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SARAH LOUISE BUTSON		
Date of Birth — Date de naissance	Country of Birth — Lieu de naissance	
FEBRUARY 17. 1954	CANADA	
Darmanant Address (Décidance five	AIREPTA	
11335-88 ST EDMONTON	HERENTILL AS BLOCKTA	
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BIOSTATISTICAL ANALYSIS OF FACTORS INFLUENCING LACTATION
PERFORMANCE OF RANGE COWS AND WEANING WEIGHTS OF THEIR
CALVES

SARA

SARAH LOUISE BUTSON

by

A THESIS

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OF MASTER OF SCIENCE

IN

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The undersigned certify that they have read, and recommend to the Faculty of Graduate Studies and Research, for acceptance, a thesis entitled BIOSTATISTICAL ANALYSIS OF FACTORS INFLUENCING LACTATION PERFORMANCE OF RANGE COWS AND WEANING WEIGHTS OF THEIR CALVES submitted by SARAH LOUISE BUTSON in partial fulfilment of the requirements for the degree of MASTER OF SCIENCE in ANIMAL GENETICS.

R.D. Berg.
Supervisor
Molet T. Harli

"Science must not be a selfish pleasure.

Those who are so fortunate as to be able to pursue it should be the first to put their knowledge at the service of mankind"

Karl Marx

Economic and Philosophic Manuscripts of
1844

ABSTRACT

An investigation was carried out to examine trends in lactation performance and factors influencing milk traits and calf weaning weight among range beef cattle.

Cows representing four beef and dairy-beef breeds and lines from the University of Alberta beef herds ranging in age from 2 to 10 years of age were used in the study. June and September measurements of milk and constituent percentages were taken on 206 and 242 cows in 1976 and 1977 respectively.

Age and breed of dam, and age and sex of calf were recorded. Other variables examined were cow weight changes during gestation and lactation and birthweights, weaning weights and preweaning ADG of their calves.

Milk variables and calf weaning weights were the traits studied.

In Chapter I the lactation performance of the four breed groups of dams was examined. Results indicated that crossbred cows with dairy breed ancestry yielded more milk and were more persistent than a purebred traditional beef breed.

Trends between breeds for constituent percentage production were less noticeable although the dairy crossbreds tended to produce less butterfat% than the purebreds.

In Chapter II factors influencing milk and constituent yields were investigated using the multiple regression

technique in Chapter II. Stepwise regressions involving age and breed of dam accounted for between 34.0% and 44.0% of the variance in any of the milk or constituent yields.

Full and restricted regression models allowed the introduction of independent variables after adjusting for age and breed of dam effects. Calf preweaning ADG showed a high association with milk yield variables. Other factors such as calf birthweight and cow weight changes however did not account for a significant proportion of the variance in any one milk or constituent yield variable.

Bécause associations and correlations between milk yield and calf ADG were significant but only moderately high, all factors influencing weaning weights of range calves were examined in Chapter III. Regression analysis indicated that age and breed of dam effects accounted for an average of 47% of the variance in calf weaning weight. When age and breed of dam were not considered, milk or constituent yields accounted for approximately 40% of weaning weight variance. However, milk variables still accounted for up to 10% of the variance after accounting for the effects of age and breed of dam.

Calf birthweight and cow weight changes had minor effects on calf weaning weight.

As a significant proportion of the variance in weaning weight was accounted for by milk variables, which in turn are largely a reflection of the breed of dam, it was concluded that selection for increased lactation performance

and the introduction of dairy breeds into the breeding program of a beef herd can effect meaningful increases in the weights of calves weaned.

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I. LACTATION PERFORMANCE OF RANGE BEEF AND DAIRY-BEEF COWS

A. INTRODUCTION

The effect of milk yield on preweaning growth response of range cattle has been reported extensively (Long, 1980). Studies have indicated that measurements of milk yield and milk constituents of beef cows may serve as useful predictors of calf growth (Gleddie and Berg, 1968; Bluntzer and Sims, 1976). However, few workers have examined differences in lactation performance among various breeds of beef and dairy-beef cattle. Gaskins and Anderson (1980) suggest that knowledge of beef lactation trends may provide information on changing energy requirements involved in both cropping and herd management decisions. Such comparisons may also be useful as predictors of variation between breeds in levels of total energy consumption by calves, which in turn possibly affect growth response of suckling calves.

Selection criteria in dairy programs have emphasized total milk volume and persistency rather than constituent percentage production (Schmidt, 1971). Based on the foregoing and the negative correlation that exists between milk yield and constituent percentages (Preston and Willis, 1974), it is expected that beef cows with dairy backgrounds may produce a lower percentage of total solids than traditional beef breeds.

As with dairy cattle, lifetime performance of beef cows is an important criterion of production, yet few

experimental regimes with groups of various breeds have been maintained over several breeding seasons for observations (Richardson et al., 1977). According to Bluntzer and Sims (1976) quantifying milk production could help to determine the most efficient types of cow-calf units for utilizing and converting range forage into high-protein human foods.

Some authors (Rutledge et al., 1970) contend that milk quantity rather than quality is more important for beef calves. However, to date there is little information concerning the constituent percentage differences between breeds of range cattle.

The purpose of this study was to determine the levels of milk yield and milk constituent percentages and to examine the seasonal variation in lactation traits of four lines or breeds of range beef and dairy-beef cows. The inter-relationships of milk variables were also examined.

B. MATERIALS AND METHODS

The breeding plan and general management of the herd have been described in detail by Berg (1978). Cows for breeding purposes have been maintained on native shortgrass range throughout the year with supplementary feed provided during the winter. Generally the policy of winter feeding is to provide the minimum pount of energy consistent with reasonable herd health. A summary of the December to March winter feeding plan from 1963 to 1975 is provided by Berg (1975).

First-calf heifers were fed to gain a moderate amount of weight over winter (10 to 20 Kg) prior to calving. Older cows were fed to maintain their precalving weight during gestation. Heifers were first bred in July and August to calve as two-year olds. All cows and heifers two years of age and older failing to calve each year were culled. First-calf heifers calved in a semi-enclosed feedlot area while all other cows calved on open range. Calves were born in April and May and remained with their dams with no creep feed. Calves of cows used in this study were weaned in the first week of October, averaging 156 and 158 days of age in 1976 and 1977 respectively.

Winter temperatures from October to March for 1975 to 1976 averaged -6°c and -5°c respectively. The winter of 1976-1977 was relatively mild with 29 days below -18°c compared with 49 days in 1975-1976. Summer precipitation, measured from April to October, was similar for both years, averaging approximately 32.0 cm. However, in 1976 the heaviest rainfall was in June and July compared with early May in 1977. Pasture conditions were therefore slightly less favorable in 1977.

Experiments relating milk production and cow size to efficiency and growth rate were conducted from 1964 through 1966 at the University of Alberta Research Ranch in Kinsella, Alberta (Peschiera, 1966; Gleddie and Berg, 1968). Relationships between milk yields and calf weaning weight were also studied. The lactation performance of purebred

Herefords was compared with that of various breed crosses of Charolais, Angus and Galloway cows.

The cows were from the experimental beef herd at The University of Alberta Research Ranch in Kinsella, and represented Hereford (HE), Beef-Synthetic (SY), Dairy-Beef (DB) and Dairy-Synthetic (DS) breed groups. The HE group is a purebred population open to artificial insemination from superior industry bulls selected on the basis of performance or progeny tests. The SY population, established in 1960, is predominantly a composite of Charolais, Angus and Galloway breeding. The DB is a crossbred group resulting from mating HE and SY cows to purebred dairy breed bulls. The DS line was begun in 1967 and is a composite of approximately 30% Holstein, 30% Brown Swiss and the rest traditional beef breeds (Berg, 1975).

A milking experiment similar to the one conducted from 1964 to 1966 (Gleddie, 1968) was repeated in 1976 and 1977. Measurements of milk were taken over 4 days each in June and September at an average of 44 and 130 days in lactation. Data were recorded over a two-year period for June and September averages of daily yields of milk and constituent percentages of butterfat (BF), protein (PROT) and lactose (LACT) for 448 animals. Cow age and breed were also recorded.

The method of milking was similar to that outlined by Jeffery and Berg (1971) using oxytocin and teat tubes for manual removal of milk. No machine or hand milking was

involved. Milking commenced at 7am immediately following separation of the cows and calves. Cows were restrained in a commercial stock squeeze and milk letdown was induced using an intrajugular injection of 20 International Units of oxytocin. Teat tubes were inserted into all quarters and milk flow began within 15 seconds following injection. No further oxytocin injections were administered as residual milk following a second injection in a trial sample of 30 cows was negligible. Cows exhibiting mastitis in any quarter were not milked. Six hours later the same procedure was repeated and the milk weighed.

The 6-hour milk yield was multiplied by the interval in minutes between the two milkings to estimate 24-hour milk yield assuming a constant rate of milk secretion. Samples from each cow were collected and 5 grams of Potassium Dichromate were added as a preservative. Samples were analyzed at the Alberta Provincial Central Milk Testing Laboratory for percent butterfat using an Infra-Red Milk Analyzer. Protein percent was determined using the Kjeldahl Method as outlined by Bradstreet (1965). Solids-not fat (SNF) percent was determined using the Golding Bead Test outlined by Golding (1964). Lactose percent was then estimated by subtracting protein percent from SNF percent. Results for lactose percent may therefore be slightly higher than average values as the small proportion of ash content would be included in the calculations.

In 1976 the experiment involved a total of 206 cows consisting of 45 HE, 102 SY, 26 DB and 33 DS. In 1977 a total of 242 cows were involved comprising 58 HE, 123 SY, 33 DB and 28 DS. Although approximately two-thirds of the cows tested in 1977 had been included in the milking experiment the previous summer, for statistical analyses cows in one year were considered to be different from cows in the other year. Cows ranged in age from 2 to 10 years.

Owing to the large number of available SY cows, random selection was made in this line from lactating dams nursing their own calves while all available dams from the HE, DB and DS groups were used.

Statistical Analyses

Least squares analyses of covariance for unequal subclass numbers (Mehlenbacher, 1978) were computed with the effects of breed of dam, age of dam, breed x age of dam interaction and sex of calf as sources of variation, and age of calf as a covariate.

Levels for the main effects were:

- 1. Breed of dam (B) classified as HE, SY, DB and DS;
- Age of dam (A) classified as 2, 3, 4 and mature (over 4 greats of age);
- Sex of calf (S) classified as male and female;
- 4. Calf age (CA), the covariate was recorded as days of age.

Least squares constants for A, B, A x B and S were computed and used to calculate least squares means for milk

variable averages of June, September and overall averages of milk yield, BF%, PROT% and LACT% in the following model:

MVijkl = u + Ai + Bj + ABij + Sk + blXijkl + eijkl where MVijkl = adjusted milk variable of the ith age of dam, jth breed of dam and kth sex of calf.

u = overall population mean for x=0.

Ai = effect of the ith age of dam.

Bj = effect of the jth breed of dam.

ABij = effect of the interaction of the ith age of dam and jth breed of dam.

Sk = effect of the kth sex of calf. /

blXijkl = partial regression of the 1th milk variable on the age of calf.

eijkl = random error.

Bonferroni's t-statistic as outlined by Kirk (1968) was used to test differences between individual means when significant differences were established by least squares analysis.

Phenotypic correlations were computed over all the data sets.

Measurement of Milk Yield Using Oxytocin

Studies involving measurements of lactation in range cattle are limited primarily because of technical difficulties involved in milk removal (Richardson et al., 1977). Furthermore there exists some controversy as to whether milk taken from range cows provides a good estimate of either the expression of breed potential or individual

dam capacity for milk production.

Several experiments have been conducted to determine the efficacy of using exogenous oxytocin to estimate milk yield or calf ingestion. Sibaja and Schmidt (1975) found that injection of oxytocin (40 U.S.P.) did not interfere with normal milk ejection. However, Schwulst et al., (1966) noted that whereas oxytocin did not significantly affect milk composition, its consistent administration tended to result in higher milk consumption by the calf and an increase in total milk production from the second to fifth months of lactation. Whereas Thompson et al., (1973) found that amounts of residual milk were highly variable among Holsteins in mid-lactation, Schwulst et al., found that average residual milk declined from 15% of the total in the first 2 to 3 weeks to 6% of total yield at any given milking. These results are similar to the 9 to 10% residual milk following suckling in the first months of lactation with Herefords as observed by Bluntzer and Sims (1976) but differ with the 25% residual observed by Thompson et al., (1973). These workers reported that a 20 I.U. dose of oxytocin was required in Holsteins in mid-lactation to release 75% of total milk contained. Hanjra et al., (1977) noted that the differences in residual milk between 10 I.U. and 20 I.U. injections in lactating buffalo were not significant. Hanjra et al., collected from 0.79 to 5.56 Kg per week of residual milk.

