q}\in\mathbb{I}_{j-1}^{q}}1_{\mathbb{R}\mathcal{R}_{j}^{q,i_{q}}}\mathbb{1}_{\mathbb{D}_{j-1}^{N}}\right]
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}}+\sum_{q=1}^{p}\sum_{j=1}^{n+1}\frac{E^{Y}\left[\sum_{i_{1}\in\
$$

and from [\(7.59,](#page-33-1)[7.47\)](#page-30-3) and [\(7.64\)](#page-33-0)

<span id="page-34-2"></span>
$$
E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}|\hat{\mathbb{K}}_{n+1}^{k}-\hat{\mathbb{K}}_{n+1}^{k}|\right|^{p}1_{\mathbb{D}_{n}^{N}}\right]
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}}+\sum_{q=1}^{p}\sum_{j=1}^{n+1}\frac{E^{Y}\left[1_{\mathbb{R}\mathcal{R}_{j}^{1}}1_{\mathbb{R}\mathcal{R}_{j}^{2}}\cdots1_{\mathbb{R}\mathcal{R}_{j}^{q}}1_{\mathbb{D}_{j-1}^{N}}\right]}{m_{N}^{p-q}}
$$
\n(7.68)

where

$$
\mathbb{R}\mathcal{R}_j^k = \mathbb{R}_j^k \bigtriangleup \mathcal{R}_j^k. \tag{7.69}
$$

<span id="page-35-0"></span>However, letting  $\widehat{\mathbb{K}}\widehat{\mathcal{K}}_j^{k,i} = |\widehat{\mathbb{K}}_j^{k,i} - \widehat{\mathcal{K}}_j^{k,i}|$ , we find by [\(7.53,](#page-32-1)[7.36,7.37\)](#page-29-2), exchangeability and [\(7.64\)](#page-33-0) that

$$
E^{Y}\left[\sum_{i_{1}\in\mathbb{I}_{j-1}^{1}\cup\mathcal{I}_{j-1}^{1}}1_{\mathbb{R}\mathcal{R}_{j}^{1,i_{1}}}\sum_{i_{2}\in\mathbb{I}_{j-1}^{2}\cup\mathcal{I}_{j-1}^{2}}1_{\mathbb{R}\mathcal{R}_{j}^{2,i_{2}}}\cdots\sum_{i_{q}\in\mathbb{I}_{j-1}^{q}\cup\mathcal{I}_{j-1}^{q}}1_{\mathbb{R}\mathcal{R}_{j}^{q,i_{q}}1_{\mathbb{D}_{j-1}^{N}}}\right]
$$
(7.70)  
\n
$$
\ll E^{Y}\left[\left(|\mathbb{A}_{j}-\sigma_{j}(1)|+\sum_{i_{1}}\widehat{\mathbb{K}}\widehat{\mathcal{K}}_{j-1}^{1,i_{1}}\right)\cdots\left(|\mathbb{A}_{j}-\sigma_{j}(1)|+\sum_{i_{q}}\widehat{\mathbb{K}}\widehat{\mathcal{K}}_{j-1}^{q,i_{q}}\right)1_{\mathbb{D}_{j-1}^{N}}\right]
$$
  
\n
$$
\ll \sum_{r=0}^{N}E^{Y}\left[|\mathbb{A}_{j}-\sigma_{j}(1)|^{q-r}\sum_{i_{1}}\widehat{\mathbb{K}}\widehat{\mathcal{K}}_{j-1}^{1,i_{1}}\sum_{i_{2}}\widehat{\mathbb{K}}\widehat{\mathcal{K}}_{j-1}^{2,i_{2}}\cdots\sum_{i_{r}}\widehat{\mathbb{K}}\widehat{\mathcal{K}}_{j-1}^{r,i_{r}}1_{\mathbb{D}_{j-1}^{N}}\right]
$$
  
\n
$$
\ll \sum_{r=0}^{q}E^{Y}\left[|\mathbb{A}_{j}-\sigma_{j}(1)|^{q-r}\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{j-1}^{k}\cup\mathcal{I}_{j-1}^{k}}\widehat{\mathbb{K}}\widehat{\mathcal{K}}_{j-1}^{k}\right|^{r}1_{\mathbb{D}_{j-1}^{N}}\right]
$$
  
\n
$$
\ll m_{N}^{-\frac{q}{2}}+E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{j-1}^{k}\cup\mathcal{I}_{j-1}
$$

<span id="page-35-1"></span>Substituting  $(7.70, 7.62)$  $(7.70, 7.62)$  into  $(7.65)$  and using Hölder's inequality, we find

$$
E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{n+1}^{k}\cup\mathcal{I}_{n+1}^{k}}|\hat{\mathbb{K}}_{n+1}^{k,i}-\hat{\mathcal{K}}_{n+1}^{k,i}|\right|^{p}1_{\mathbb{D}_{n}^{N}}\right]
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}} + \sum_{q=1}^{p}\sum_{j=1}^{n+1}\frac{m_{N}^{-\frac{q}{2}}}{\prod_{j=1}^{m_{N}}}\left|E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i}|\hat{\mathbb{K}}_{j-1}^{k,i}-\hat{\mathbb{K}}_{j-1}^{k,i}|\right|^{p}1_{\mathbb{D}_{j-2}^{N}}\right]\right)^{\frac{q}{p}}
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}}
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}}
$$
\n
$$
(7.71)
$$

<span id="page-35-2"></span>so

$$
E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{l}^{k}\cup\mathcal{I}_{l}^{k}}|\widehat{\mathbb{K}}_{l}^{k,i}-\widehat{\mathcal{K}}_{l}^{k,i}|\right|^{p}1_{\mathbb{D}_{l-1}^{N}}\right] \stackrel{N}{\ll}m_{N}^{-\frac{p}{2}} \quad \forall l\leq n+1.
$$
\n(7.72)

Similarly, replacing [\(7.65\)](#page-34-0) with [\(7.67\)](#page-34-1), we have

$$
E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{l}^{k}\cup\mathcal{I}_{l}^{k}}|\mathbb{K}_{l}^{k,i}-\mathcal{K}_{l}^{k,i}|\right|^{p}1_{\mathbb{D}_{l-1}^{N}}\right] \stackrel{N}{\ll}m_{N}^{-\frac{p}{2}} \quad \forall l\leq n+1.
$$
\n(7.73)

Turning to the tracking system and following [\(7.70\)](#page-35-0), we find by [\(7.55](#page-32-1)[,7.64\)](#page-33-0) and exchangeability that

$$
E^{Y} \left[ 1_{\underline{\mathbb{R}} \underline{\mathcal{R}}_{j}^{1}} 1_{\underline{\mathbb{R}} \underline{\mathcal{R}}_{j}^{2}} \cdots 1_{\underline{\mathbb{R}} \underline{\mathcal{R}}_{j}^{q}} 1_{\underline{\mathbb{D}}_{j-1}^{N}} \right]
$$
\n
$$
\ll m_{N}^{-\frac{q}{2}} + E^{Y} \left[ \left| \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \sum_{i} \left| \underline{\hat{\mathbb{K}}}_{j-1}^{k} - \underline{\hat{\mathbb{K}}}_{j-1}^{k} \right| \right|^{q} 1_{\underline{\mathbb{D}}_{j-1}^{N}} \right]
$$
\n(7.74)

so by [\(7.68\)](#page-34-2) and the method of [\(7.71](#page-35-1)[-7.72\)](#page-35-2) one has that

$$
E^{Y} \left[ \left| \frac{1}{m_N} \sum_{k=1}^{m_N} \left| \underline{\hat{\mathbb{K}}}_l^k - \underline{\hat{\mathbb{K}}}_l^k \right| \right|^p 1_{\mathbb{D}_{l-1}^N} \right] \stackrel{N}{\ll} m_N^{-\frac{p}{2}} \quad \forall l \le n+1. \tag{7.75}
$$

 $\Box$ 

With this coupling and prior preliminary results, we can establish our main result.

Proof of Theorem [5.1.](#page-21-0)

We can work directly on the coupled algorithms by  $(7.29,7.30)$ . **Mllns:** Taking  $p > \frac{2}{1-2\beta}$ , we then find by [\(7.44\)](#page-30-0), Theorem [7.2](#page-30-2) and Fubini's theorem that

$$
E^{Y} \left[ \sum_{N=1}^{\infty} \left| \frac{N^{\beta}}{m_{N}} \sum_{k=1}^{m_{N}} B_{n}^{N,k}(f) \right|^{p} \right]
$$
  
\n
$$
\ll \sum_{N=1}^{\infty} E^{Y} \left[ \left| \frac{N^{\beta}}{m_{N}} \sum_{k=1}^{m_{N}} \sum_{i} \left| \mathbb{K}_{n}^{k,i} - \mathcal{K}_{n}^{k,i} \right|^{p} \mathbf{1}_{\mathbb{D}_{n-1}^{N}} \right]
$$
  
\n
$$
\ll \sum_{N=1}^{\infty} N^{p\beta} m_{N}^{-\frac{p}{2}} \ll \sum_{N=1}^{\infty} N^{(\beta - \frac{1}{2})p} < \infty
$$
\n(7.76)

and it follows by  $N^{th}$ -term divergence that

$$
\sum_{N=1}^{\infty} \left| \frac{N^{\beta}}{m_N} \sum_{k=1}^{m_N} B_n^{N,k}(f) \right|^p < \infty \Rightarrow \frac{N^{\beta}}{m_N} \sum_{k=1}^{m_N} B_n^{N,k}(f) \to 0 \text{ a.s. } [Q^Y]. \tag{7.77}
$$