Hall (1971) reported that when oxytocin was injected every two weeks carves began taking all of the available milk by the 10th week. By the 16th, no residual milk was obtained. Hall suggested therefore that the maximum demand of the calf coincides with total milk decline. According to Sibaja and Schmbdt (1975) it is generally accepted that administration of oxytocin in the first 5 weeks post-partum will not be a true measure of calf ingestion.

It is not clear whether the large variability in residual milk percentages of total milk is due to breed differences in hormone levels or due to differences in experimental methods. According to literature reviewed by Hart et al., (1975), some authors found significantly higher levels of growth hormones and non-esterified fatty acids, but lower concentrations of prolactin, insulin and glucose in the circulation of dairy cows compared with beef cows. However, Cowie (1976) reported no significant differences in prolactin levels between Freisians and Hereford x Sussex throughout lactation.

In the present study, a small number of cows were given additional injections of oxytocin approximately 15 minutes after the completion of the initial milk letdown. Negligible amounts of milk were collected but were not weighed. It was assumed that milk extracted using the single 20 I.U. injection served as an approximation of the 6-hour milk yield of the dam.

It is likely in this experiment that June milk yields extracted at approximately 44 days of lactation will represent a smaller proportion of total milk contained but will be greater than the amount ingested by the calf. Although early yields may not correspond as directly to calf ingestion, they may reflect more accurately the inherent genetic potential of the milk yield capacity of the dam. However, by September at an average of 130 days, individual milk yields will more probably correspond to levels of calf ingestion.

C. RESULTS AND DISCUSSION

Average Milk Yields

Least square means and standard errors of average daily milk yields by breed and breed-age groups of dam are presented in Table I.1 for 1976 and 1977. The averages for 24-hour milk yields in 1976 and 1977 were respectively 6.9± 0.1 and 7.1± 0.1 Kg/day over all breed and age groups examined, ranging from 4.5±0.3 Kg/day for 2-year old HE dams in 1977 to 9.3±0.4 Kg/day for mature DS dams in 1976. The difference between the 1976 and 1977 yields was significant and may be attributed to variation in weather conditions. The 1976 data showed an unexpected low production for DB dams which was possibly biased by the low yield for the single mature DB observation.

I.i. Least aquares means and standard errors of sverage daily milk yields by breed and breed-age proups of dam. Kinesis, ozz.,ozz.

	٠.			Hilk yiel	Hilk yield (Kg/day)		
			1976		ļ }	1.61	
	Breed	<u>§</u>	1	35	£	T C	35
Grand total	. A11 .	20 20 20 20 20 20 20 20 20 20 20 20 20 2	9.9	0.1	242	7.100	0
	¥	5	5.7	0.0	3	8	c
	٠	102	8	0	123		
	8 2	۲ ;	<u>.</u>	7.0	2	. to	0
	3	3		6 .0	59	7.70	0 8.0
Age of dam (veers)							
'n	14	7.8	6.2v	0.3	82	\$	0.7
	Ĭ	=		•	:	•	•
	SY	36	. 7	0	e g		9 6
	8	t	7.9	0			
	SO	<u>.</u>	۲. ۲	6 .0	•	6.7	6
6		96	9 . 6 .	0.3	\$2	7.2×	0.3
•	¥	-	D	9.0	1		•
	<u>ئ</u> د	ē.	m i	6.0	8	6.7	0
	s s	• •		• se	v , a		00
•	LIA	8	7.5x	0.3	*	7.4×	
	. ±	•	•	•	•	: (;
	S	ō	-	•	`:	P) •	0
	8		*		•	•	• i
	08	6	6.7	0.7	• ~		
Mature	A11	70	7.1x	0.3	2	7.9x	0.3
	¥	22	9	0.3	86		•
	, S,	7	0.	0.7	; ‡		
	8	-	4.7	7.7	•	•	
	2	•	•				•

1977 average daily milk yield significantly different from 1976 (Pc0.01).

Breed of Dam Differences

The average values for HE and SY of 5.8 and 6.9 Kg/day respectively are within the range reported by Gleddie and Berg (1968), Hall (1971) and Nicol (1976) for traditional beef breeds (Table I.2) but are higher than yields of Herefords noted by Kress and Anderson (1974).

Average daily milk yields for the two dairy groups were and 8.0 Kg/day for DB and 7.9 and 7.7 Kg/day for DS in each year respectively. These values considerably lower than the 14.0 and 19.0 Kg/day averages of commercial dairy cows including Jersey, Ayrshire, Holstein and Brown Swiss (Schmidt, 1971; Preston and Willis, 1974; Cerbulis and Farrell, 1975) (Table I.2).

Table I.3 presents the analyses of covariance of average daily milk yields with calf age as the covariate. The F values indicate that the effects of both breed and age of dam accounted for a significant (P<0.01) source of variation in average milk yield for both years, similar to results reported by Gaskins and Anderson (1980).

As illustrated in Figure I.1 the DS yielded more milk than other breeds in 1976. However, DS yields in 1977 were not significantly different from yields of DB cows (Table I.1). Although it is likely that the single low yield recorded for the DB mature dam in 1976 biased the breed average for the DB group, it is not clear whether the DB and DS have the same genetic potential for milk production. The breeding background of the DS (approximately one-third

Table I.2. Average daily milk yields: literature breed averages.

	AVE. MIJK	Ave. Milk Yield (Kg/day)	•
Breed	Ave. 305	Ave. 305 - day lactation	Source
A. DAIRY		``	
Holstein		19.5	Cerbulis and Farrel, 1973
Brown Swiss		19.9	j,,de³
Milking Shorthorn		14.8	Schwidt, 1971
Jersey		IO	
2 3 3 3 3 3 3 3 3 3 3 3 3 3 3 3 3 3 3 3	2-6 W##KB	200-210 days	
. Hereford	ID ID	1.7	Kress and Anderson, 1974
Angus	7.9	2.5	Gleddie and Berg, 1968
Shorthorn	6.8	2.3	Notter, 1967
Angus x Hereford	7.4	4.7	Gaskins and Anderson, 1980
Charolais	7.4	3.4	•

Table 1.3. Analyses of covariance of average daily milk yields, Kinsella 1976-1977.

			F TESTS	: S1
	d.f.		Average dafly milk	fly milk
Source	1976	1977	1976	1977
Age of dam	3	m	5.65**	17.08**
Breed of dam	. m	ю	15,90**	14.17**
Age x Breed	6	້	2.11*	0.31
Sex of calf	-	₩.	0.0	0.33
Age of calf	-	-	2.60	0.0
Error	188	224		
Total	205	241		

Significant at P<0:05; ** Significant at P<0.01.

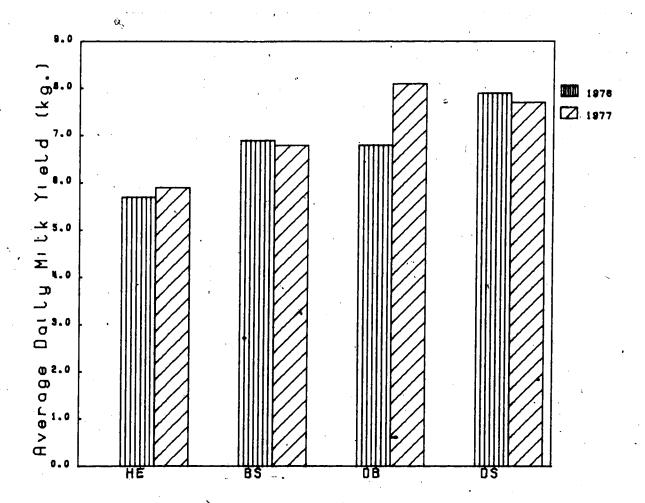


Figure I.1 Average milk yield by breed of dam, Kinsella 1976-1977

Holstein, one-third Brown Swiss and the remainder, traditional beef breeds), would suggest that calf suckling does not promote the full expression of the genetic potential of the dam.

Both the SY and DB groups produced approximately 1.0 and 1.5 Kg/day more than HE cows in 1976 and 1977 respectively, while the DS breed group yielded 2.2 and 1.8 Kg/day more than the HE respectively each year.

Results vary among reports. Marshall et al., (1976) found that reciprocal crosses of Angus and Charolais cows milked less than Angus purebreds, while Hall (1971) and McGinty and Frerichs (1971) reported higher yields for Charolais and Brown Swiss-Hereford crossbreds than purebred Herefords respectively. Deutscher and Whiteman (1971) found that 2-year old Angus-Holstein crossbreds produced from 0.9 to 2.7 Kg more milk daily than purebred Angus cows. Similar breed differences were reported by Gleddie and Berg (1968), Wilson et al., (1971), Rutledge et al., (1971), Notter (1976) and Gaskins and Anderson (1980); suggesting that in general the average daily milk yield tends to increase with the proportion of dairy breeding. Cruikshank et al., (1976) reported a 3-fold increase in the average annual income from the weaned calf and milk produced by various Friesian-beef crosses compared to purebred beef cows under range conditions.

It is not clear from results in this study whether differences in yields between the beef and dairy types are

due to the specific contribution of the ancestral backgrounds of the DS (Holstein, Brown Swiss and other) or whether the higher yields for all the crossbreds are due to heterotic effects of crossbreeding.

Age of Dam Differences

Least squares means of average daily milk yield by age of dam groups (Table I.1) are illustrated in Figure I.2. Averaged over both years they were 6.1, 7.0, 7.5 and 7.5 Kg/day for all 2, 3, 4-year old and mature dams respectively. Yearly differences between 1976 and 1977 were significant only for the 3-year olds (P<0.01).

Mature dams tended to yield significantly (P<0.01) more than 2- and 3-year olds although differences between 4-year olds and mature cows were not significant in either year (Table I.1). In 1976 3-year olds yielded approximately 0.4 Kg/day more than 2-year olds; 4-year olds produced 0.9 Kg/day more than 3-year olds and 0.4 Kg/day more than mature cows. In 1977 however, 3-year olds produced 1.2 Kg/day more than 2-year olds and only 0.2 Kg less than the daily yield of 4-year olds. Mature cows produced an averaged of 0.5 Kg/day more than 4-year olds.

The 1977 results are more nearly similar to data in published material than to the 1976 results. Gaskins and Anderson (1980) noted that there was a positive linear trend (b=1.0 Kg/year) in daily milk production as age of cow increased from two to four years.

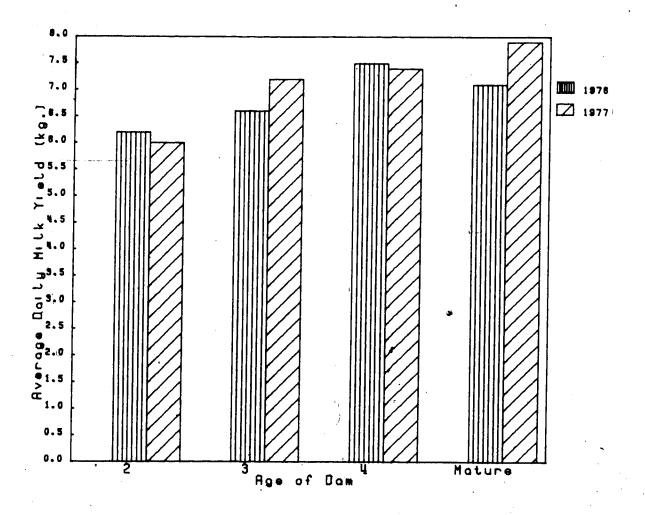


Figure I.2 Average milk yield by age of dam, Kinsella 1976-1977

Schmidt (1971) noted that for dairy cattle, more milk is produced with an increase in age due to the increased body weight including weight of the reproductive and digestive tract. Schmidt found that a 45.5 Kg increase in body weight was associated with an increase of 195.5 Kg total milk in a 310-day lactation. When body weight was held constant however, an increase in milk independent of body weight was expected until maturity. Ramirez and Porte (1976) reported a negative correlation between cow body weight and total milk yield, yet concluded that the first, second and third lactation number affected yield. Rutledge et al., (1970, 1971) reported quadratic effects of age of dam on milk yield in Herefords with a maximum of 8.4 years, while Notter (1976) found that average milk yield for 4-year old Herefords was 25% greater than for 3-year olds.

Christensen et al., (1973) found that age at first and second calving had more significant influences on yield than overall age effects alone. Neville et al., (1974) suggested that although milk yield increased for cows up to six years of age before reaching a plateau, lactation number may influence milk production as much as age of dam at calving. However, it is possible that in studies by both Neville et al., (1974) and Ramirez and Porte (1976), the lactation number is largely confounded by the age of dam effect. Seasonal Variation in Milk Yields

Lactation curves for Holstein cows based on 305-day lactations (Schmidt, 1971) compared with those for various

traditional beef breeds (Gleddie, 1968; Hall, 1971 and Gaskins and Anderson, 1980) are illustrated in Figure I.3. Milk yield for Holsteins peaks at about 3 to 6 weeks before remaining constant according to Schmidt (1971). Schmidt contends that pregnancy inhibits milk secretion by the 7th or 8th month of lactation when there is a sudden decline in milk. Ramirez and Porte (1976) and Wood (1972) found additional peaks corresponding to periods of flushes of grass growth or during periods of feeding silage when indoors, but the usual trend for dairy cattle is as outlined by Schmidt (1971).

It is generally thought that beef dams peak at about 4 weeks, but whether this is a breed difference between beef and dairy cows, confounded by the effect of calf suckling compared to regular machine milking, or simply due to differences in methods of measurement is not known. For example, Kress and Anderson (1974) found maximum production in 4- and 5-year old Herefords at 20 days (7.3 Kg/day) using the calf weigh-suckle-weigh method for measuring calf ingestion or milk production of the dam. However, Totusek et al., (1973) found that when beef calves were weighed before and after suckling, the lactation curve for beef cows was more nearly similar to the dairy, peaking at 7 compared to 4 weeks with hand-milking. Neidhardt (1979) and Ramirez and Porte (1976) found a similar peak in the second month of lactation among Brahman and Hereford cows respectively on range in Chile and Venezuela using the weigh-suckle-weigh

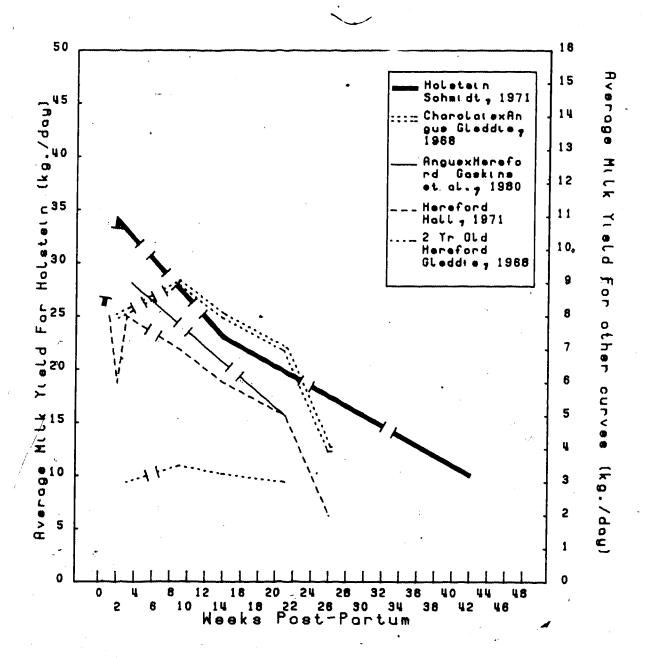


Figure I.3 Lactation curves of dairy and beef cows: literature breed averages

method.

Based on the foregoing, it is possible that the method of measurement might shift the milk yield curve. Klett et al., (1965) suggested that the range beef lactation curve is probably more flexible than that of dairy cows due to greater flexibility in milk production response to changing feed conditions. Deutscher and Whiteman (1971) observed that milk production curves of Angus-Holstein crossbreds on range paralleled range feed conditions.

Gleddie (1968) reported a decline in milk of various beef breeds from the first to fifth month (Figure I.3) which is different than dairy breeds with a rise from the first to the second month of lactation. He contended that the difference was due to the calf suckling effect. Hall (1971) noted a similar pattern, except that the Herefords he examined exhibited an initial high level which was lower than the maximum peak, followed by a sudden drop between 25 to 40 days, then a rapid decline (Figure I.3). Totusek et al., (1973) suggested that variation in estimates of milk yield within the first 30 days post-partum of Hereford, Angus and Shorthorn cows was indicative of limited calf capacity, while greater variation later was due to individual cow differences in persistency.

It is clear that based on published reports, no specific trend for beef cow lactation curves has been established. Gaskins and Anderson (1980) noted that lactation curves measured for Jersey, Angus, Hereford and

Simmental crossbreds were generally curvilinear throughout lactation, but were more convex for cows which had higher milk production - (ie: the 3- and 4-year olds and the Jersey a Angus and Simmental x Angus cows) and were more linear for cows with lower milk production.

June and September Milk Yields

Analyses of covariance for average June and September milk yields are presented in Table I.4. The effect of breed of dam accounted for a highly significant (P<0.001) source of variance in June and September milk yields each year. This effect is illustrated in Figure I.4.

Tables I.5 and I.6 show the least squares means of June and September milk yield averages for each year by age and breed of dam. Between breeds, trends for June and September yields were similar to those for the overall average daily milk yields.

In 1976, June yields were significantly higher (P<0.01) for the DS line, averaging 1.9, 0.9 and 1.4 Kg/day more than HE, SY and DS groups respectively. In June of 1977 however, differences between the SY, DB and DS groups were not significant although these breeds yielded an average of 1.3 Kg/day more than HE dams.

September yields were significantly (P<0.01) lower than June yields each year for all breeds.

Dairy-Synthetic cows produced approximately 2.9, 1.1 and 0.9 Kg/day more in September than the HE, SY and DB cows respectively.

Table 1.4. Analyses of covariance of average daily June and September milk yields, Kinsella 1976-1977.

see.	₽. D	·	June M	June M11k (1)	September	September Milk (1)
Source	1976	1977	1976	1977	1976	1977
Breed of dam	က	m	6.97***	4.31***	17.80***	30.18***
Age of dam	ო	e	2.88*	10.09**	6.78**	29.4100
Age x Breed	o	o	1.35	1.39	2.81**	1.67
Sex of calf	₩.	-	0.17	0.23	0.24	0.58
Age of calf	-	,-	1.07	0.86	2.95	3.87*
Error	188	224				
Total	205	241	•			a

** Significant at P<0.01, *** Significant at P<0.001. * Significant at P<0.05,

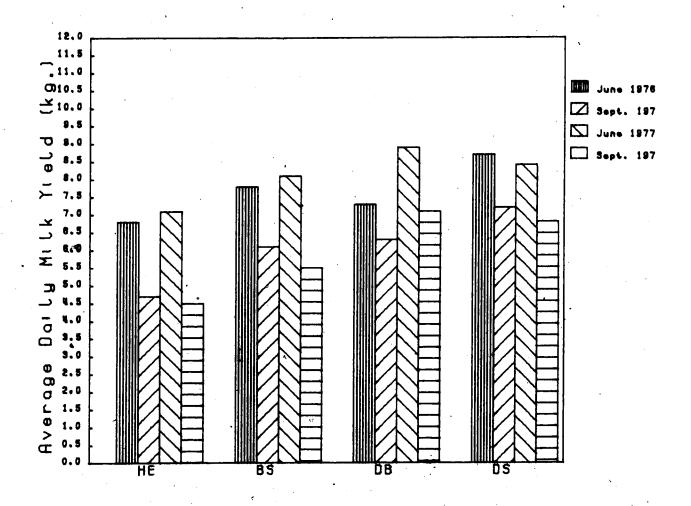


Figure I.4 June and September milk yields by breed of dam, Kinsella 1976-1977

Between ages, (Tables I.5 and I.6) mature dams tended to yield more than dams under 5 years of age in both milking periods, although differences were not always significant.

As measurements of milk variables were taken only in June and September, no lactation curves were extrapolated and prediction of trends was difficult. Figure I.4 illustrates the total differences between the two periods. The DB and DS dams demonstrated the highest persistency with total yields declining between 1.0 and 1.8 Kg between June and September. The HE dams were least persistent, averaging 2.5 Kg decreases each year between June and September.

Notter (1976) found significant breed differences between milk yields measured in beef cattle during all but the last month of lactation (176 days). In addition, the increase in persistency was inversely proportional to average milk yields measured from 121 to 176 days. This trend was noted for all but the low-yielding Charolais.

It is difficult to ascertain whether there are marked breed differences in lactation curve trends between beef and dairy types on range and if the differences in milk yields between HE and dairy types on range is always more noticeable at the end of the lactation. It is possible that the dairy-beef cows in this study were more persistent due to selection for 10-month lactations in their dairy breeding backgrounds, and because of greater stimulus due to the genotype of the calf.

Table I.S. Lesst squares means and standard errors of June and September average deliy milk yields, Kinselia 1976.

			eune .	2	September	- L	
	Breed	8	E .	35	£	SE	Difference
Grand Total	A11	206	7.6	0.2	-	0.1	
	뿦	£.	6.8	0.3	. 3	0.2	2 800
	Š	102	7.8b	0.5	.	0	- 1
	2 2	, 6	9	O	8	4.0	-0.
Age in years	Š	P)	9 . 10	F. 0	7.20	6.0 .3	
~	A11	75	7. tx	0.2	9.3x	0.0	
	¥	=		6	•	•	•
	λS	36	7.7	0		, ,) c
	8	E	7.2	•			
	S 0	Ç	7.0	0	7.	0	.
М	A11.	36	7.2×	6 .0	80 X	0.7	1.8.
	¥	7	7.0	•		· C	•
	Š	÷.	7.2	7			• •
	8	•	e.	•	0.	O	
•	80	9	7.0	0.1	.2	.0	0
•		20	8. fy	•	7.9	4.0	• • •
	¥	15	7.0	0.7	9		•
	S	2		8		•	. 6.
	2	n	. n	-		.0	.
	Š	e	.	•.0	0.	•. O	•
Mature	A11	75	8.2y	4.0		•	2.2.
	¥	33	7.7	0.9	9.0	0.9	2.4
	S	=	6.7	ю. О	7.3	0.0	7
	2	-	-	•	3.2		2.9
	2						

** Significant at P<0.01.

a.b.c * Least squares means of wilk yield by breed of daw within years with different slphabetic letters are significantly different (P<0.01).

 Least squares means of milk yield by age of dam within years with different alphabetic letters are significantly different (P<0.01)

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			erub.	2	Sept	September		
	Breed	8	E .	SE	Mean	ES EN	Difference	
Grand Total	II	242	8.2	0.2	0.	- 6	2.20	
	뿦	80	7.28	6		,		
	S.	123	9.10	0	5	, -	9.7	
٠.	28 2	. 33	.	0	7. 1c	- M	7.0.7	
	SO	59	8. 6b	7 .0	6.80	9.0		
Age in years	-							
	A11	82	Ø. 9	6.0	4. Bx	÷.	•	•
	¥,	9		,	•		•	
•	SY	100 100 100 100 100 100 100 100 100 100			,	F. 6	• - · -	
	2	2		?	•	0.0	2.2	
	80	•	. m			O	7	
				;		•	3.2	
6	A)	52	8.8y	0.3	5.87	0.3	2.8.	
	¥	7	0	6	•	,	•	
•	24	õ	. 7		, c	• •	F. 6	
	30	ø	•			, ,	P (
	Š	1 a	9	9 .0		•) r.	
•	Ŧ	36	8. 1y	8.0	6.62	6.0		
•	1		: •		1		•	
	.	7	P. 6		0	.	2.3	
	8		, e	e -	9		7.8	
	SO	. 4	. .	- -		- = 0 0	• • • •	
Ma ture	411	8 2	è	•			<u>.</u>	
		;		•	Z	0.3	2.2.	
	¥	25	8.2	4.0	9.7	6	. ,	
	<u>ن</u>	7	•	0.3	•	0.0		
	8 2	•	•	•	9.5		: -	
	20	đ	•				,:	

.. Significant at P<0.01

a.b.c = Least squares means of milk yield by breed of dem within years with different alphabetic letters are significantly different (P<0.04).

alphabetic latters are similarity age of dem within years with different

Differences in persistency between beef and dairy types may also be due to the fact that the genetic potential of dairy cows is more clearly expressed because of the method of milk removal. The high and consistent demand placed on dairy cows for 305-day lactations as compared with the relatively nominal demand placed on beef cows suckling their calves may account for the differences in trends. The genetic potential of even the high-producing older dairy-beef crossbreds in this study was not fully expressed, as measured with respect to milk yields. It is possible that the drop in persistency for all nursing cows becomes more dependent upon the individual cows' responses to suckling, rather than strictly a reflection of breed.

Age of dam also accounted for a significant (P<0.05) source of variance in June and September milk yields (Table I.4). As shown in Tables I.5 and I.6, among age groups mature animals tended to yield more in both June and September than dams under 5 years of age, although differences were not always significant. The results in this study indicate no consistent pattern for persistency trends for dam's age among range beef cattle.

Constituent Percentages

Average constituent percentages for traditional dairy and beef breeds are shown in Table I.7. According to Schmidt (1971) although there is a large variation in composition between breeds of dairy cattle, lactose and mineral content are not as variable and are not affected noticeably by

Table I.7. Average constituent percentages for dairy and beef breeds: literature breed averages.