Moreover, by Borel-Cantelli and [\(7.35\)](#page-29-2)

$$
\sum_{N=1}^{\infty} Q((\mathbb{D}_{n-1}^{N})^{C}) \leq \sum_{N=1}^{\infty} 2(n-1)e^{-\epsilon_{n-1}m_{N}} < \infty
$$
\n
$$
\Rightarrow Q((\mathbb{D}_{n-1}^{N})^{C} \text{ i.o.}) = 0.
$$
\n(7.78)

Finally, we know

$$
\frac{1}{m_N} \sum_{k=1}^{m_N} \mathcal{K}_n^{k,i} f(\mathcal{X}_n^{k,i}) \ll N^{-\beta} \text{ a.s. } [Q^Y]. \tag{7.79}
$$

by  $(4.1)$  $(4.1)$  $(4.1)$  as well as Theorems 4.1 and [7.1](#page-29-3) so this part follows by  $(7.44)$ .

slln: This part follows from the Mllns, using the same  ${f_i} \subset C(E)$  as in the proof of Theorem [4.1.](#page-17-4)

<span id="page-37-0"></span>Establish i),ii) of Theorem [5.2:](#page-21-1) It follows by  $(7.44)$  and  $(7.36,7.37)$  that

$$
E^{Y}|B_{n}^{N,1}(f)|^{2} \ll E^{Y} \left[ \sum_{i \in \mathbb{I}_{n}^{1} \cup \mathcal{I}_{n}^{1}} |\mathbb{K}_{n}^{1,i} - \mathcal{K}_{n}^{1,i}|^{2} 1_{\mathbb{D}_{n-1}^{N}} \right]
$$
\n
$$
\ll E^{Y} \left[ \sum_{i \in \mathbb{I}_{n}^{1} \cup \mathcal{I}_{n}^{1}} |\mathbb{K}_{n}^{1,i} - \mathcal{K}_{n}^{1,i}| 1_{\mathbb{D}_{n-1}^{N}} \right]
$$
\n(7.80)

for  $f \in \overline{C}(E)_+$ . Hence, by [\(7.80\)](#page-37-0), exchangeability and [\(7.49\)](#page-31-0) with  $p = 1$ 

$$
N^{\frac{1}{2}}E^{Y}|B_{n}^{N,1}(f)|^{2} \ll \left(\frac{N}{m_{N}}\right)^{\frac{1}{2}} \to 0.
$$
\n(7.81)

<span id="page-37-1"></span>and Theorem [5.2](#page-21-1) i) is true with  $\rho = \frac{1}{2}$  and  $\psi_{N,k} = B_n^{N,k}(f)$ . It follows by  $(7.44)$  and  $(7.49)$  with  $p = 2$  that

<span id="page-37-2"></span>
$$
NE^{Y}|B_{n}^{N,1}(f)B_{n}^{N,2}(f)|
$$
\n
$$
\leq N|f|_{\infty}^{2}E^{Y}\left|\sum_{i\in\mathbb{I}_{n}^{1}\cup\mathcal{I}_{n}^{1}}\sum_{j\in\mathbb{I}_{n}^{2}\cup\mathcal{I}_{n}^{2}}|\mathbb{K}_{n}^{1,i}-\mathcal{K}_{n}^{1,i}||\mathbb{K}_{n}^{2,j}-\mathcal{K}_{n}^{2,j}|1_{\mathbb{D}_{n-1}^{N}}\right|
$$
\n
$$
\leq \frac{N}{m_{N}(m_{N}-1)}|f|_{\infty}^{2}E^{Y}\left[\left|\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{n}^{k}\cup\mathcal{I}_{n}^{k}}|\mathbb{K}_{n}^{k,i}-\mathcal{K}_{n}^{k,i}|\right|^{2}1_{\mathbb{D}_{n-1}^{N}}\right]
$$
\n
$$
\leq \frac{N}{m_{N}} \to 0.
$$
\n(7.82)

# Apply Exchangeability Result, Reduced System clt:

1  $N^{\frac{1}{2}}$  $\stackrel{N}{\blacklozenge}$  $k=1$  $B_n^{N,k}(f) \to^P 0$  by [\(7.82\)](#page-37-1), [\(7.81\)](#page-37-2) and Theorem [5.2](#page-21-1) with  $\rho = \frac{1}{2}$ , with  $E = E^Y$  and  $\psi_{N,k} = B_n^{N,k}(f)$ . Therefore, it follows by [\(7.44\)](#page-30-0), [\(7.35\)](#page-29-2) that for any  $\epsilon > 0$ 

$$
Q^{Y}(N^{-\frac{1}{2}}\sum_{k=1}^{N}\sum_{i}(\mathbb{K}_{n}^{k,i}f(X_{n}^{k,i}) - \mathcal{K}_{n}^{k,i}f(X_{n}^{k,i})) > \epsilon)
$$
\n
$$
\leq Q^{Y}(N^{-\frac{1}{2}}\sum_{k=1}^{N}B_{n}^{N,k}(f) > \epsilon) + Q^{Y}((\mathbb{D}_{n-1}^{N})^{C}) \to 0.
$$
\n(7.83)

The clt in Theorem [5.1](#page-21-0) now follows from the clt in Theorem [4.1](#page-17-4) and Theorem [7.1.](#page-29-3)  $\Box$ 

# 8. Appendix I: Proof of Proposition [4.1,](#page-15-1) variance calculation

<span id="page-37-3"></span>Abbreviating  $M_n^k = M_n^{\mathcal{B}^k}(f)$ , one notes from [\(4.11\)](#page-14-2) and [\(4.8\)](#page-14-0) that

$$
M_n^k = \sum_{l=0}^n \left[ \mathcal{B}_l^k \left( A_{l+1,n} f \right) - \mathcal{B}_{l-1}^k \left( A_{l,n} f \right) \right], \tag{8.1}
$$

<span id="page-38-0"></span>where  $\mathcal{B}_{-1}^k = \pi_0$ . The variance of the ' $l = 0$ ' term is

$$
E^{Y} \left| \mathcal{B}_{0}^{k}(A_{1,n}f) - \mathcal{B}_{-1}^{k}(A_{0,n}f) \right|^{2} = E^{Y} \left| (A_{1,n}f(\chi^{k})) - \pi_{0}(A_{1,n}f) \right|^{2}
$$
\n
$$
= \pi_{0} \left( (A_{1,n}f)^{2} \right) - \left( \pi_{0}(A_{1,n}f) \right)^{2}.
$$
\n(8.2)

The martingale differences for  $l \ge 1$  are by [\(4.1\)](#page-12-0), [\(4.4\)](#page-13-1), [\(4.2\)](#page-12-1), [\(3.2\)](#page-11-0), Step 4 of the algorithm and [\(4.7\)](#page-13-0)

$$
\mathcal{B}_{l}^{k}(A_{l+1,n}f) - \mathcal{B}_{l-1}^{k}(A_{l,n}f)
$$
\n
$$
= \sum_{i=1}^{\mathcal{N}_{l-1}^{k}} \left\{ \sum_{j=\overline{\mathcal{N}}_{l}^{k,i}}^{\overline{\mathcal{N}}_{l}^{k,i}} \mathcal{L}_{l}^{k,j} A_{l+1,n}f(\mathcal{X}_{l}^{k,j}) - \widehat{\mathcal{L}}_{l-1}^{k,i} K A_{l+1,n}f(\mathcal{X}_{l-1}^{k,i}) \right\}
$$
\n
$$
= \sum_{i=1}^{\mathcal{N}_{l-1}^{k}} \overline{\mathcal{L}}_{l-1}^{k,i} \left\{ \sum_{j=\overline{\mathcal{N}}_{l}^{k,i}}^{\overline{\mathcal{N}}_{l}^{k,i}} A_{l+1,n}f(\mathcal{X}_{l}^{k,j}) - \mathcal{L}_{l-1}^{k,i} \mathcal{L}_{l}^{k,i} \right\}
$$
\n
$$
-E^{Y} \left[ \sum_{j=\overline{\mathcal{N}}_{l}^{k,i-1}+1}^{\overline{\mathcal{N}}_{l}^{k,i}} A_{l+1,n}f(\mathcal{X}_{l}^{k,j}) \left| \mathcal{F}_{l-1}^{l}\vee \mathcal{F}_{l}^{\mathcal{V}} \right| \right] \right\}.
$$
\n(8.3)