1. 3. (1. 3. (1.) 3. (Constituent	Holstein	Jersey 1	Brown Swiss	Milking! Shorthorn	Angus i	=	
3.7 - 4.0 5.4 4.3 3.6 1) 3.9 - 4.4 1) 3.6 - 4.1 3.1 4.1 3.8 3.2 1) 3.4 1) 3.4 1) 3.4 1) 3.4 1) 3.4 1) 3.5 11.3.5			٥				neretord	,
3.2 (1) 3.4 (1) 3.4 (1) 3.4 (1) 3.4 (1) 3.5 (1	Surrentet %	3.7 - 4.0	ሊ 4.	4 .3	9.e	1) 3.9 - 4.4	1) 3.6 - 4.1	1) 3.4
12.9 - 5.0 5.0 5.2 4.8 12.9 1) 12.5 13.5	Protein.	3.1	4	8.6	3.2	4.6	1) 3.4	ff) 6.3 f)
12.9 +) 12.3 - 12.9 +) 12.5	Lactose	0.8 - 6.4	5.0	5.2	4 80.	a; }	14) 3.5	- (#
	Total Solids	12.9	:	1		1) 12.3 - 12.9	4) 12.5	

Based on 305-day lactation (Schwidt, 1971; Cerbulis and Farrel, 1975).

* i) Based on 2 to 6 weeks post-partum, (Hall, 1971; Schwulst et al., 1966; Cundiff et al., 1974; Gleddie and Berg, 1968). ii) Based on weaning averages at 200-210 days (above referentin).

selection for milk yield. Various reports however (Schmidt, 1971; Christensen et al., 1973 and Gaunt, 1973) note that butterfat, protein and solids-not-fat gercentages decrease with an increase in milk yield due to a negative correlation between these components and milk yield.

Some workers (Gleddie, 1968 and Hall, 1971) report that beef cows tend to produce more total solids, but much of the data is incomplete.

Tables I.8, I.9 and I.10 present the least squares means and standard errors of average constituent percentages of BF, PROT and LACT for 1976 and 1977. The differences between years for each constituent percentage for all cows were significant (P<0.05).

Results are within the range reported in the literature for protein, but tend to be higher than the value for BF% and total solids reported by Neidhart et al., (1979).

Butterfat

The BF% tended to be greater in 1977 than in 1976; differences were not as large for PROT% and LACT% between years.

BF% exhibited a greater range than PROT or LACT%, from 4.05 ± 0.42 for 4-year old DB cows in 1976 to $5.32\pm0.60\%$ for the single mature cow in 1976. However, as the small sample size of mature DB animals resulted in large standard errors, the BF% range may be more accurately illustrated by the SY group, which exhibited a range of from $4.27\pm0.19\%$ in 1976 to $5.3\pm0.30\%$ in 1977.

Least squares means and standard errors of average butterfat percentages. Kinsella 1976-1977.

•77

					Butterfat %	×	
·	. Breed	No. of cous 1976 1977	1977	1976 Mea n	S.E.	1877 Me an	S.E.
Grand Total	411	206	242	4.621	0.07	4.978	0.01
	Ĭ	4	**	7	9		
	λS	102	123	907.7	9.9	5.07e	٠ و
	2	56	EC	10.1	3 5	0.0	8
		33	3 8	4.87	2	0 0 0 m	9 5
Age in years				-			:
7	· V	73	93	4.47x	8		;
•	;				3	4 . / axy	<u>-</u>
	¥	Ξ	õ	4.71	91.0	7 96	
0	٠ د د	90	99	4.4	0.0	2 7	2 =
	3 2	-	2	4.42	0.15		0
	ŝ	P	=	4.31	0.17	4.30	0.27
6	A11	96	93	4.66×	0.11	5.01x	5
	¥	1	1	4.62	0.22	5	
	<u>ئ</u>	D	8	4.67	15		2 6
•	8	•	•	4.74	0.21	3 5	
	S	•	đ,	4.63	0.25	60 60 10 10 10 10 10 10 10 10 10 10 10 10 10	9.0
-	A11	20	38	4.31x	9. 16	5.26x	0.17
	¥	10	1	4.53	0.27		;
	S	9	<u>:</u>	4.27	61.0		9
	8	(4)	6	8.	0.42	. 20 . 20 . 20 . 20 . 30	
	ŝ.	Ю.	~	4.40	0.34	5.20	0
le ture .	TIV .	75	82	5.04y	0. f š	4.86y	0.1
	星	22	33	9. 12	6, 13	•	3
	۰ کو	\	Į	4.75	0	4	
	•	- ;	•	5.32	°.	5.01	2
	ŝ	=	a	70 7	•		

BFX by breed of dam within term are significantly different (P<0.05)

į

×. y=

Lesst squares seans of overall average BFX forall cous with different alphabetic letters are significantly different (Pc0.01)
Lesst squares means of overall average BFX by age of dem within each year with different alphabetic letters are significantly different (Pc0.05)

Least squares means and standard errors of average protein percentages, Kinsells 1976-1977. Table 1.9.

					Protein X	×	
	Breed	No. of COMB 1976 1977	1977	1976 Mean		1977 Mean	S. F.
Grand Total	A11	206	242	3.581	9.0	3.34m	0.02
	¥	4	3	2 55.0	2		
	. \$٧	102	123	3.428	000		5 6
	8	36	66	3.63	-	3.42	3 8
	So	33	78	3.52	0.07	3.45	8
Age in years		•					
en ;	1	7.	2	3.64xy	9.0	3. 50x	8
<u> </u>	3	:	,	;	,	•	
	2 3	= #	<u> </u>	. e	2 {	 	8
-	8	ū	8 8	- 10 - 10 - 17	58	d. 37	8
	80	Ē		3.93	80.0	9.00	5 6
6	, I.A	96	. 29	3.54×	90.0	3.48xy	8
	뿣	1	7	3.52	. 61		8
	SY	D	8	9.80	8		9 2
	2	.	φ	3.56	0.72	6	2 5
	SO	•	•	9.8	0. 1	3.83	0.4
•	A13	20	38	3.42×	8.0	3.39×y	90.0
•	Ŧ	8 7	•				
	SY	5	· Ξ	77.	0	7 6	5 5
	8	~	m	3.26	0 23		8
	SO,	6	~	9.48	•	3.26	
Mature	LIA	7.5	. 28	3.63y	80.0	3.38y	9.0
	¥	22	25	40.6	0.07	. 20.0	2
•.	δY	Ę	;	3.56	8	2.3	2
	2	-	•	3.80	0.33	9.43	0
	ž		•				

Least squares means of overall average PROTX by breed of dae within each year with different alphabetic letters are significantly different (P<0.05) Least squares means of overall average PROTX for all cows with different alphabetic letters are significantly different (P<0.01) Least squares means of overall average PROTX by age of dae within each year with different alphabetic letters are significantly different (P<0.05)

×. y.

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	Breed	No. of cows 1976 1977	1977	1976 Mean	. S.E.	1977 Mean	77 S.E.
Grand .Total	A11	208	242	5.481	9.0	B. 25a	0.0
	¥	5	9	. 50	8	48.	8
	<u>ن</u> و	2	123	5.56	8	S. 330	8 8
	8 2	3 C	P e	9. 330	17.0	a. 18c	8
	3	3*	٩,	8.4280	0.01	9.040	0.01
Age in years							
.,	14	. 75	93	5.50x	80.0	5.28x	80.0
7	¥	:	•	:	!		
	.		: ;		9.0		0.0
	8	<u></u>	3 8	, E	5 6	9.0	8
	S	2	•	\$ \$	3 5	7.	
)	}	2	\$	0
n	. I A	8	25	5.45x	8.0	5.25x	8
	里	7	1	3.0	0. 13	77 5	6
	<u>بر</u>	5	8	6.63	8.0	8	2 2
	8 2	•	•	9	0.12	50.00	2
	ŝ	•	-	5.24	0 5	6.31	0. 5
	A11	20	*	5.52×	0.0	5.20x	8.0
	¥	10	7	10 40		•	8
	Š	9	7	20.00	-	. E	3 3
	2	~	c	9.6	0.24		
	8		7	5.43	0.20	7	200
Keture	11	7.5	2	5.47x	0. 10	5.24x	8
	¥	33	22	10. 10.	0.01	100 100 100 100 100 100 100 100 100 100	5
	<u>,</u>	-	2	B. 52	80.0		
	8	-	₹.	7.0	0.38	3	2 5
	50	_	•	**			

Least aquares means of oversil average LACT % by breed of dam within each year with different alphabetic letters are significantly different (P<0.05) Least, squares means of oversil average LACT % for all covs with different alphabetic letters are significantly different (P<0.01) Least squares means of oversil average LACT % by age of dam within each year with different alphabetic letters are significantly different (P<0.05)

Protein

PROT% ranged from an average of 3.26% for DB and DS 4-year olds in both years to 3.90% for the single DB cow in 1976.

Lactose

LACT% ranged from 4.94% for the 3-year old DS in 1977 to an average of 5.67% for the HE and SY 2-year olds in 1976.

Effect of Breed of Dam on Constituent Percentages

Analyses of covariance of average constituent percentages (Table I.11) show that the effect of breed accounted for a highly significant source of variation (P<0.001) only for the 1977 LACT%. Differences between all breeds for BF% and PROT% were not significant (P>0.05) as indicated in Tables I.8-10, although generally the HE tended to produce the highest BF% each year.

Breed differences for PROT% were negligible. For LACT% however, HE and SY produced significantly (P<0.05) more than DB and DS groups.

As noted, BF% values for the dairy crossbreds in this study were higher than those reported in the literature. Because dairy selection programs have usually emphasized milk volume rather than constituent content (Cerbulis and Farrell, 1975) BF% may either be a reflection of breed effect or an effect of the suckling or milking regime. Christensen et al., (1973) found that BF content was less affected by environmental factors like calving interval and

Table I.11. Analyses of covariance of average constituent percentages, Kinsella 1976-1977.

		•			F TESTS	STS		
Source	2		ā	BF %	PRO	PROT %	LAC	LACT X
	1976	1977	1976	1977	1976	1977	1976	1977
Age of dam	ო	6	4.61.**	2.41	1.92	1.91		
Breed of dam	(9	m	- 8.	-8	0.66		<u>.</u>	5 6
Age x Breed of dam	co	. 60	0.48	1.41	0.76		4.7 4.38	1.48
Sex of calf	-	-	0.27	0.88	0.36	Ç.	ć	;
Sex x Breed	ю	m .	0.89	1.27	0.0		5 6	96
Age of calf		-	0.40	3.85*	0.12		* .	80
Error	185	221			!		9	2.38
Total	202	241						

* P<0.05; ** P<0.01; *** P<0.005

herd average than either milk or BF yields. As these dairy types in the present study were all crossbreds, the BF content may also be affected by crossbreeding or may be a factor confounded with the method of milk removal.

Most published reports (Schmidt, 1971; Gaunt, 1973 and Preston and Willis, 1974) concluded that milk and BF yields are traits with low heritabilities of between 0.2 and 0.3, although Tong et al., (1977) calculated heritabilities for milk and BF yields of up to 0.5. While crossing results in heterosis for traits of low heritability like milk production, the heritability of constituent percentages is generally higher (Preston and Willis, 1974). There is probably a combination of factors involved that influence the BF content of milk and further investigation is required.

Effect of Age of Dam

The effect of age of dam accounted for a significant (P<0.01) source of variation only for 1976 BF% (Table I.11). As shown in Tables I.8-10, mature dams yielded significantly (P<0.01) more BF% than other age groups. Generally, there were no marked age trends for the other constituent percentages of PROT and LACT.

Whereas Christensen et al., (1973) found that BF% was not considerably influenced by either age of dam or calving interval, Schmidt (1971) noted 0.2 and 0.4 unit decreases in butterfat% from the first to fifth lactation, but a negligible change in protein with increasing age.

June and September Constituent Percentages and Seasonal Variation

1

As illustrated in Figure I.5, for dairy cows the negative relationship between milk yield and component percentages results in curves for component percentages that are inversely proportional to those of milk (Schmidt, 1971; Christensen et al., 1973; Preston and Willis, 1974, and Cerbulis and Farrell, 1975). According to Schmidt (1971) this may be due to the fact that butterfat, protein and solids-not-fat percentages are high in colostrum in the first part of the lactation. He suggested that the protein and solids-not-fat percentages rise around the sixth month of lactation due to the possible hormonal effects of pregnancy because if the cow is pregnant, these levels tend to remain constant.

Lactose percent, according to Schmidt (1971) is low in colostrum, but usually increases to a high level at the start of lactation and exhibits a small decline near the end.