Therefore, by the independence of the  $\{\mathcal{U},\mathcal{V},\mathcal{Z}\}$ 

$$
E^{Y} \left[ (\mathcal{B}_{t}^{k}(A_{l+1,n}f) - \mathcal{B}_{l-1}^{k}(A_{l,n}f))^{2} | \mathcal{F}_{l-1}^{l\mathcal{U}} \vee \mathcal{F}_{l}^{V} \right] \n= \sum_{i_{1},i_{2}=1}^{N_{t-1}^{k}} \overline{\mathcal{L}}_{l-1}^{k,i_{1}^{k}} \overline{\mathcal{L}}_{l-1}^{k,i_{2}} \n\left\{ E^{Y} \left[ \sum_{j_{1}=\overline{\mathcal{N}}_{t}^{k,i_{1}-1}+1}^{\overline{\mathcal{N}}_{t}^{k,i_{1}}} A_{l+1,n} f(\mathcal{X}_{l}^{k,j_{1}}) \sum_{j_{2}=\overline{\mathcal{N}}_{t}^{k,i_{2}-1}+1}^{\overline{\mathcal{N}}_{t}^{k,i_{2}}} A_{l+1,n} f(\mathcal{X}_{l}^{k,j_{2}}) \Big| \mathcal{F}_{l-1}^{l\mathcal{U}} \vee \mathcal{F}_{l}^{V} \right] \n- E^{Y} \left[ \sum_{j_{1}=\overline{\mathcal{N}}_{t}^{k,i_{1}-1}+1}^{\overline{\mathcal{N}}_{t}^{k,i_{1}}} A_{l+1,n} f(\mathcal{X}_{l}^{k,j_{1}}) \Big| \mathcal{F}_{l-1}^{l\mathcal{U}} \vee \mathcal{F}_{l}^{V} \right] \n\left\{ \sum_{j_{2}=\overline{\mathcal{N}}_{t}^{k,i_{2}-1}+1}^{\overline{\mathcal{N}}_{t}^{k,i_{2}-1}+1} \mathcal{A}_{l+1,n} f(\mathcal{X}_{l}^{k,j_{2}}) \Big| \mathcal{F}_{l-1}^{l\mathcal{U}} \vee \mathcal{F}_{l}^{V} \right] \right\} \n= \sum_{i=1}^{N_{t-1}^{k}} \left| \overline{\mathcal{L}}_{l-1}^{k,i} \right|^{2} \left\{ E^{Y} \left[ \sum_{j=\overline{\mathcal{N}}_{t}^{k,i_{1}-1}+1}^{\overline{\mathcal{N}}_{t}^{k,i_{1}}+1} A_{l+1,n} f(\mathcal{X}_{l}^{k,j}) \right]^{2} \Big| \mathcal{F}_{l-1}^{l\mathcal{U}} \vee \math
$$

However, by the independence of the  $\{\mathcal{U},\mathcal{V},\mathcal{Z}\}$  again as well as  $(4.6)$ 

$$
E^{Y}\left[\left(\sum_{j=\overline{\mathcal{N}}_{l}^{k,i}}^{\overline{\mathcal{N}}_{l}^{k,i}}A_{l+1,n}f(\mathcal{X}_{l}^{k,j})\right)^{2}|\mathcal{F}_{l-1}^{l\mathcal{U}\mathcal{X}}\vee\mathcal{F}_{l}^{V}\right]
$$
\n
$$
=E^{Y}\left[\mathcal{N}_{l}^{k,i}\left\{K(A_{l+1,n}f)^{2}(\mathcal{X}_{l-1}^{k,i})-(KA_{l+1,n}f)^{2}(\mathcal{X}_{l-1}^{k,i})\right\}
$$
\n
$$
+\left(\mathcal{N}_{l}^{k,i}KA_{l+1,n}f\right)^{2}(\mathcal{X}_{l-1}^{k,i})|\mathcal{F}_{l-1}^{l\mathcal{U}\mathcal{X}}\vee\mathcal{F}_{l}^{V}\right]
$$
\n
$$
=\frac{\hat{\mathcal{L}}_{l-1}^{k,i}}{\overline{\mathcal{L}}_{l-1}^{k,i}}\left\{K(A_{l+1,n}f)^{2}(\mathcal{X}_{l-1}^{k,i})-(KA_{l+1,n}f)^{2}(\mathcal{X}_{l-1}^{k,i})\right\}
$$
\n
$$
+\left\{\left|\frac{\hat{\mathcal{L}}_{l-1}^{k,i}}{\overline{\mathcal{L}}_{l-1}^{k,i}}\right|^{2}+\frac{\hat{\mathcal{L}}_{l-1}^{k,i}}{\overline{\mathcal{L}}_{l-1}^{k,i}}-\left|\frac{\hat{\mathcal{L}}_{l-1}^{k,i}}{\overline{\mathcal{L}}_{l-1}^{k,i}}\right|-\left|\frac{\hat{\mathcal{L}}_{l-1}^{k,i}}{\overline{\mathcal{L}}_{l-1}^{k,i}}\right|^{2}\right\}(KA_{l+1,n}f)^{2}(\mathcal{X}_{l-1}^{k,i}),
$$
\n(8.5)

since

$$
E^{Y}\Big[|\mathcal{N}_l^{k,i}|^2\left|\mathcal{F}_{l-1}^{l\lambda X}\vee\mathcal{F}_l^{\mathcal{V}}\right]=\left|\frac{\widehat{\mathcal{L}}_{l-1}^{k,i}}{\overline{\mathcal{L}}_{l-1}^{k,i}}\right|^2+\left|2\left|\frac{\widehat{\mathcal{L}}_{l-1}^{k,i}}{\overline{\mathcal{L}}_{l-1}^{k,i}}\right|+1\right|Q^{Y}\Bigg(U_l^{k,i}\leq\frac{\widehat{\mathcal{L}}_{l-1}^{k,i}}{\overline{\mathcal{L}}_{l-1}^{k,i}}-\left|\frac{\widehat{\mathcal{L}}_{l-1}^{k,i}}{\overline{\mathcal{L}}_{l-1}^{k,i}}\right|\Bigg),\right.
$$

and

$$
\begin{split}\n&\left| E^{Y} \left[ \sum_{j=\mathcal{N}_{l}^{k,i}}^{\overline{\mathcal{N}_{l}^{k,i}}} A_{l+1,n} f(\mathcal{X}_{l}^{k,j}) \left| \mathcal{F}_{l-1}^{U\mathcal{X}} \vee \mathcal{F}_{l}^{V} \right| \right]^{2} \right| \\
&= \left| E^{Y} \left[ \mathcal{N}_{l}^{k,i} K A_{l+1,n} f(\mathcal{X}_{l-1}^{k,i}) \left| \mathcal{F}_{l-1}^{U\mathcal{X}} \vee \mathcal{F}_{l}^{V} \right| \right]^{2} \\
&= \left| \frac{\widehat{\mathcal{L}}_{l-1}^{k,i}}{\overline{\mathcal{L}}_{l-1}^{k,i}} \right|^{2} (K A_{l+1,n} f)^{2} (\mathcal{X}_{l-1}^{k,i}).\n\end{split} \tag{8.6}
$$

<span id="page-40-0"></span>Combining the last three equations, letting  $f_{l,n} = A_l(A_{l+1,n}f)^2 - \alpha_l(KA_{l+1,n}f)^2$ , breaking over the resample and non-resample cases, and averaging over the  $\mathcal{V}_l^{k,i}$ , one finds by  $(4.17, 4.18, 4.2, 4.3)$  $(4.17, 4.18, 4.2, 4.3)$  $(4.17, 4.18, 4.2, 4.3)$  $(4.17, 4.18, 4.2, 4.3)$  $(4.17, 4.18, 4.2, 4.3)$  $(4.17, 4.18, 4.2, 4.3)$  that

$$
E^{Y} \left[ (\mathcal{B}_{l}^{k}(A_{l+1,n}f) - \mathcal{B}_{l-1}^{k}(A_{l,n}f))^{2} | \mathcal{F}_{l-1}^{lVX} \right]
$$
\n
$$
= E^{Y} \left[ \sum_{i=1}^{N_{l-1}^{k}} \overline{\mathcal{L}}_{l-1}^{k,i} \widehat{\mathcal{L}}_{l-1}^{k,i} \left\{ K(A_{l+1,n}f)^{2}(\mathcal{X}_{l-1}^{k,i}) - (KA_{l+1,n}f)^{2}(\mathcal{X}_{l-1}^{k,i}) \right\} \right]
$$
\n
$$
+ \sum_{i=1}^{N_{l-1}^{k}} \left( \overline{\mathcal{L}}_{l-1}^{k,i} \right)^{2} r \left( \frac{\widehat{\mathcal{L}}_{l-1}^{k,i}}{\overline{\mathcal{L}}_{l-1}^{k,i}} \right) (KA_{l+1,n}f)^{2}(\mathcal{X}_{l-1}^{k,i}) | \mathcal{F}_{l-1}^{lVX} \right]
$$
\n
$$
= \sigma_{l}(1) \sum_{i=1}^{N_{l-1}^{k}} \mathcal{L}_{l-1}^{k,i} \overline{\nu}_{l}(\alpha_{l}(\mathcal{X}_{l-1}^{k,i}) \mathcal{L}_{l-1}^{k,i}) \left\{ f_{l,n}(\mathcal{X}_{l-1}^{k,i}) \right\}
$$
\n
$$
+ \sigma_{l}^{2}(1) \sum_{i=1}^{N_{l-1}^{k}} \overline{\nu}_{l}(\alpha_{l}(\mathcal{X}_{l-1}^{k,i}) \mathcal{L}_{l-1}^{k,i}) r \left( \frac{\widehat{\mathcal{L}}_{l-1}^{k,i}}{\sigma_{l}(1)} \right) (KA_{l+1,n}f)^{2}(\mathcal{X}_{l-1}^{k,i})
$$
\n
$$
+ \sum_{i=1}^{N_{l-1}^{k}} \left( \mathcal{L}_{l-1}^{k,i} \right)^{2} \alpha_{l}(\mathcal{X}_{l-1}^{k,i}) \nu_{l}(\alpha_{l}(\mathcal{X}_{l-1}^{k,i}) \mathcal{L}_{l-1}^{k,i}) \left\{ f_{l,n}(\mathcal{X}_{l-1}^{k,i}) \right\}
$$

since r  $\left(\frac{\widehat{\mathcal{L}}_{i-1}^{k,i}}{\sum_{i=1}^{k}1}\right)$  $\widehat{\mathcal{L}}_{l-1}^{k,i}$  $\setminus$  $= 0$ . Now, in the case ' $l = 1$ ' we have  $\mathcal{L}_{l-1}^{k,i} = 1 = \mathcal{N}_{l-1}^{k}$  and

$$
E^{Y}[(\mathcal{B}_{1}^{k}(A_{2,n}f) - \mathcal{B}_{0}^{k}(A_{l,n}f))^{2}]
$$
\n
$$
= \sigma_{1}(1)E^{Y}[\overline{\nu}_{1}(\alpha_{1}(\chi^{k})) \{A_{1}(A_{2,n}f)^{2}(\chi^{k}) - \alpha_{1}(\chi^{k})(KA_{2,n}f)^{2}(\chi^{k})\}]
$$
\n
$$
+ \sigma_{1}^{2}(1)E^{Y}[\overline{\nu}_{1}(\alpha_{1}(\chi^{k}))r\left(\frac{\alpha_{1}(\chi^{k})}{\sigma_{1}(1)}\right)(KA_{2,n}f)^{2}(\chi^{k})]
$$
\n
$$
+ E^{Y}[\alpha_{1}(\chi^{k})\nu_{1}(\alpha_{1}(\chi^{k})) \{A_{1}(A_{2,n}f)^{2}(\chi^{k}) - \alpha_{1}(\chi^{k})(KA_{2,n}f)^{2}(\chi^{k})\}].
$$
\n(8.8)

<span id="page-41-0"></span>Moreover, for any  $l \ge 2$ ,  $m \in \{1, 2, ..., l-1\}$  and bounded function  $\phi_m$ , we have by [\(4.4\)](#page-13-1), [\(4.6\)](#page-13-3), [\(3.1\)](#page-11-0), independence and division over resampled and non-resampled cases that

$$
E^{Y}\left[\sum_{i=1}^{N_{m}^{k}}\mathcal{L}_{m}^{k,i}\phi_{m}(\mathcal{X}_{m}^{k,i},\mathcal{L}_{m}^{k,i})\Big|\mathcal{F}_{m-1}^{U\mathcal{V}\mathcal{X}}\right]
$$
\n
$$
=\sum_{j=1}^{N_{m-1}^{k}}E^{Y}\left[\sum_{i=\overline{\mathcal{N}}_{m}^{k,j}}^{\overline{\mathcal{N}}_{m}^{k,j}}\overline{\mathcal{L}}_{m-1}^{k,j}\phi_{m}(\mathcal{X}_{m}^{k,i},\overline{\mathcal{L}}_{m-1}^{k,j})\Big|\mathcal{F}_{m-1}^{U\mathcal{V}\mathcal{X}}\right]
$$
\n
$$
=\sum_{j=1}^{N_{m-1}^{k}}E^{Y}\left[\sum_{i=\overline{\mathcal{N}}_{m}^{k,j}}^{\overline{\mathcal{N}}_{m}^{k,j}}\overline{\mathcal{L}}_{m-1}^{k,j}K\phi_{m}(\mathcal{X}_{m-1}^{k,j},\overline{\mathcal{L}}_{m-1}^{k,j})\Big|\mathcal{F}_{m-1}^{U\mathcal{V}\mathcal{X}}\right]
$$
\n
$$
=\sum_{j=1}^{N_{m-1}^{k}}E^{Y}\left[\hat{\mathcal{L}}_{m-1}^{k,j}K\phi_{m}(\mathcal{X}_{m-1}^{k,j},\overline{\mathcal{L}}_{m-1}^{k,j})\Big|\mathcal{F}_{m-1}^{U\mathcal{V}\mathcal{X}}\right]
$$
\n
$$
=\sum_{j=1}^{N_{m-1}^{k}}\mathcal{L}_{m-1}^{k,j}E^{Y}\left[A_{m}\phi_{m}(\mathcal{X}_{m-1}^{k,j},\overline{\mathcal{L}}_{m-1}^{k,j})\Big|\mathcal{F}_{m-1}^{U\mathcal{V}\mathcal{X}}\right]
$$
\n
$$
=\sum_{j=1}^{N_{m-1}^{k}}\mathcal{L}_{m-1}^{k,j}\phi_{m-1}(\mathcal{X}_{m-1}^{k,j},\mathcal{L}_{m-1}^{k,j}),
$$
\n(8.9)

<span id="page-41-3"></span>where

$$
\phi_{m-1}(\mathcal{X}, \mathcal{L})
$$
\n
$$
= A_m \phi_m(\mathcal{X}, \sigma_m(1)) \overline{\nu}_m(\alpha_m(\mathcal{X})\mathcal{L}) + A_m \phi_m(\mathcal{X}, L) \Big|_{L=\alpha_m(\mathcal{X})\mathcal{L}} \nu_m(\alpha_m(\mathcal{X})\mathcal{L}).
$$
\n(8.10)

<span id="page-41-1"></span>[\(8.9\)](#page-41-0) implies that

$$
E^{Y}\left[\sum_{i=1}^{\mathcal{N}_{1}^{k}} \mathcal{L}_{1}^{k,i} \phi_{1}(\mathcal{X}_{1}^{k,i}, \mathcal{L}_{1}^{k,i})\right]
$$
\n
$$
= \pi_{0}\left[A_{1}\phi_{1}(\cdot, \sigma_{1}(1))\overline{\nu}_{1}(\alpha_{1}(\cdot)) + A_{1}\phi_{1}(\cdot, \alpha_{1}(\cdot))\nu_{1}(\alpha_{1}(\cdot))\right].
$$
\n(8.11)

<span id="page-41-2"></span>Now, recall [\(4.23\)](#page-15-4) and suppose that

$$
E^{Y}\left[\sum_{i=1}^{N_{m-1}^{k}} \mathcal{L}_{m-1}^{k,i} \phi_{m-1}(\mathcal{X}_{m-1}^{k,i}, \mathcal{L}_{m-1}^{k,i})\right]
$$
\n
$$
= \sum_{j=0}^{m-1} \sum_{1 \le i_1 < i_2 < \dots < i_j \le m-1} \pi_0[A_{1,m-1}\phi_{m-1}(\cdot, \alpha_{i_j,m-1})\nu_{i_j,m-1}\overline{\nu}_{i_1,i_2,\dots,i_j}]
$$
\n(8.12)

<span id="page-42-0"></span>for some  $m \in \{2, ..., l-1\}$ , which is known when  $m = 2$  by  $(8.11)$  and  $(4.19, 4.20, 4.21)$ . (For clarity, the "j = 0" term on the right of [\(8.12\)](#page-41-2) is simply  $\pi_0[A_{1,m-1}\phi_{m-1}(\cdot,\alpha_{0,m-1})\nu_{0,m-1}]$ .) Then, it follows from  $(8.9,8.10,8.12)$  $(8.9,8.10,8.12)$  $(8.9,8.10,8.12)$  and  $(4.19,4.20,4.21)$  by letting  $r = j + 1$  that

$$
E^{Y}\left[\sum_{i=1}^{N_{m}^{k}}\mathcal{L}_{m}^{k,i}\phi_{m}(\mathcal{X}_{m}^{k,i},\mathcal{L}_{m}^{k,i})\right]
$$
\n
$$
=\sum_{j=0}^{m-1}\sum_{1\leq i_{1}< i_{2}< \cdots < i_{j}\leq m-1}\pi_{0}[A_{1,m}\phi_{m}(\cdot,\sigma_{m}(1))\overline{\nu}_{i_{j},m}\nu_{i_{j},m-1}\overline{\nu}_{i_{1},i_{2},\ldots,i_{j}}]
$$
\n
$$
+\sum_{j=0}^{m-1}\sum_{1\leq i_{1}< i_{2}< \cdots < i_{j}\leq m-1}\pi_{0}[A_{1,m}\phi_{m}(\cdot,\alpha_{i_{j},m})\nu_{i_{j},m}\overline{\nu}_{i_{1},i_{2},\ldots,i_{j}}]
$$
\n
$$
=\sum_{r=1}^{m}\sum_{1\leq i_{1}< i_{2}< \cdots < i_{r}\leq m}\pi_{0}[A_{1,m}\phi_{m}(\cdot,\alpha_{i_{r},m})\nu_{i_{r},m}\overline{\nu}_{i_{1},i_{2},\ldots,i_{r}}]
$$
\n
$$
+\sum_{j=0}^{m-1}\sum_{1\leq i_{1}< i_{2}< \cdots < i_{j}\leq m-1}\pi_{0}[A_{1,m}\phi_{m}(\cdot,\alpha_{i_{j},m})\nu_{i_{j},m}\overline{\nu}_{i_{1},i_{2},\ldots,i_{j}}]
$$
\n
$$
=\sum_{j=0}^{m}\sum_{1\leq i_{1}< i_{2}< \cdots < i_{j}\leq m}\pi_{0}[A_{1,m}\phi_{m}(\cdot,\alpha_{i_{j},m})\nu_{i_{j},m}\overline{\nu}_{i_{1},i_{2},\ldots,i_{j}}].
$$
\n(8.13)

Hence,  $(8.13)$  holds for all  $m = 1, ..., l - 1$  by induction and  $(4.22)$  follows by  $(8.1)$ ,  $(8.2)$ ,  $(8.7)$  and  $(8.13)$ , considering the three cases:

$$
\phi_{l-1}(\mathcal{X}, \mathcal{L}) = \{f_{l,n}(\mathcal{X})\} \overline{\nu}_l(\alpha_l(\mathcal{X})\mathcal{L})
$$
\n(8.14)

$$
\phi_{l-1}(\mathcal{X}, \mathcal{L}) = \mathcal{L}\alpha_l(\mathcal{X}) \{f_{l,n}(\mathcal{X})\} \nu_l(\alpha_l(\mathcal{X})\mathcal{L}) \tag{8.15}
$$

$$
\phi_{l-1}(\mathcal{X}, \mathcal{L}) = \frac{1}{\mathcal{L}} \overline{\nu}_l(\alpha_l(\mathcal{X}) \mathcal{L}) r\left(\frac{\alpha_l(\mathcal{X}) \mathcal{L}}{\sigma_l(1)}\right) (KA_{l+1,n} f)^2(\mathcal{X}), \tag{8.16}
$$

where  $f_{l,n} = A_l (A_{l+1,n} f)^2 - \alpha_l (K A_{l+1,n} f)^2$ .  $\Box$ 

# 9. Appendix 2: Proof of Lemma [7.1](#page-31-1)

Essentially, we observe that this result would hold trivially for the weighted particle system and then use induction and the coupling to show the necessary differences between the resampled and weighted systems converge appropriately.