Trends for seasonal variation of constituents are shown for Herefords, Angus, Galloway and crossbreds of the same breed composition (Figure I.5) (Gleddie and Berg, 1968; Hall, 1971). Gleddie and Berg reported an increase in BF% from the beginning of lactation, first measured at 35 days, followed by a sharp increase at the end of 155 days. Seasonal fluctuations were small for protein. The total solids and SNF% decreased to the second month and then

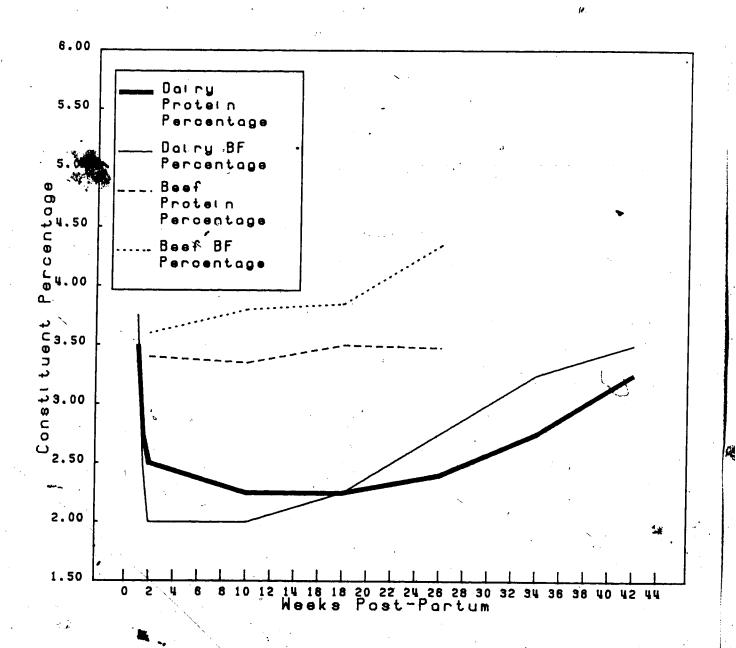


Figure I.5 Milk constituent trends for dairy and beef cows

increased gradually, more nearly similar to the dairy breeds.

It may be possible that differences noted in seasonal fluctuations for constituents may be attributed to the variation in times of measurement during the lactation. For example, the beef cows in Gleddie's experiment may have been first measured after the initial peak in milk yield. On the other hand, Gleddie (1968) reported a small but positive correlation coefficient between milk yield and BF%, contrary to results noted in most published reports, which might account for the differences in patterns. In another study conducted in Rhodesia, Richardson et al., (1977) found that BF% increased signficantly at 5 weeks post-partum during the peak yields. These workers suggested that the BF% increase was related to a corresponding increase in cow weight.

Based on the foregoing, there seem to be no consistent trends for seasonal variation in constituent percentages for beef cattle.

Results taken from the present study are illustrated in Figure I.6 averaged across breeds.

Least squares means and standard errors of June and September constituent percentages are presented in Tables I.12 through I.17. Differences between the two periods for all constituents were significantly (P<0.01) higher for September when averaged across all breeds. As illustrated in Figure I.6 BF% consistently had the highest increases, averaging 0.41 between periods over both years. Protein %

Table 1.12. Least squares means and standard errors of average delly June and September

	x -			Aver	Average Daily 8f	×	
	Breed	No. of cows	Mean	.E.	September Mean S	S.E.	Change
Grand Total	A11	206	4.4	90.0	4.60	0.07	0.36
	¥		4.71a	0.1	4.780	5	6
	<u>ئ</u> ج	102	4.32b	800	4.758	8	
	s o	9 E	4.470	0.36	4.4	0.22	0.34
Age in years	l	}		e 5		2	0.55
~	I I V	7.8	4.22×	0.10	4.71x	8.0	0.4
	¥	Ξ	4.61	0.24		\$	
	Š	96	4.23	0.13	4	3:	2 5
٠	2	•	4.31	0.0	4.52	2	7 6
<u>.</u>	So	E.	3.74	0.23	4.87	. 0	0.5
6	Ĩ.	98	4.50x	2.0	4.83xy	0. 12	0.33
	꽃;		7.	0.30	4.70	0.25	9
	5 8	I	# ·	0.21	8	-	92.0
	2	.	7	0.28	4.70	0.24	0
	S	9		7.	2 .	0.30	0.52
• .	Ī	20	4. 19x	0.21	4.49x	0.1	8 .0
	불.	10	# *	9.0	90,7	5	;
,	<u>ک</u>	9	4.1	0.25	4.4	2.0	3.0
	8 2	CV (3.62	96.0	4.47	0.47	0.83
Ċ	S	m	8	9. 78	4.50	0.3	0.20
Mature	A11	75	4.80y	0.22	5. 18y	0.1	0.28
	뿦	22	5.22	0.17	60.03	9.14	2 • •
	.	₹ .	7.62	0.13	7	5.0	0.78
	8 8	-;	9.7		. 56 50	9	0.42
	3	-					

a.b . Least dquares means of constituent percentages by breed of dem within months with different letters are significantly different (P<0.05)

X.y * Least squares means of constituent percentages by age of dam within months with different latters are significantly different (P<0.05)</p>
* significant at P<0.05; ** significant at P<0.01</p>

٠				\	:		
				Average	, A	Protein X	
	D .	No. of covs 1976	Mean .	. E.	September Mean S	S.E.	Change
Grand Total	A11	206	3.47	80.0	J. 64	0.03	0.17•
	뿦	45	200	6		8	
	.xs	20,	3.38	8	2 62	88	
•	200	38	3.60	0.14	3.69	; <u>0</u>	8
Age in years	SO	33	3.80	8	3.548	80.0	9
· .	`:	,					
N		78	3. 62 ×	8.0	3.65x	9.0	0.03
	¥	Ξ	3.50	0.13	3.76	0 0	36
	λ\$:	9 ¢	9.43	0.01	3.8	8	9
`	2	5	3.83	0.1	3.75	000	-0.20
\	SO	Ç	3.30	0. 12	2.51	80.0	8.0
. 6	T I V	36	3.48×	90.0	3.58×y	80.0	0.0
\ \	뷮	•	3.43		:	· (
\	SY	. 2	3.33	==	3.67	- C	0.5
,	2		96 · E	0.13	30.00	-	0.00
	SO	•	3.62	E	3,46	0.14	-0.16
4	A11	20	3.31x	0.11	3.54x	0.0	0.23
	呈	40	3.20	5	3.75	0	0.47
	<u>ک</u> (5	3.28	0.13	3.6	0.0	0.33
	8 2	~ ~	3.23	8	7.5	0.22	0.0
	3	•	,	0.24	 	 E	0.14
Mature	- V	75	3.46x	0. 12	3.78y	8.0	0.33**
	¥	23	80	8	7		
	SY		3.4	0.0	3.62		25.0
•	2	- :	08.	0.43	4.21	0.32	800
	ŝ	=	8 .5	0 .T	3.62		0.24

a - Least aguares meaner of constituent percentages by breed of dam within months with different (P<0.05)

- Least squares means of constituent percentages by age of dam within months with different letters are significantly different (Pc0.05)

* significant at P<0.05; ** significant at P<0.01

Table I. 14. Least Squares means and standard errors of average daily dune and September lactors percentages. Kinsils 1978

•				Aver	Average Daily Lactose X	actose X	-
	Deer 8	No. of cows 1976	Mean	. S.E.	Sept	September n S.E.	Change
Grand Total	AII	206	10 -	8.0	5.86	80 0	0.75
•	¥	57	5.28	0.07		8	
	ζ	102	5.26	8	\$	3 8	
•	2	56	4.90p	0. 12	5.79a	9 6	
Age in years	Š	33	4. 9 4b	8.0	5. 85s	0.0	
	, LEV	75	5. 12x	8	, 60	8.0	0.76
	里	:	8.36		6.0	41.0	
	ž į	36	9.34	8	88.	0	
	8 2		4.63	o. 0	90	0.12	2 2
•	S	<u>.</u>	5.13	÷.	10.9	0. 13	9
	A 11	38	5.04x	0.01	10°	8.0	0.82**
	¥	-	5.29	9.1	9.0	0.17	•
	5 2	ē.	2.28	0.0	6.87	0.12	0
	8 8	•			9 .	0.0	0.76
	ŝ	• •	4 67	• •	5.80	0. t	1.13
₹ .	114	50	5. 16×	0.10	. 89 × 80	0.12	0.73**
	¥	1 0	8.20	0.17		0.01	•
	<u>ک</u> ا	5	5.26	0. 12	7.0		3 4
	8 2	ev e	88	0.27	6.15	0.32	10
`		•	3	0. 22	3 78	0.26	9.0
Meture	A11	. 75	5, 12x	D. 10	. 5.82×	0.12	0.70
	¥	22	5.27	80.0	6.81	0, 10	2
	<u>ک</u> د	-	9 . 1	80.0	9.80	6	0.75
	8 2	- ;	8.	0.30	7	0.4	0
	S	Ξ	8	÷	-	:	

a.b = Lesst squares means of constituent percentages by breed of dem within months with different letters are significantly different (P<0.05)</p>

x ... Least aquares means of constituent percentages by age of dem within months. with different letters are significantly different (P<0.05)

significant at P<0.05; ** significant at P<0.01

•				Avera	Average Daily BF	*	
	B	No. of Cows	J. Least	June	September Meen S	Page 7	Change
Grand Total	, A11	242	4.74	0.0	5.20	8	0 46
•	業	8	5.04	0. 13	5.11	. 6	6
	ž č	123	4.976	8	9	8	0.62
Age in years	SO	3 2	4.68 4.68 60	0.5	8. 12s	0.24	0.45
2		. 83	4. 80x	0.11	. 8 8		0.50
	¥	.	į	61.0	,		
	≿	38	4.58	0	2	2	5 6
	2 2	2	6.6 1	0.36	5.29	6	
•	S	-	3.83	0.34	4.84	0.34	0.0
	I V	32	4.80xy	0.14	5.22×	9. 7	0.42
	¥	7	4.18	5		;	;
	λS	8	78.4		2 #		0.58
	8	•	8	98.0	. 4	9 6	
	SO	•	9.44	0.48	G. 83		
	. T	%	8. 10y	0.22	5.42x	0.22	0.32
	·¥	4,	76 #	;		. (;
,	λS	=			3 2	0.42	8 6
	8		5.07	0.32			
	SQ	~	4.82	0.93	4	9 6	5
Meture	A11		4.57×	0.1	5. 16x	0.14	0.58**
	¥	23.	5.22	8	:		. 6
	λS	7	2	3		3 6	8
•	8	*	7	3 2	2 6	8 6	9

a.b - Least squares means of consitutent percentages by breed of dam within months with letters are significantly different (P<0.05).

x,y . Least aguarss means of constituent precentages by age of dam within months with letters are significantly different (P<0.05).

* Significant at P<0.05; ** Significant at P<0.01.

Table 1.16, Least squares means and standard errors of average delly June and September protein percentages, Kinesia 1977.

•			į	Averag	Average Daily Protein X	otein X	
	81000	No. of cous	Mean	S. F.	September Mean S) 1 2 3 3 4	Change
Grand Total	A11	242	3.31	0.03	3.57	. 60.0	0.26**
	¥	6	3.31	8	3.65a	8	
	<u>ئ</u> د	123	3.28s	0.03	3.56		
	8 8	D 00	3.30	0.0	3.50 3.50 8.50 8.50	0.0 0.07	0.16
supply the same		•		•			
7	I V	8	3.38xy	9.0	3.61x	80.0	0.22**
	¥	9	2.48	90.0		5	į
	λS	38	3.27	0	2 47	3 8	8 6
	z	. 02	3.31	=	3.71	3 2	9 6
	SO		3.50	ە. ق	3.53	.0	80.0
		52	3.31x	9.0	3.66xy	9.0	0.35**
	¥	,	3 33	8	•	;	;
	λS	8	3.43	2	2	3 5	
	2	•	3.37	-	333	2	5 6
	SQ	a	3.13	1	3.83		200
•	A11	26	. 3.24xz	0.07	3.54×	90.0	0.30
	뿦	7	3.32	6		8	
•	λS	<u>:</u>	3.4	0.12		3 =	
	2	m	3.28	0	9 6	2 5	÷ (
	SO	~	3.22	0. 16	3.31	0.21	8
Mature	T V	. 82	3.28×y	9.0	3.49×z	8	0.20**
	뿦	28,	3. 12	90		:	ç
	λS	7	3.30	0.12	4	: <u>:</u>	7 :
	8	. •	3.38	0	3 47	2 5	
	50	a	97. 7			? :	3

1 = Least aquares means of consituent percentages by bread of dam within months with different letters are significantly different (P<0.05).</p>

 κ,y,z^{μ} . Least squares feans of constituent percentages by age of dam within months with different letters are significantly different (P<0.05).

* Significant at P-6.05; ** Significant at P-0.01.

Table 1.17. Least equares means and standard errors of average daily June and September lactoss percentages, Kinsella 1977.

	,				Aver	Average Deily Lactose X	Lactose 9	
	•	į	No. of COUR	Cune		September		
	Breed	£	1977	Hean	8. F.	E .		Change
Grand Total	A11		242	5.16	9.0	5.33	9.0	0.17
	¥		48	8 48.	5	-	1	
	λ	_	23	5.278	5 6		8 6	8
	8		33	4.830	5	37.	3	0.12
Age in years	S			8.016	0.0	5.07b	8	8
8	413		. 8	5.26x	0.01	5.32x	0.0	6
	¥		<u>a</u>	5.52	ç	97		; ;
	Š		35	8.30	0.0		3 8	5.0
*	8		20	5. 13			: 6	3:
	ń O		•	7.18	0.18	8		9
e	Y I		52	5. 18x	0.07	. 9. x	0.01	0.13
	¥		1	87 · S	9	. \$;
	۵,		8	4.07		, E	8 5	8 9
	2		•	8.				2 6
	S		a	9,48	0.24	g. 13	0.2	9 9 9 9
•	A11		26	5.00x	0.12	5.32×	0. 0	0.23
•	¥		7	200	:	;	;	
	S.		:	9.4	200		2:	7 7 0 0
	2		m	4.82	0. 17	93		3 2
•	SO		~	4.75	0.28	5.12	0.25	0.37
Hature .	F F	_	92	5. 13×	0.01	. 30 X	0.0	0.22
	里	•	52	9.90	0	96	2	8
	5	•	•	B. 26	0.21	5.45		3 2
	8 2		<i>y</i>	4.84	5	5.43	9	
	ć		0		:			

a.b $^{\circ}$ Least squares means of constituent percentages by bread of dem within months with different latters are significantly different (Pc0.03).

* Significant at P<0.05; ** Significant at P<0.01.

x = Least squeres means of constituent percentages by age of dem within months with different letters are significantly different (P<0.05).</p>

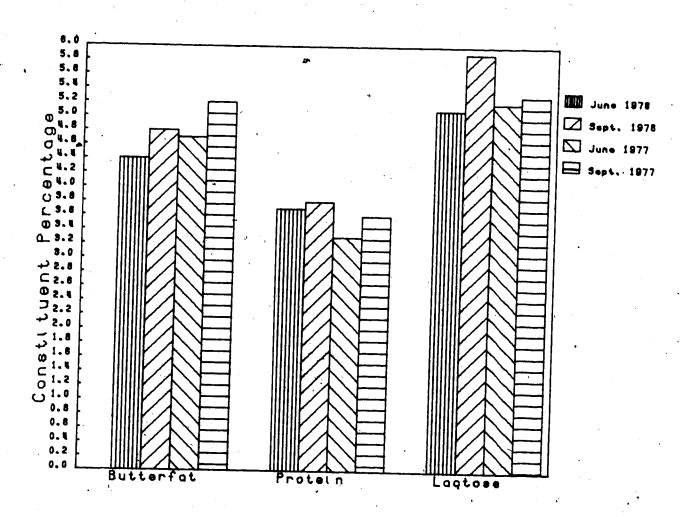


Figure I.6 June and September averages for butterfat, protein and lactose percent, Kinsella 1976-1977

changed less noticeably, increasing by 0.22 units. Lactose % increased by 0.75 in 1976 but by only 0.17 in 1977. The 1977 results are comparable to trends reported by Schmidt (1971).

For breed-age categories (Tables I.12-17), there was no consistent trend for variation in components. Protein and lactose % increases were significant for most breeds. Although some age differences for PROT% were also demonstrated, no specific trends for seasonal variation could be determined to the dam.

September constitution tages. The effects of breed of dam accounted for significant variation in June BF% (P<0.05) in 1977, in June LACT% (P<0.01) for both years, and for September LACT% in 1977 (P<0.01). The HE breed produced significantly (P<0.01) higher BF% than all other breeds in June each year. The DB and DS groups produced significantly less LACT% (P<0.05) than HE and SY in June. However, no other significant differences were observed for seasonal variation between breeds.

Table I.18 shows that the age of dam accounted for a significant (P<0.05) source of variance for June and September constituents in 1976. Results in Tables I.12 through I.17 indicate that only mature dams produced more BF% than other age groups. Although some age differences for PROT% are also demonstrated, no specific trends for seasonal variation could be determined by age of dam.

Table I.18. Analyses of covariance of June and September butterfat, Kinsella 1976-1977

**.			į	-					F TEST			,		
<i>t</i>	D.F	u.	· June BF	. BF	Sept. BF	8F	June PROT	PROT	Sept. PROT	ROT	dune LACT	LACT	, toes	Sept ACT
Source	1976	1976 1977	1976	1977	1976	1977	1976	1976 1977	1976	1977	1976	1977	1976	1977
ige of dam	e	m	3.37* 2.19	2. 19	2.75* 1.06	7.06	2.03	2.03 1.71	1.82	1.77	0.38	0.79	0.07	0.07 0.07
Breed of dam	ო	m	2.18	3.23*	0.25	0.51	1. 10	1.10 0.17	1.53	1.03	5.39**		0.23	4
Age x Breed	Ø	O	69.0	1.72	0.51	0.64	0.83	1.56	1. 12	2.04	1.60		8	4
Sex of calf	-	-	0.14	0.26	0.22	96.0	0.02	90.0	20.	0.92	0.27	0.79	9 6	
Age of calf	- -	-	1.48	0.02	6.46*	9.21**	3.15	0.74	9.54**	0.30	5.92*	1.69	90	, -
irror	188	224)) ;	,
rotal	205	241				•					•			

* Significant at P<0.05; ** Significant at P<0.01.

Relationships between June, September and Average Milk Yields and Constituent Percentages

Table I.19 presents the phenotypic correlations of all milk yields and constituent percentages for 1976 and 1977.

Table I.20 presents correlations examined in this study.

a) Correlations of June with September Milk Measurements:

June and September milk yields were moderately correlated (Table I.20 a); r = 0.55 and 0.58 for 1976 and 1977 indicating that other factors during the lactation would influence the daily yields. Each of June and September constituents had low to moderate values ranging from 0.22 for June with September LACT% in 1977 to 0.47 for June with September PROT% in 1976.

b) Correlations of Milk Yields with Constituent Percentages:

Generally, all measurements of milk yield were negatively correlated (P<0.05) with all constituent percentages (Table I.20 b) showing little variation in trends for overall, June and September measurements.

Average milk yield with average BF, PROT and LACTX had r values of -0.13, -0.29 and -0.01 respectively for 1976 and similar correlations for 1977. These values correspond to those in the literature reviewed for dairy (Schmidt, 1971; Preston and Willis, 1974, and Cerbulis and Farrell, 1975; but differ with some of the

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.77		.24	; !	. 23	8	.36	80.	69.	027
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8.	Ø	=	.03	2.		. 47	0.50	99	6
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. 02ns	2ns	.07ms	-04ns	44	41	. 34.	1	.83	98
.03ns	3ne	. 05ns	01ns	34	47	5	63	i	7
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is = not significant, at P>0.0

ible 1.20. Phenotypic correlations of milk yields and milk constituent percentages.

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1977	80	7.7	25	1977			. 73	29	- 19	. Q	1977	B 1	2		6		-	ě	7	76	8
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•	*?			â						•	າ.* ີ່ ວ									Ŋ	•

A m Average, U m Gune, S m September, M m Milk yield. B m Butterfat M, P m Protein M. L m Lactose M. coefficient values for beef breeds reported by Gleddie and Berg (1968). These authors found that BF% was positively correlated with milk yield and negatively correlated with PROT and solids-not-fat%. Jeffery (1971) cited literature with similar values between BF% and milk yield. In the present study, there was a tendency for negative correlations between the overall averages of milk yield and protein percentages with r values of -0.29 and -0.22 for 1976 and 1977 respectively (P<0.01).

percentages were variable, ranging from -0.66 between June PROTE and LACTE in 1976 to a low and not significant correlation between average BF% and LACTE (Table 1.20 c). Generally, "BF% and PROTE had moderate to low positive r values. The correlations between BF% and LACTE were low and not significant, and PROTE and LACTE were moderately and negatively correlated. These trends were similar for both months.

Jeffery (1971) reports all r values as positive among % milk components, similar to Gleddie and Berg (1968) and other workers cited in his study.

D. CONCLUSIONS

A milking experiment was conducted to examine differences in milk selds and constituent percentages between Hereford, Beef-Synthetic, Dairy-Beef and

Dairy-Synthetic breed groups of range cows. There were significant breed differences between Hereford and Beef-Synthetics, the latter yielding more milk than Herefords. However, the Dairy-Beef and Dairy-Synthetic cows yielded significantly more milk than the beef breeds and exhibited greater persistency each year, suggesting that there are differences in both average daily milk yields and seasonal variation between dairy cosses and beef cows on range.

Age trends for milk production were less noticeable but generally, 4-year old and mature cows yielded more milk.

Constituent percentages for butterfat, protein and lactose are within the range reported in the literature for range beef cattle although the 4.8% average for butterfat averaged across all breeds for both years was slightly higher than values reported in other published material. Differences between breeds for butterfat% and protein% were not significant although Hereford dams tended to produce the highest butterfat content each year. Breed differences for protein% were negligible. The Hereford and Bell-Synthetic cows produced a significantly higher lactose% content than the dairy crosses. All constituent percentages were significantly higher in September than in June. No consistent trend for variation in composition was noted between breeds.

It is not clear whether differences in milk yield are the result of breed differences and heterosis resulting from

crossbreeding, or are more a reflection of the milking regime. Results indicate however that the introduction of a dairy breed into a beef line will result in higher milk production under range suckling conditions compared to beef breeds and crosses.

The results indicate a high butterfat content for the dairy crosses in this study which raises some interesting points for further examination. Is butterfat, a reflection of breed effects per se or rather, more influenced by the suckling or milking regime? Or, as dairy cows in this experiment were all crossbred, is the butterfat, also affected by crossbreeding? It is commonly held that crossbreeding results in heterosis for traits of low heritability like milk production; however, according to the literature reviewed, the heritability for all constituent percentages is much higher at approximately 0.5. More investigations are required.

Based on the high butterfat% and average daily milk yield of dairy crosses used in this study, it seems likely that calves of DS and DB dams ingested more total energy than HE and SY calves, thereby affecting preweaning growth response. For butterfat% alone, HE and SY calves ingested approximately 325 grams per day compared to 395 grams per day ingested by calves of dairy cross dams. This in turn may influence growth response over and above the inherent growth potential of the crossbreds, and consequently result in higher weaning weights.

II. FACTORS INFLUENCING LACTATION PERFORMANCE OF RANGE BEEF AND DAIRY-BEEF COWS

A. INTRODUCTION

Many studies have indicated that increasing milk production in a cow-calf operation can be effected by crossing dairy with beef breeds (Long, 1980).

In the preceding chapter significant differences were found in milk yield among breed and within age of dam categories. Whether the effects of preed and age of dam are confounded to a great degree by other factors such as the nature of the suckling regime is not clear. However, the results suggested that breed and age of dam effects exert some influence on the level of milk production and associated constituent yields. The nature of the effects of breed and age of dam on milk yield has not been reported extensively. For example, it is not known if breed of dam of effects on milk yield are more pronounced at the beginning or end of lactation. This may have practical implications for operations that consider early weaning as part of the program.

In addition, of interest to the producer are factors other than the breed and age of dam such as calf age, sex and birthweight, and cow weight changes before and after calving which may influence milk yields and constituent percentage yields.

The purpose of this part of the study was to determine the influence of the above cow and calf factors on June, September and average milk yields using multiple regression analyses. A limited analysis of the effect of these factors on constituent yields was also conducted.

B. MATERIALS AND METHODS

Data were collected in 1976 and 1977 from a milking experiment involving June and September measurements of lactation. The 206 (1976) and 242 (1977) cows ranging from 2 to 10 years of age represented four beef and dairy-beef groups: Hereford (HE), Beef*Synthetic (SY), Dairy-Beef (DB) and Dairy-Synthetic (DS). The general management and breeding program of the experimental herd were described in Chapter I. A detailed account of the milking experiment was also provided earlier.

Statistical Analyses

Dependent variables recorded for the purpose of this study were June and September milk yields and their overall averages and constituent yields of butterfat (BF), protein (PROT) and lactose (LACT) for each year.

Independent variables analyzed as categories were breed of dam (B), age of dam (A) and sex of calf (S). Quantitative independent variables, ie: covariates, that were included either alone or together were age of calf (CA), calf birthweight (BW), cow winter weight loss from October of the preceding year to calving (WWLS) and cow post-calving

average daily gain from calving to September (cow ADGCS).