*Proof.* Recall  $\mathbb{A}_0^{m_N} = 1$  so  $\mathbb{A}_0^{m_N} = \sigma_0(1)$ . It follows from  $(7.19, 7.10, 7.14, 7.15)$  $(7.19, 7.10, 7.14, 7.15)$  $(7.19, 7.10, 7.14, 7.15)$  that for all  $n \geq 0$ 

<span id="page-42-1"></span>
$$
\mathbb{A}_{n+1}^{m_N} = \frac{1}{m_N} \sum_{k=1}^{m_N} \sum_{i \in \mathbb{I}_n^k} \alpha_{n+1}(X_n^{k,i}) \mathbb{K}_n^{k,i}.
$$
\n(9.1)

**Base Case:** For notational reasons we consider the case  $n = 0$  separately. One then finds by [\(7.1](#page-26-0)[,7.2\)](#page-27-1) that [\(9.1\)](#page-42-1) reduces to

$$
\mathbb{A}_1^{m_N} = \frac{1}{m_N} \sum_{k=1}^{m_N} \alpha_1(\chi^k), \tag{9.2}
$$

where  $\{\alpha_1(\chi^k)\}_{k=1}^{\infty}$  are i.i.d., bounded and mean  $\sigma_1(1)$  with respect to  $Q^Y$ . Hence, by the Marcinkiewicz-Zygmund and Jensen inequalities there is a constant  $C_p > 0$  such that

$$
E^{Y} |\mathbb{A}_{1}^{m_{N}} - \sigma_{1}(1)|^{p} \leq \frac{C_{p}}{m_{N}^{p}} E^{Y} \left| \sum_{k=1}^{m_{N}} \left( \alpha_{1}(\chi^{k}) - \sigma_{1}(1) \right)^{2} \right|^{p} \leq \frac{C_{p}}{m_{N}^{p}} \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} E^{Y} |\alpha_{1}(\chi^{k}) - \sigma_{1}(1)|^{p} \ll m_{N}^{-\frac{p}{2}}
$$
\n(9.3)

for any  $p\geq 1.$ 

<span id="page-43-0"></span>**Case**  $n \geq 1$ : It follows from  $(7.3, 7.2, 7.15, 7.21, 7.45)$  $(7.3, 7.2, 7.15, 7.21, 7.45)$  that  $\mathbb{K}_0^{k,1} = 1$  and

$$
\mathbb{K}_{j}^{k,i} = \mathbb{A}_{j}^{m_{N}} \mathbf{1}_{\mathbb{R}_{j}^{k,p_{j}(i)}} + W_{j}^{k,p_{j}(i)} \mathbb{K}_{j-1}^{k,p_{j}(i)} \mathbf{1}_{\mathbb{S}_{j}^{k,p_{j}(i)}} \ \forall i \in \mathbb{I}_{j}^{k}, j \in \mathbb{N}.
$$
\n(9.4)

<span id="page-43-1"></span>Using  $(9.1, 7.34)$  $(9.1, 7.34)$  and  $(9.4)$  recursively, one has

$$
A_{n+1}^{m_{N}} \n= \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \sum_{i_{n-1}=1}^{N_{k}} \sum_{i_{n-1}=1}^{N_{k,i_{n-1}}} W_{n+1}^{k,i_{n}} A_{n}^{m_{N}} 1_{\mathbb{R} \mathbb{S} \mathbb{I}_{n,n}^{k,i_{n-1},i_{n}}} + \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \sum_{i_{n-2}=1}^{N_{k-2}} \sum_{i_{n-1}=N_{k,i_{n-2}}^{k,i_{n-2}}} W_{n+1}^{k,i_{n}} W_{n}^{k,i_{n-1}} A_{n-1}^{m_{N}} 1_{\mathbb{R} \mathbb{S} \mathbb{I}_{n-1,n}^{k,i_{n-2},i_{n-1},i_{n}}} + \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \sum_{i_{n-3}=1}^{N_{k-3}} \sum_{i_{n-2}=N_{k,i_{n-3}}^{k,i_{n-2}}} W_{n+1}^{k,i_{n}} W_{n}^{k,i_{n-1}} W_{n-1}^{k,i_{n-2}} A_{n-2}^{m_{N}} 1_{\mathbb{R} \mathbb{S} \mathbb{I}_{n-2,n}^{k,i_{n-3},\dots,i_{n}}} + \cdots + \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \sum_{i_{1}=1}^{N_{k}} W_{n+1}^{k,i_{n-1}} W_{n}^{k,i_{2}} W_{2}^{k,i_{1}} A_{1}^{m_{N}} 1_{\mathbb{R} \mathbb{S} \mathbb{I}_{n,n}^{k,i_{0},\dots,i_{n}}} + \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \sum_{i_{1}=1}^{N_{k}} W_{n+1}^{k,i_{n}} W_{n}^{k,i_{n-1}} \cdots W_{3}^{k,i_{2}} W_{2}^{k,i_{1}} A_{1}^{m_{N}} 1_{\mathbb{R} \mathbb{S} \mathbb{I}_{n,n}^{k,i_{0},\dots,i_{n}}} + \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} W_{n+1}^{k,1} W_{n}^{k,1} \cdots W_{3}^{k,1} W_{2}^{k,1} W_{1}^{k,1}
$$

where the non-summed indices satisfy  $i_0 = 1$ ,  $i_l = \overline{N}_l^{k, i_{l-1}-1} + 1$  (since no resampling). For clarity, here and below  $\mathbb{RSI}_{0,n}^{k,1,\dots,1} = \mathbb{S}_{1}^{k,1} \mathbb{S}_{2}^{k,1} \cdots \mathbb{S}_{n}^{k,1} \{1 \in \mathbb{I}_{n}^{k}\}\$ . Noting  $\mathbb{S}_{1}^{k,1} \mathbb{S}_{2}^{k,1} \cdots \mathbb{S}_{n}^{k,1} \{1 \in \mathbb{I}_{n}^{k}\} = (\underline{\mathbb{R}}_{1}^{k})^{C} (\underline{\mathbb{R}}_{2}^{k})^{C} \cdots (\underline{\mathbb{R}}_{n}^{k})^{C}$  (by <span id="page-44-0"></span>[\(7.21,](#page-28-1) [7.25,](#page-28-1) [7.2,](#page-27-1) [7.3,](#page-27-1) [7.5,](#page-27-1) [7.15,](#page-28-0) [7.18\)](#page-28-2)) and letting

$$
W_{1,n+1}^{k,1} \triangleq W_{n+1}^{k,1} W_n^{k,1} \cdots W_2^{k,1} W_1^{k,1}
$$
  
=  $W_{1,n+1}^{k,1} 1_{\mathbb{R} \mathbb{S}^{\mathbb{R}^{k,1},\ldots,1}_{0,n}} + W_{1,n+1}^{k,1} 1_{\mathbb{R}^k_1(\mathbb{R}^k_2)^C \cdots(\mathbb{R}^k_{n-1})^C(\mathbb{R}^k_n)^C} + \cdots$   
+  $W_{1,n+1}^{k,1} 1_{\mathbb{R}^k_{n-2}(\mathbb{R}^k_{n-1})^C(\mathbb{R}^k_n)^C} + W_{1,n+1}^{k,1} 1_{\mathbb{R}^k_{n-1}(\mathbb{R}^k_n)^C} + W_{1,n+1}^{k,1} 1_{\mathbb{R}^k_n},$  (9.6)