Calf average daily gain from birth to September (calf ADGBS)

was also included as an independent variable.

The analyses of covariance for unequal subclass numbers (Mehlenbacher, 1978) and levels for the main effects of B, A, S and CA are as described in Chapter I.

o determine the influence of cow and calf variables with milk variables, multiple stepwise regression analyses were used.

As trends were similar for each year, regressions for 1976 were not computed.

Stepwise Multiple Regressions

Stepwise multiple regressions of milk and constituent variables on combined cow and calf variables were computed using the Special REGRESSION as outlined by Nie et al., (1975). Some of the stepwise regression models involved the sequential regression of variables which entered the equations in a hierarchical order based on partial correlations. Some of the variables which entered were quantitative or continuous; others such as A, B, A x B and S were discreet or non-continuous. In the other stepwise regressions, some of the variables were forced to enter the equations.

preliminary analyses of the data examined the combined effects of cow and calf variables on average milk and constituent yields for both years using three models of regression. In the first model, the effects of

calf age and sex were not included (Table II.1, Equations 1a and 1b). This stepwise multiple regression analysis provided some indication of the variance of overall average milk yield explained primarily by genetic effects of the breed and age of dam. The association of the effects of management factors of cow winter weight loss during gestation and post-calving weight gain on lactation performance were also examined.

In the second stepwise regression model, age and breed of dam and age and sex of calf were ignored (Table II.1, Equations 2a and 2b). This method permitted the introduction of only calf birthweight and cow weight changes, factors that may be affected by management.

All variables in the last two models were introduced on the basis of the highest partial correlations with the dependent milk variable. No variables were forced to enter the equation. If two independent variables were correlated, some of the effect of the first is removed by the subsequent variable. Consequently, the additional variance explained by a variable over and above the preceding variables is conditional to the preceding variables entered into the regression equation.

A third model of stepwise regression analysis involved the sequential regression of average milk, BF, PROT and LACT yields on all cow and calf factors in 1977 (Tables II.2 and II.3). However, independent variables

of interest were forced to enter the equation based on a priori reasoning. This model allowed examination of the relative weight of association of one variable over another with the dependent milk variable. It also provided information as to the feasibility of using various milk component traits as indicators of milk performance.

Other regression models were computed as outlined by Overall and Klett (1972). These regressions permitted the examination of the influence of each of the cow and calf factors on the dependent milk variables of everage milk, June and September yields. All other variables were ignored (Tables II.5 through II.10).

These regressions were computed by: a) ignoring the effects of A, B, A x B, S and CA and then forcing in the variables of interest; b) forcing the effects of A, B, and A x B to enter the equation before forcing in the variable of interest, and c) forcing all the effects of A, B, A x B, S and CA to enter the equations before the variable of interest. Methods b) and c) allowed examination of the effects of individual variables over and above main effects.

C. RESULTS AND DISCUSSION

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Factors Influencing Average Daily Milk and Constituent Yields

Stepwise Multiple Regression Analyses

Analyses of the data are presented in Table II.1 for regressions using methods 1 and 2 for comparison of the 1976 and 1977 results.

Results indicated the age of dam x breed of dam interaction (AxB) alone accounted for an average of 34.0% of the variance (P<0.01) in average milk yields each year (Table II.1, Equations 1a and 1b). All other cow and calf variables including birthweight and cow weight changes during gestation and lactation accounted for a small and not significant proportion of the variance.

As the first three independent variables explained most of the variance in milk yield, the equations were limited to the third step of the regression analysis.

In Equations 2a and 2b (Table II.1) calf birthweight, cow winter weight loss and ADG from calving to September explained little of the variance. When compared with the high association with milk yield shown by B and A effects, is sults would suggest that more effects of BW, WWLS and cow ADGCS were largely accounted for by the age and breed of dam variables.

Equations 1a,b, c and d (Table II.2) present stepwise regressions of 1977 average daily milk, butterfat, protein and lactose yields on cow and calf variables forced to enter the equations.

quation	Independent Variables entered sequentially into	X Total variate explained (R' x 100)	% Additional variance explained	Partial D	SED*
a) 1976	map Jo X X BO T CASH	34.4	1	0.17	0.03
	Age of dam'	96.9 96.3	4.80 0.60	0.04	0.05
b) 1977	B x A of dam Calf 84rthwarght Age of dam	33.88 36.88 37.8	00°	0.06	0.03
a) 1976	Calf Birthweight Cow WWLS Cow ADG C=51	- 25 E	្ត	. 0 - 0 - 0 - 0 - 0 - 0 - 0 - 0 - 0 - 0	0.02
b) 1977	Calf Birthweight' Com ADG C-S' COW WWLS'	3.0 3.1 3.1	2.80 0.10	40.0 - 0.0 -	0.30

* Years
* Kg/day
* Standard error + p<0/08



00.000 00000 000000 000000 SED. Stepwigg regressions of average delly wilk, butterfat, protein and lactose yields (Kg/dey) on cow and/calf variables, Kinsella 1977. Partial b 6. 45** -0. 10 -0.95 0. 17 -0.36 0.00 0.00 0.00 0.00 0.00 0.40 00000 Additional variance explained £0.0 20.0 000 × X Total variance explained (Rf x 100) Order of forced independent variables Breed of dem (B) A × B Calf ADG B-S¹ Calf Sex Calf Birthweight Birthweight: Calf Birthweight dem (A) COW ADG C-S' AG 8-5" COM WWLS" CON WALS ŏ Calt Dependent · Ave. milk yield? Ave. PROT yield? Ave. LACT yield* Ave. Br yield? Table 11.2. Equation. ç ê ٠ .

Standard error of b • P<0:01

Age and breed of dam and the A x B of dam interaction accounted for a highly significant proportion of the variance (P<0.01) of milk variables. These effects together explained an average of approximately 42% of the variance for all milk variables except lactose yield, for which they explained 33.1% of the variance. In these equations, the amount of additional variance explained by the A x B of dam interaction was not significant, similar to results by Gaskins and Anderson (1980).

Calf ADG from birth to September and calf age both accounted for significant amounts of additional variance in all the milk variables. They each explained between 5 and 11% of the additional variance.

In Table II.3 equations 1a,b,c and d indicate that when calf ADGBS was forced to enter the equations first, it explained between 22.2 and 23.5% of the variance in each of the milk variable yields. Given that its correlation with milk yields and constituent yields is moderately high (Table II.4, r= 0.6) its resulting association is predictably significant (P<0.01).

The relationship between milk and weaning weight, a function of daily gain, is discussed in Chapter III.

Many authors have noted the high correlation between milk and gain in several species. Rae (1977) concluded that milk production can be increased within a flock of sheep by indirect selection through lamb growth. Rae

quet fon	Dependent variable	Order of forced independent variables	K. Total. variance explained (R* x 100)	X Additional variance explained	100	
(6	Ave. milk	Calf And R-C			1:	
t -	V.(e) V			· ,	7.34	0.61
)	3		7.17	-0.12	0.0
		X BO LESS		2.7	0.27	8 0.0
	· •		7	. . .	-0.23	0.03
	٠ ٠	COW ADG C-S	** *** **** **************************	o	-0.63	0.47
	* *	COM MALS	6.63	4 .0	0.47	0.0
ار 10 14	Ave. BF	Calf And B-c	440 60			
				· · · · ·	8	8
٠.		- 4		N .	90.0	0
		X DO LIND		3.2	7 .0	0.0
•	•	•	47.0	4.0	0.26	0.0
		-	9.7	9. 0	9.0	0.0
			47.6	•	-0.01	0.0
ΰ.	Ave. PROT	Calf Aba B-S	22.2**			8
	~	•		7 16	***	500
		Calf Sax				9.0
	•			¥. (5.0	0.
•			<i>x</i>		0.0	0.0
•				•	10.0- V	0
	•	COW WELS	62.0	•	0.0	0.0
0	AVE. LACT	Calf ADG B-S	23.5.	. 	2	8
	yield	•	- 46.4**	22.0	2	
	•		40.30			
		Calf ADG B-S	9.07			
			- 67			5 6
		.3) (
		COW ADG C-S	0.08	. 0	3 0) (
	•				3	3

explained not significant (P<0.05) by introduction of 4th vari

lable yields with cow and cal

Variables!	CA	2	¢	. 8 − 8 − 8 − 8 − 8 − 8 − 8 − 8 − 8 − 8	Calf ADG 6-S	COW ADG	- Ave. Milk	
Ave. Milk June Milk Sept. Milk		*464		. 63 . 48 . 62	80.08		1 1	**
Ave. Br June Br Sept. Br	000	464		86 88 4	. 5.48	. <u>\$</u> 89	9	• .*
Ave. PROT. June PROT. Sept. PROT.	56	. 33 22 75		20.40 € 60.00		5 88	- 8.	4 . •
AveLACT. June LACT. Sept. LACT.	868	4 5 4 0 4 0 4 0	*	6 6 6 6	823	. 8 .6.8		

| Variables defigned in Chapter E. = not commuted

reported correlations of 0.63 between ewe's rilk yield and lamb preweaning ADG.

Similarly, calf age in this study accounted for a significant proportion of the additional variance (Table II.3, Equations lamb,c and d). These results suggest that older, more vigorous calves increase their capacity for milk as intel energy requirements increase. However, results do not indicate at what point in the calf's prevening period the demand for milk decreases as forage and milk yield was low and not significant (Table II.4, r= 0.03). Interpretation of the effect of calf age however if difficult, as this factor may be interpreted to mean either days in lactation of the dam of significant due to growth of the calf.

Calf sex, birthweight and cow weight changes during gestation and lactation accounted for a small proportion of the additional variance over and above A, B and A x B in each of the milk variables (The II.2). In Table 11.3 however, calf sex demonstrated a more significant (P<0.05) association with milk yield variables when it was forced to enter the third step of the equations.

The foregoing analyses demonstrated that each of the independent variables accounted for approximately as much variance in one dependent milk variable as in another. The high correlations between average milk yields (Table II.4, r= 0.9) and constituent yields, and

Byreas

the similar association between the cow and calf factors and each of the milk variables suggest that one measurement of either a milk yield variable or constituent yield variable is likely a sufficient indicator of the association between milk traits and cow or calf variables. Because of the technical difficulties involved in the determination of constituent percentages, milk yield alone would serve as an accurate predictor for all milk traits.

The preceding equations provide an indication of the major contributing factors in the major milk and constituent yields. They strongly that age and breed of dam and calf preweaning growth are highly associated with milk yield.

They also indicate that cow weight changes and call birthweight, although influenced by breed and age of dam effects, have little association with lactation performance. Further examination of these factors of follows in the individual multiple regressions performed.

The remaining analyses determined the specific effects of individual cow and calf variables on the dependent milk variables of average, June and September yields for 1976 and 1977.

Tables II.5 and II.6 show effects of calf sex, age, birthweight, cow winter weight loss and summer weight change from calving to weaning at approximately 150 days

of lactation on average daily milk yields for both years.

Effect of Age and Breed of Dam

When forced to enter the equation first, A, B and A x B interaction tegether accounted for 43.4 and 43.0% of the total variance in average daily milliples in 1976 and 1977 respectively. As indicated in Table II.2, dam age alone accounted for 24,6% and breed, 17.5% of additional variance.

Earlier analyses (Table II.1.) suggested there is a confounding effect exerted by age and breed of dam on calf birthweight, cow winter weight loss and cow ADG from calving to September. These results agree with reports of Singh et al., (1971), Fahmy and Lalande (1973) and Smith et al., (1976).

Many authors report significant, effects of dam age on milk production. This factor was discussed in some detail in the preceding chapter.

the effects of cow age alone accounted for 10.5 and 4.6% of the variance in milk yield. Breed of dam alone explained a significant 17 and 33% of the variance in average daily milk yield each year over the effects of calf and cow age in his study.

Gleddie and Berg (1968) estimated that 82.5% of the variance in milk yield was associated with breed of dam effects. However, this high estimate has not been reported in any other literature geviewed.

dividual cow and calf variables, Kinsella 1976. Regression of average daily milk yield (kg/d

Variable of Interest	Equation	Order of Forced Equation Variables		Regression Coefficient for Variable of Interes b (Kg/day) SEb	ficient Interest SEb	Total Variance Explained ER: x 100 (%)	Additional variance Explained by variable of Interest after Forcing other Variables (x):
Sex of Calf	1. 'a') (b) (c)	AXB. S				0.23	0.00
Ca14 Age	600	CA A. B. AXB. CA A. B. AXB. CA		3	0.00	44.22	0.78
Calf Birthweight'	(C)	BW AXB BW A B AXB BW A B AXB S	7.	0.00	000	20.22*** 44.95**	0.02
Cow ADS from Calvirg to September	4. <u>9</u> 90	COW ADG C-S A. B. AxB. Co A. B. AxB. S.	. g.,	0.79		45.39* 4.03** 44.44 44.63	1.17
Weight Loss*	я. СОД	WWES A. B. AxB. WW A. B. AxB. S.	CA, WLS	0.00	: 1 1 1	4.26** 44.47* 44.90*	0.28 0.68

.. S. CA accounted for 44.22% of total variance.