<span id="page-44-1"></span>we have by  $(9.5-9.6)$  $(9.5-9.6)$  that

$$
m_{N}A_{n+1}^{m_{N}} - \sigma_{n+1}(1)
$$
\n
$$
= \sum_{k=1}^{m_{N}} \left| W_{1,n+1}^{k,1} - \sigma_{n+1}(1) \right|
$$
\n
$$
+ \sum_{k=1}^{m_{N}} \left| \sum_{i_{n-1}=1}^{N_{n-1}^{k}} \sum_{i_{n-1}=1}^{N_{n-1}^{k,i_{n-1}}} W_{n+1}^{k,i_{n}} A_{n}^{m_{N}} 1_{\text{RSI}_{n,n}^{k,i_{n-1},i_{n}} - W_{1,n+1}^{k,1} 1_{\underline{\mathbb{R}}_{n}^{k}} \right|
$$
\n
$$
+ \sum_{k=1}^{m_{N}} \left| \sum_{i_{n-2}=1}^{N_{n-2}} \sum_{i_{n-1}=N_{n-1}^{k,i_{n-2}}}^{N_{n-1}^{k,i_{n-1}} - 1} W_{n+1}^{k,i_{n}} W_{n}^{k,i_{n-1}} A_{n-1}^{m_{N}} 1_{\text{RSI}_{n-1,n}^{k,i_{n-2},i_{n-1},i_{n}} - W_{1,n+1}^{k,1} 1_{\underline{\mathbb{R}}_{n}^{k}} \right|
$$
\n
$$
+ \sum_{k=1}^{m_{N}} \left| \sum_{i_{n-2}=1}^{N_{n-3}} \sum_{i_{n-1}=N_{n-2}^{k,i_{n-3}}}^{N_{n-2}^{k,i_{n-2}} - 1} W_{n+1}^{k,i_{n}} W_{n}^{k,i_{n-1}} W_{n-1}^{k,i_{n-2}} A_{n-2}^{m_{N}} 1_{\text{RSI}_{n-2,n}^{k,i_{n-3},\dots,i_{n}} - W_{1,n+1}^{k,1} 1_{\underline{\mathbb{R}}_{n}^{k}} \sum_{i_{n-2},n}^{N_{n-2}} W_{n+1}^{k,i_{n-1}} W_{n}^{k,i_{n-1}} W_{n-1}^{k,i_{n-2}} A_{n-2}^{m_{N}} 1_{\text{RSI}_{n-2,n}^{k,i_{n-3},\dots,i_{n}} - W_{1,n+1}^{k} 1_{\underline{\mathbb{R}}_{n}^{k}} W_{n}^{k,i_{n-1}} \dots W_{3}^{k,i_{2}} W_{2}^{
$$

Now,  ${W}_{1,n+1}^{k,1} - \sigma_{n+1}(1)$ <sub>k=1</sub> are i.i.d., zero mean and bounded with respect to  $Q^Y$ . Therefore, it follows as above by the Marcinkiewicz-Zygmund and Jensen inequalities that

$$
E^{Y} \left| \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} W_{1,n+1}^{k,1} - \sigma_{n+1}(1) \right|^{p} \ll m_{N}^{-\frac{p}{2}} \tag{9.8}
$$

<span id="page-45-0"></span>for any  $p \ge 1$ . Next, we consider a typical (non-first) term in [\(9.7\)](#page-44-1) in terms of  $l \in \{1, ..., n\}$ 

$$
\sum_{i_{l-1}=1}^{N_{l-1}^k} \sum_{i_l=N_l^{k,i_{l-1}}+1}^{\overline{N}_l^{k,i_{l-1}}} W_{n+1}^{k,i_n} \cdots W_{l+2}^{k,i_{l+1}} W_{l+1}^{k,i_l} \mathbb{A}_l^{m_N} 1_{\mathbb{R} \mathbb{S}_{l,n}^{k,i_{l-1},\dots,i_n}}
$$
\n
$$
-W_{1,n+1}^{k,1} 1_{\mathbb{E}_l^k(\mathbb{E}_{l+1}^k)} c \cdots (\mathbb{E}_n^k) c = \mathbb{T}_1^k + \mathbb{T}_2^k + \mathbb{T}_3^k + \mathbb{T}_4^k,
$$
\n
$$
(9.9)
$$

where

$$
\mathbb{T}_{1}^{k} = \sum_{i_{l-1}=1}^{N_{l-1}^{k}} \sum_{i_{l}=\overline{N}_{l}^{k,i_{l-1}}}^{\overline{N}_{l}^{k,i_{l-1}}} W_{n+1}^{k,i_{n}} \cdots W_{l+2}^{k,i_{l+1}} W_{l+1}^{k,i_{l}} \sigma_{l}(1) 1_{\mathcal{R\mathcal{S}}\mathcal{I}_{l,n}^{k,i_{l-1},\ldots,i_{n}}} \tag{9.10}
$$
\n
$$
-W_{1,n+1}^{k,1} 1_{\underline{\mathcal{R}}_{l}^{k}(\underline{\mathcal{R}}_{l+1}^{k})} \cdots (\underline{\mathcal{R}}_{n}^{k})^{C}
$$

$$
\mathbb{T}_{2}^{k} = \n\sum_{\substack{N_{l-1}^{k} \\ i_{l-1} = 1}} \sum_{\substack{N_{l}^{k, i_{l-1}} \\ i_{l-1} = 1}}^{N_{l}^{k, i_{l-1}}} W_{n+1}^{k, i_{n}} \cdots W_{l+2}^{k, i_{l+1}} W_{l+1}^{k, i_{l}} (\mathbb{A}_{l}^{m_{N}} - \sigma_{l}(1)) 1_{\mathbb{RSI}_{l,n}^{k, i_{l-1}, \dots, i_{n}}}
$$
\n(9.11)

$$
\mathbb{T}_{3}^{k} = \sum_{i_{l-1}=1}^{N_{l-1}^{k}} \sum_{i_{l}=\overline{N}_{l}^{k,i_{l-1}}+1}^{\overline{N}_{l}^{k,i_{l-1}}} W_{n+1}^{k,i_{n}} \cdots W_{l+2}^{k,i_{l+1}} W_{l+1}^{k,i_{l}} \sigma_{l}(1)
$$
\n
$$
\times \left(1_{\mathbb{R}\mathbb{S}_{l,n}^{k,i_{l-1},\dots,i_{n}}-1} \mathcal{R}_{\mathcal{S}\mathcal{I}_{l,n}^{k,i_{l-1},\dots,i_{n}}}\right)
$$
\n(9.12)

and

$$
\mathbb{T}_4^k = W_{1,n+1}^{k,1} 1_{\underline{\mathcal{R}}_l^k(\underline{\mathcal{R}}_{l+1}^k)^C \cdots (\underline{\mathcal{R}}_n^k)^C} - W_{1,n+1}^{k,1} 1_{\underline{\mathbb{R}}_l^k(\underline{\mathbb{R}}_{l+1}^k)^C \cdots (\underline{\mathbb{R}}_n^k)^C}.
$$
\n(9.13)

**Bound**  $\mathbb{T}_1$ : The sums in  $\mathbb{T}_1$  only involve the reduced system so by Theorem [7.1](#page-29-3) we can just work in the original (prior to coupling) reduced system setting. Now, recalling  $\nu_l, \alpha_{i,m}, \overline{\nu}_{i,m}$  from [\(4.18](#page-15-3)[,4.19,4.21\)](#page-15-5) and using [\(7.33\)](#page-29-4), one has by [\(7.34](#page-29-4)[,5.5\)](#page-20-2), independence, the fact  $\sigma_l(1)1_{\mathcal{H}_l^{k,i_{l-1}}} = \overline{\mathcal{L}}_{l-1}^{k,i_{l-1}}$  $\frac{d_{l-1}}{l-1}1_{\mathcal{H}_l^{k,i_{l-1}}},\,(4.6),\,(4.4)\,\,\text{and}\,\,$  <span id="page-46-0"></span>[\(4.1\)](#page-12-0) that

$$
E^{Y}\left[\sum_{i_{l-1}=1}^{N_{l-1}^{k}}\sum_{i_{l-1}=1}^{N_{l-1}^{k,i_{l-1}}} \mathcal{W}_{l+1,n+1}^{k,i_{l},...,i_{n}}\sigma_{l}(1)1_{\mathcal{H}_{l}^{k,i_{l-1}}}(\mathcal{H}_{l+1}^{k,i_{l}})^{C} \cdots(\mathcal{H}_{n}^{k,i_{n-1}})^{C}\middle|\mathcal{F}_{l-1}^{U\mathcal{Y}\mathcal{X}}\right]
$$
\n
$$
=\sum_{i_{l-1}=1}^{N_{l-1}^{k}}E^{Y}\left[\sigma_{l}(1)\sum_{i_{l}=\overline{\mathcal{N}}_{l}^{k,i_{l-1}-1}+1}^{N_{l+1-1}^{k,i_{l-1}-1}+1}(\mathcal{H}_{l}^{k,i_{l}},...,\mathcal{X}_{n-1}^{k,i_{n-1}})\right]
$$
\n
$$
\times \mathcal{W}_{n}^{k,i_{n-1}}\cdots \mathcal{W}_{l+2}^{k,i_{l+1}}\mathcal{W}_{l+1}^{k,i_{l}}1_{\mathcal{H}_{l}^{k,i_{l-1}}}( \mathcal{H}_{l+1}^{k,i_{l}})^{C} \cdots(\mathcal{H}_{n-1}^{k,i_{n-2}})^{C}\middle|\mathcal{F}_{l-1}^{U\mathcal{Y}\mathcal{X}}\right]
$$
\n
$$
=\sum_{i_{l-1}=1}^{N_{l-1}^{k}}E^{Y}\left[\sigma_{l}(1)\sum_{i_{l}=\overline{\mathcal{N}}_{l}^{k,i_{l-1}}-1}^{N_{l,i_{l-1}-1}^{k,i_{l-1}}+1}(\mathcal{X}_{l}^{k,i_{l}})^{1}_{\mathcal{H}_{l}^{k,i_{l-1}}}\middle|\mathcal{F}_{l-1}^{U\mathcal{Y}\mathcal{X}}\right]
$$
\n
$$
=\sum_{i_{l-1}=1}^{N_{l-1}^{k}}\hat{\mathcal{L}}_{l-1}^{k,i_{l-1}}E^{Y}\left[\Gamma_{l,n+1}^{k}(X_{l}^{k,i_{l}})1_{\mathcal{H}_{l}^{k,i_{l-1}}}\middle|\mathcal{F}_{l-1}^{U\mathcal{Y}\mathcal{X}}\right]
$$
\n
$$
=\sum_{i_{l-1}=1}^{N_{
$$

and by [\(7.34\)](#page-29-4), [\(7.26,](#page-28-1) [7.18,](#page-28-2) [7.11,](#page-27-0) [7.6\)](#page-27-1), [\(4.18](#page-15-3)[,4.19,4.21\)](#page-15-5) that

<span id="page-46-1"></span>
$$
E^{Y}\left[W_{1,n+1}^{k,1}1_{\underline{\mathcal{R}}_{l}^{k}}(\underline{\mathcal{R}}_{l+1}^{k})^{C} \cdots (\underline{\mathcal{R}}_{n}^{k})^{C}\right] \mathcal{F}_{l-1}^{UVX}\right]
$$
\n
$$
= W_{1,l}^{k,1} \left(\overline{\nu}_{l} \circ \alpha_{l} K \Gamma_{l,n+1}^{k}\right) (X_{l-1}^{k,1})
$$
\n
$$
= \prod_{m=1}^{l} \alpha_{m} (X_{m-1}^{k,1}) \left(\overline{\nu}_{l} \circ \alpha_{l} K \Gamma_{l,n+1}^{k}\right) (X_{l-1}^{k,1})
$$
\n(9.15)

for  $l = 1, ..., n$ , where

$$
\Gamma_{l,n+1}^{k}(x_{l})
$$
\n
$$
= (\alpha_{l+1}\nu_{l+1} \circ \alpha_{l,l+1} K (\alpha_{l+2}\nu_{l+2} \circ \alpha_{l,l+2} \cdots K (\alpha_{n}\nu_{n} \circ \alpha_{l,n} K \alpha_{n+1}))) (x_{l}).
$$
\n(9.16)

Hence, by [\(7.31\)](#page-29-1), [\(9.14\)](#page-46-0), [\(4.1\)](#page-12-0), [\(4.9\)](#page-14-1), [\(1.1\)](#page-1-0), [\(1.2\)](#page-1-1) and [\(9.15\)](#page-46-1)

$$
E^{Y} \left[ \sum_{i_{l-1}=1}^{N_{l-1}^{k}} \sum_{i_{l}=\overline{N}_{l}^{k,i_{l-1}}}^{\overline{N}_{l}^{k,i_{l-1}}} W_{n+1}^{k,i_{n}} \cdots W_{l+2}^{k,i_{l+1}} W_{l+1}^{k,i_{l}} \sigma_{l}(1) 1_{\mathcal{RST}_{l,n}^{k,i_{l-1},\ldots,i_{n}}} \right]
$$
\n
$$
= \sigma_{l-1} (\alpha_{l} \overline{\nu}_{l} \circ \alpha_{l} K \Gamma_{l,n+1}^{k})
$$
\n
$$
= E^{Y} \left[ W_{1,n+1}^{k,1} 1_{\overline{\mathcal{R}}_{l}^{k}} (\overline{\mathcal{R}}_{l+1}^{k})^{C} \cdots (\overline{\mathcal{R}}_{n}^{k})^{C} \right]
$$
\n
$$
(9.17)
$$

and  $\{\mathbb{T}_1^k\}_{k=1}^{m_N}$ , the first terms of [\(9.9\)](#page-45-0), are i.i.d., bounded (w.r.t.  $E^Y$ ) and zero mean. Therefore, it follows as above by the Marcinkiewicz-Zygmund and Jensen inequalities that

$$
E^{Y} \left| \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \mathbb{T}_{1}^{k} \right|^{p} \ll m_{N}^{-\frac{p}{2}}
$$
\n(9.18)

for any  $p \geq 1$ .

**Bound**  $\overline{T_2}$ : One has by the induction hypothesis, [\(7.47](#page-30-3)[,7.36,7.37\)](#page-29-2) and Jensen's inequality that for any  $p\geq 1$ 

$$
E^{Y} \left| \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \mathbb{T}_{2}^{k} 1_{\mathbb{D}_{n}^{N}} \right|^{p} \ll E^{Y} \left| \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \left( \mathbb{A}_{l}^{m_{N}} - \sigma_{l}(1) \right) 1_{\mathbb{D}_{l-1}^{N}} \right|^{p} \ll m_{N}^{-\frac{p}{2}}.
$$
\n(9.19)

<span id="page-47-1"></span>**Bound**  $\mathbb{T}_3$ : One finds by [\(7.47\)](#page-30-3) that

$$
\sum_{k=1}^m \mathbb{T}_3^k \left| 1_{\mathbb{D}_n^N} \right. \tag{9.20}
$$
\n
$$
\ll \sum_{k=1}^m \sum_{i_{l-1}=1}^{N_{k-1}^k} \sum_{i_l=N_l^{k,i_{l-1}-1}+1}^{N_{l-1}^k} \left| 1_{\mathbb{R}\mathbb{S}_{l,n}^{k,i_{l-1},\ldots,i_n} - 1_{\mathcal{R}\mathcal{S}_{l,n}^{k,i_{l-1},\ldots,i_n}} \right| 1_{\mathbb{D}_{l-1}^N}
$$
\n
$$
\leq \sum_{j=l+1}^n \sum_{k=1}^{N_N} \sum_{i_{l-1}=1}^{N_{k-1}^k} \sum_{i_l=N_l^{k,i_{l-1}-1}+1}^{N_{k,i_{l-1}-1}^k} 1_{\mathbb{R}_l^{k,i_{l-1}} \leq \sum_{j=1}^{k,i_{l-1}} \leq \mathcal{S}\Delta_{j}^{k,i_{j-1}} \mathcal{S}_{j+1}^{k,i_{j-1}} \cdots \mathcal{S}_{j+1}^{k,i_{l-1}} \cdots \mathcal{S}_{j+1}^{k,i_{l-1}} \cdots \mathcal{S}_{j+1}^{k,i_{l-1}-1} \mathbb{I}_{\mathbb{D}_{l-1}^N}
$$
\n
$$
+ \sum_{k=1}^{m_N} \sum_{i_{l-1}=1}^{N_{k-1}^k} \sum_{i_l=\overline{N}_l^{k,i_{l-1}-1}+1}^{N_{k,i_{l-1}-1}^k} 1_{\mathbb{R}\Delta \mathcal{R}_l^{k,i_{l-1}} \leq \mathcal{S}_{l+1}^{k,i_{l-1}} \cdots \mathcal{S}_{n}^{k,i_{n-1}} \mathbb{I}_{\mathbb{D}_{l-1}^N}^N,
$$
\n(9.20)

where

$$
\mathbb{S} \,\triangle \mathcal{S}_j^{k,i_{j-1}} = \mathbb{S}_j^{k,i_{j-1}} \,\triangle \,\mathcal{S}_j^{k,i_{j-1}} \text{ and } \mathbb{R} \,\triangle \mathcal{R}_j^{k,i_{j-1}} = \mathbb{R}_j^{k,i_{j-1}} \,\triangle \,\mathcal{R}_j^{k,i_{j-1}}.\tag{9.21}
$$

<span id="page-47-0"></span>Recalling  $\mathcal{G}_k^j$  from [\(7.52\)](#page-31-2), one has that

$$
E^{Y} \left| \sum_{k=1}^{m_{N}} \sum_{i_{l-1}=1}^{N_{l-1}^{k}} \sum_{i_{l}=\overline{N}_{l}^{k,i_{l-1}}}^{N_{l}^{k,i_{l-1}}} 1_{\mathbb{R}_{l}^{k,i_{l-1}} \leq \mathbb{S}_{l+1}^{k,i_{l-1}} \leq \mathbb{S}_{l}^{k,i_{l-1}} \leq \mathbb{S}_{l+1}^{k,i_{l-1}} \leq \mathbb{S}_{l+1}^{k
$$

for  $j = l, ..., n$ , where

$$
\Delta_{j}^{k} = \n\sum_{i_{l-1}=1}^{N_{l-1}^{k}} \frac{\overline{N}_{l}^{k,i_{l-1}}}{\sum_{i_{l-1}=1}^{k}} 1_{\mathbb{R}_{l}^{k,i_{l-1}} \leq s_{l+1}^{k,i_{l-1}} \leq \Delta S_{j}^{k,i_{j-1}} \leq \Delta S_{j}^{k,i_{j-1}} \leq s_{n}^{k,i_{n-1}} \mathbb{D}_{l-1}^{N}}\n- E^{Y} \left[ \sum_{i_{l-1}=1}^{N_{l-1}^{k}} \sum_{i_{l}=\overline{N}_{l}^{k,i_{l-1}}}^{N_{l}^{k,i_{l-1}} \mathbb{D}_{l}^{k,i_{l-1}} \leq s_{l+1}^{k,i_{l-1}} \leq s_{n}^{k,i_{j-2}} \mathbb{E}_{S_{j}^{k,i_{j-1}} \leq s_{n}^{k,i_{j}} \cdots S_{j+1}^{k,i_{n-1}} \mathbb{D}_{l-1}^{N}} \right] \mathcal{G}_{k-1}^{j}
$$
\n
$$
= E^{Y} \left[ \sum_{i_{l-1}=1}^{N_{l-1}} \sum_{i_{l}=\overline{N}_{l}^{k,i_{l-1}-1} + 1}^{N_{l}^{k,i_{l-1}} \mathbb{D}_{l}^{k,i_{l-1}} \leq s_{n}^{k,i_{j-2}} \mathbb{E}_{S_{j}^{k,i_{j-1}} \leq s_{n}^{k,i_{j}} \cdots S_{n}^{k,i_{n-1}} \mathbb{D}_{l-1}^{N}} \right] \mathcal{G}_{k-1}^{j}
$$
\n(9.23)

<span id="page-48-0"></span>are bounded  $\{\mathcal{G}_k^j\}$ -martingale differences (in k). Therefore, it follows by the Burkholder-Gundy-Davis inequality and Jensen's inequality as well as exchangeability that

$$
E^{Y} \left| \sum_{k=1}^{m_{N}} \Delta_{j}^{k} \right|^{p} \ll E^{Y} \left| \sqrt{\sum_{k=1}^{m_{N}} (\Delta_{j}^{k})^{2}} \right|^{p}
$$
\n
$$
\ll m_{N}^{\frac{p}{2}-1} \sum_{k=1}^{m_{N}} E^{Y} \left[ |\Delta_{j}^{k}|^{p} \right] = m_{N}^{\frac{p}{2}} E^{Y} \left[ |\Delta_{j}^{1}|^{p} \right] \ll m_{N}^{\frac{p}{2}}
$$
\n(9.24)

<span id="page-48-1"></span>for  $p \ge 2$ . Now, by Hölder's inequality we can take p to be an integer. Moreover, by [\(7.36,7.37\)](#page-29-2)

$$
\sum_{i_{l-1}} \sum_{i_l} E^Y \bigg[ 1_{\mathbb{R}_l^{k, i_{l-1}} \mathbb{S}_{l+1}^{k, i_l} \cdots \mathbb{S}_{j-1}^{k, i_{j-2}} \mathbb{S} \triangle \mathcal{S}_j^{k, i_{j-1}} \mathbb{S}_{j+1}^{k, i_{j}} \cdots \mathbb{S}_n^{k, i_{n-1}}} \bigg| \mathcal{G}_{k-1}^j \bigg] \mathbf{1}_{\mathbb{D}_{l-1}^N}
$$
\n
$$
\ll 1_{(\mathbb{D}_{j-1}^N)^C} + \sum_{i \in \mathbb{I}_{j-1}^k \cup \mathcal{I}_{j-1}^k} E^Y \bigg[ \mathbf{1}_{\mathbb{S}_j^{k, i} \triangle \mathcal{S}_j^{k, i}} \bigg| \mathcal{G}_{k-1}^j \bigg] \mathbf{1}_{\mathbb{D}_{j-1}^N}
$$
\n(9.25)

<span id="page-48-2"></span>and one has by [\(7.54\)](#page-32-1) that

$$
E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{j-1}^{k}\cup\mathcal{I}_{j-1}^{k}}E^{Y}\left[1_{\mathbb{S}_{j}^{k,i}\triangle S_{j}^{k,i}}\middle|\mathcal{G}_{k-1}^{j}\right]\right|^{p}1_{\mathbb{D}_{j-1}^{N}}\right]
$$
\n
$$
\ll E^{Y}\left[\left|\mathbb{A}_{j}^{m_{N}}-\sigma_{j}(1)\right|^{p}1_{\mathbb{D}_{j-1}^{N}}\right]+E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{j-1}^{k}\cup\mathcal{I}_{j-1}^{k}}\left|\widehat{\mathcal{K}}_{j-1}^{k,i}-\widehat{\mathbb{R}}_{j-1}^{k,i}\right|\right|^{p}1_{\mathbb{D}_{j-1}^{N}}\right]
$$
\n(9.26)

<span id="page-49-0"></span>so by [\(9.22\)](#page-47-0), [\(9.24\)](#page-48-0), [\(9.25\)](#page-48-1), [\(9.26\)](#page-48-2) and [\(7.35\)](#page-29-2)

$$
E^{Y} \left| \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \sum_{i_{l-1}} \sum_{i_{l}} 1_{\mathbb{R}_{l}^{k,i_{l-1}} \mathbb{S}_{l+1}^{k,i_{l}} \cdots \mathbb{S}_{j-1}^{k,i_{j-2}} \mathbb{S} \Delta \mathcal{S}_{j}^{k,i_{j-1}} \mathbb{S}_{j+1}^{k,i_{j}} \cdots \mathbb{S}_{n}^{k,i_{n-1}}} \right|^{p} 1_{\mathbb{D}_{l-1}^{N}}
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}} + \sum_{j=l+1}^{n} E^{Y} [|\mathbb{A}_{j}^{m_{N}} - \sigma_{j}(1)|^{p} 1_{\mathbb{D}_{j-1}^{N}}]
$$
\n
$$
+ \sum_{j=l+1}^{n} E^{Y} \left[ \left| \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \sum_{i \in \mathbb{I}_{j-1}^{k} \cup \mathcal{I}_{j-1}^{k}} |\widehat{\mathcal{K}}_{j-1}^{k,i} - \widehat{\mathbb{R}}_{j-1}^{k,i}| \right|^{p} 1_{\mathbb{D}_{j-1}^{N}} \right].
$$
\n(9.27)

<span id="page-49-1"></span>Similarly to [\(9.22](#page-47-0)[-9.27\)](#page-49-0), one finds that

$$
E^{Y} \left| \sum_{k=1}^{m_{N}} \sum_{i_{l-1}=1}^{N_{l-1}^{k}} \sum_{i_{l}=\overline{N}_{l}^{k,i_{l-1}}}^{N_{l}^{k,i_{l-1}}} 1_{\mathbb{R}\Delta\mathcal{R}_{l}^{k,i_{l-1}}\mathcal{S}_{l+1}^{k,i_{l}}\cdots\mathcal{S}_{n}^{k,i_{n-1}}} \right|^{p} 1_{\mathbb{D}_{l-1}^{N}}
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}} + E^{Y} [|\mathbb{A}_{l}^{m_{N}} - \sigma_{l}(1)|^{p} 1_{\mathbb{D}_{l-1}^{N}}]
$$
\n
$$
+ E^{Y} \left[ \left| \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \sum_{i \in \mathbb{I}_{l-1}^{k} \cup \mathcal{I}_{l-1}^{k}} |\widehat{\mathcal{K}}_{l-1}^{k,i} - \widehat{\mathbb{K}}_{l-1}^{k,i} \right|^{p} 1_{\mathbb{D}_{l-1}^{N}} \right].
$$
\n(9.28)

Therefore, one has by  $(9.20)$ ,  $(9.27)$ ,  $(9.28)$  and the lemma hypothesis that

$$
E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\mathbb{T}_{3}^{k}\right|^{p}1_{\mathbb{D}_{n}^{N}}\right]
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}} + \sum_{j=l}^{n} E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\sum_{i\in\mathbb{I}_{j-1}^{k}\cup\mathcal{I}_{j-1}^{k}}|\widehat{\mathcal{K}}_{j-1}^{k,i} - \widehat{\mathbb{K}}_{j-1}^{k,i}|\right|^{p}1_{\mathbb{D}_{j-1}^{N}}\right]
$$
\n(9.29)

for any  $p \geq 1$ . **Bound**  $\mathbb{T}_4$ : We find by [\(7.47\)](#page-30-3) and analogous to [\(9.20-](#page-47-1)[9.25\)](#page-48-1) that

$$
E^{Y} \left[ \left| \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} \mathbb{T}_{4}^{k} \right|^{p} 1_{\mathbb{D}_{n}^{N}} \right]
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}} + E^{Y} \left[ \left| \sum_{j=l}^{n} \frac{1}{m_{N}} \sum_{k=1}^{m_{N}} E^{Y} \left[ 1_{\mathbb{E}_{j}^{k} \triangle \mathcal{R}_{j}^{k}} \middle| \mathcal{G}_{0}^{j} \right] \right|^{p} 1_{\mathbb{D}_{j-1}^{N}} \right].
$$
\n(9.30)

Hence, by [\(7.56,7.55\)](#page-32-1) and the lemma hypothesis

$$
E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}\mathbb{T}_{4}^{k}\right|^{p}1_{\mathbb{D}_{i-1}^{N}}\right]
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}} + \sum_{j=l}^{n} E^{Y}[\left|\mathbb{A}_{j}^{m_{N}} - \sigma_{j}(1)\right|^{p}1_{\mathbb{D}_{j-1}^{N}}]\n+ \sum_{j=l}^{n} E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}|\hat{\mathcal{L}}_{j-1}^{k} - \hat{\mathbb{K}}_{j-1}^{k}\right|^{p}1_{\mathbb{D}_{j-1}^{N}}\right]
$$
\n
$$
\ll m_{N}^{-\frac{p}{2}} + \sum_{j=l}^{n} E^{Y}\left[\left|\frac{1}{m_{N}}\sum_{k=1}^{m_{N}}|\hat{\mathcal{L}}_{j-1}^{k} - \hat{\mathbb{K}}_{j-1}^{k}\right|^{p}1_{\mathbb{D}_{j-1}^{N}}\right]
$$
\n(9.31)

for any  $p \geq 1$ .  $\Box$ 

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