Reghession of average daily milk yield (Kg/day) on individual cowand calf variables, Kinsella 1977.

Variable of Interest	, Equation	Order of Forced Equation Variables	Regression Coefficient Sor Variable of Inte	lent Interest SEb	Total Variance Explained R' x 100 (%),	Additional Variance Explained by Variable of Interest after Forcing other Variables (%):
I 1977 Sex of Calf	· (• · ·	· · · · · · · · · · · · · · · · · · ·			4	•
	G G	A, B, AxB, S			0.14	6 0 0
2			1	3c.	43.12	0.01
	% € Ō.⊙	CA A, B, AxB, CA A, B, AxB, S, CA	and the second	0.0	43.03	0.01
Calf Biethweight	3 / 8				43.12	60.0
		A. B. AxB. BW A. B. AxB. S. CA. BWI	0.04	0 0 0 0 0 0 0	21.00***	1.25
Cow Abd * Calving to September*	. 4. (0.00)	Cow ADG C-S A, B, AxB, Cow ADG A, B, AxB, S, CA, Cow	ADG 🥰	0.0 c	0.20	0.75
Cow Winter Weight Loss	 	WWLS A. B. AXB. WWLS A. B. AXB. S. CA. WWLS	olt	,,' 		0. 0. 32

ccounted for 43.02% of total variance, in 1977; A,

In the present study, the addition of calf sex (S) and age (CA) to the forced variables each year accounted for only 0.1 and 0.8% of the variance over and above A, B and A x B. All these factors accounted for respectively 44.2 and 43.1% of the total variance in 1976 and 1977.

Effect of Sex of Calf

Sex of calf accounted for only 0.23% (1978) and 0.14% (1977) of the total variance in average daily milk yield. Following the introduction of A, B, A x B and CA; calf sex was treated as a non-continous variable with no numerical value assigned to it, the regression coefficients computed

difference between male and female calves.

illa et al. (1974) reported that calf sex was not

ricantly associated with milk production while Rutledge

et al., (1970) found that females took more milk than males.

Neidhardt et al., (1979) however found that the sex of calf

exerted more influence on milk yield than age of dam or the

dam's body weight at calving.

Peschiera (1966) reported correlations of 0.6 between calf sex and average daily milk and moted that male calves were more vigorous sucklers than females. Richardson et al., (1977) found that when milk yields were corrected for birthweight, sex did not have as strong an effect. Marshall et al., (1976) found that sex had no significant effect on milk production although the breed of dam x sex of calf interaction was significant.

Although results in published reports vary considerable; segarding the significance of sex of calf on milk yield, most agree that makes, because of their larger size, tend to suckle more frequenctly than females.

Effect of Calf Age

Calf age (or days in lactation of the dam) in September did not account for a significant percentage of additional variance in average daily milk yields each year. The regression coefficients also indicate no association (b=0) between calf age and milk yield (Tables II.5 and II.6) unlike the results using stepwise regressions of combined cow and calf factors. Milk yield and calf age were not significantly correlated with r values which averaged 0.03 (Table II.4).

The results of today do not agree with those reported by Rutledge et al., (1974), Neville et al., (1974), Marshall et al., (1976) and Neidhardt et al., (1979). These workers found that age of calf significantly affected milk production and generally exerted a quadratic effect. Neville et al., (1976) reported that a 1 day increase in calf age was associated with 0.014 Kg/day increase in milk production. Both Marshall et al., (1976) and Neidhardt et al., (1979) toted that milk estimates decreased as calf age increased, indicating a decline in milk production as lactation progressed. Similarly, Gleddie and Berg (1968) computed a regression of milk on day of lactation (age of calf) of 0.02 Kg with a significant correlation of -0.46.

No differences in average milk yield were reported by Richardson et al., (1977) when calves were weaned at 150 rather than 240 days. However, cows with early-weaned calves were heavier and gave birth to heavier calves the following year because a long dry period permitted body weight to increase. Richardson et al., (1977) did not specify the advantages of the latter system.

Calf birthweight accounted for 20.2% (1976) and 21.0% (1977) of the variance (P<0.01) when all other variables were ignored. A 1 Kg increase in BW was associated with a 0.13 Kg/day increase in daily milk yield. After main effects were forced to enter the equation first, BW accounted for an additional 1.17 and 1.19% of the variance, suggesting that the combined effects of A, B, A x B and S removed a large proportion of the variance associated with calf wirthweight.

Although calf birthweight and milk yield had moderately high correlations of 0.5 (P<0.01), similar to those computed by Rutledge et al., (1971), increasing birthweight by crossbreeding and selection is not a favourable method for increasing milk production. The incidence of dystocia is widely reported (Smith et al., 1976) when birthweight is increased. This factor will be discussed in Chapter III.

However, results suggest that the effects of age and breed of dam and calf sex confound the individual effect of birthweight as a factor which influences milk production. Richardson et al., (1979) demonstrated with regression

with calf birthweight both of which reflections of which and age of dam. Rutledge et al 1971) noted that were with higher birthweight consumed more and were more vociferous. However, these authors contended this was more elated to an overall larger size than birthweight per se. Effect of Cow Average Daily Gain from Calving to September

Cow ADG during lactation accounted for only 4.0% (1976) and 0.2% (1977) of the total variance when all other variables were ignored. It accounted for less than 1.0% each year after the introduction of main effects. The b coefficient prior to the introduction of any variable for 1976 indicated a negative association between the milk and cow daily gain. The correlation between the two variables was low and negative.

Wilton (1976) observed that dairy cows that gained the most weight while lactating were least efficient. Most workers reported that cow weight loss increased with milk produced, suggesting that for high-producers, milk yield is maintained at the expense of body weight (Jeffers et al., 1971; Hohenboken et al., 1973 and Niedhardt et al., 1979). McGinty and Frerichs (1971) however found that among crossbred Hereford cows, those producing more milk did not necessarily lose the most weight. Economides et al., (1973) found that this factor is heavily influenced by feeding as cows that were fed only to maintenance lost more, weight and suffered a

drastic reduction in milk yield compared to others on a

Richardson et al., (1977) found that although weight change during lactation was negatively related to milk in the first 35 days, it had no significant effect from 36 to 91 days post-partum. These workers found that fat content increased five weeks after calving with an increase in cow weight post-partum (r= 0.2, P<0.05). They attribute the BF% increase to plane of nutrition rather than normal lactation trends for BF% which increased several weeks post-partum. Effect of Cow Winter Weight Loss

The results of this study indicated little association and relationship in this study of milk yields the winter weight loss from October of the preceding year calving. Before the introduction of main effects, cow we for 4.3 and 2.6% of the variance in 1976 and 1977 respectively, but less than 1% after the main effects were accounted for. The b coefficients were slightly negative but not significant.

Factors Influencing June and September Yields

Tables II.7 through I 10 present ordered regressions of June and September milk yields on individual cow and calf variables for both years. The procedure involved in the following regressions is as described for average milk yield. As results for 1976 were similar to those for 1977, only 1977 figures are discussed.

Regression on June milk yield (Kg/day) on individual cow and calf variables, Kinsella 1976.

Variable of Interest	Equation	of Forced fo	(Kg/day)	٠. هم ١ در	Total Variance Explained R' x 100 (%)	A C	Explained by Variable of Interest after Forcing other Variables (%):	
1976	3	3	•					~
Sex of Calf*		, i	-0.09	-0.0	2.35		•	•
	6 (A.63, AX63, S	0.08 0.08	8	28.29		90.0	0.25
	3	A.B. AXB. CA. S	60.0	0.03	28.70		0.48	0.17
Calf Age (June)	2. 8)	CA.	-0.0	-0.003			•	4
	<u>.</u>	63	-0.01	-0.07	28.63		43	-0.00
	ີ່ເບ	A.8.Ax8.S.CA	-0.01	-0.68	28.70	•	•	
Calf, Birthweight'	3.	3) ja	5
	(q	A.B. AxB. CA	200	÷ ;	13.95			33.08
*	ົບ	A.B. AXB. S. CA. KW	0.0	 	29.36		08.0	2.13
Coe. Ang.		•						2.
from calving	. .	A.B. AxB. Cov. And	5.25	0.0	2.75	•	•	4.32
to dune	0	A.B. AXB. CA.S.	, C	2 6	9 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6		0.00	2.27
, , , , , , , , , , , , , , , , , , ,		COW ADG, J-S		}	97.87		88.0	1.56
Cow Winter			6					
Meight Loss) }	· ·	5 8		3.50	•	•	7.60
	0	A W AKE CA	7.0	5 6	28.30		8 0.0	0.24
	7	•	3		7 28 71		0.04	0.03
Calf ADG	6. 8).	Calf ADG B	0.67	36	100			
th to June;	Ĝ	A, B, AxB,	0.56	0.76	16.00 A6.03			128.65
	,	Calf ADG. J-S	. /		**		70.0	63.34
	0	Calf And Lan	0.58	0.48	47.44		18.75	66.69

1-continuode variable - no value assigned

Regression of Jun

Variable of Interest	Equat 1on	Order of Forced Variables	Regress for Van	Veriable of	fficient Interest SEb	Explain	(X) of	Addigional Explained k Interest	fonel Variance fred by Variabl rest_after Forc	<u> </u>	· u .
977			-	de .				Ge		-	٠. د
Sex of Calf	G:	,		-0.04	0.14			,	•		
	60	A.B. AXB. S		0.03	0. 13	30.67		J	90	°	8.0
	×.		n	8	0.13	.31.02			0.08	, ,	0.23
Carr Age:	2. (9)	5		•	0.01	7				•.) }
	` G G	A, B, AxB, CA		0.01	0.91	30.84			.27		÷.
	3 .	J. C. GYW. G. K.	, , , , ,	0.01	0.01	31.02	1 y		0.17	- :	12.0
Calf Birthweight'	3. a)	2	ı, ·	0.16	0			•	• .		
•	; `a`	A.B. AXB. BV	1	8.9	0.0	9			, ;	ਨ ੰ	F. 62
	ò	A.B.AxB.CA.	S.BW	10	0.03	32.56			4 K		
COW ADG from	4. 83	8-0 908 A00						-		•	99
Calving to	a A	A.B. AXB	•		0 0	1.68		• .	,	e.	+1
September	•	GOW ADG C-S	•	3		31.18		0	0.51	. E.	1.54
	0	A.B. AXB. CA.	'n.	-0.25	0.27	34.32			08.0		
		COM ADG C-	V 1		·.			, 	3		C
Cow Winter	5. B)	WALS		0.0	**			ا نه	ı		
	ີ ົ	A.B. AXB. WWL	.		ı	30.67	*,			φ	6.62
	3 .	S. WILS	.		•	31.03	•		e n	• c	3 8
	•			· .				.		•	3
Riveh to bush	(9	Calf ADG B-	-	6.90	0.60	33 48		6	•	•	
	õ	A.B. AXB.	-65 (-8	56.20	0.70	46.96	ŕ			135 7.5	35.85
	(0	A.B.AXB.CA) 'L	8	4 CF C	4		٠		?	
	集 /	Call Abg. J-			?		i.	₽,	18.38	74	7.
			.	•.	• •		•	,	*		

on-continuous variable

X

Ka/dav

A. G. AxB explained 30.67%; A. B. AB. CA C explained 31.00

Table II.9. Regression of September milk yield on cow and calf variables, Kinsella 1976.

Variable of Interest	Equation	Order Reg of Forced for Variables b	Regression Coefficient for Variable of Interest b (Kg/day) SEb	officient Interest SEb	Total Variance Explained R' x 100 (%)	Additional Variance Explained by Variable of Interest after Forcing other Variables (%):	L.
I. 1976			•				
Sex of Calf!	- B	S		9			
	Q	A.B.AXB.S	5	2 6	06.0 0	•	5
	ົບ	A.B.AxB.CA.S	-0.05	8 8	45.11 45.07	0.04	0.15
	,				. n . c .	0.86	0.24
. DO - 100	•	C.A.	-0.001	, 00 00	0		
•	۵	A.B.AXB.CA	-0.012) (i	•	4.71
	ົບ	A.B.AxB.S.CA	-0.012	0.0	45.90 50.80	0.03	2.90
Cale Distriction					n .	0.07	2.97
Canal District Banks			0.7	000	09 11		
•	Q.	A.B.AxB,BW	90.0	3 6	09.71	•	43.76 ′
	ົວ	A.B. AXB. CA C. RW		. 0	46.72	1.63	7
3	•	•		20.02	47.18	1.22	7
COW ADG from		COM ADG 11-5	27.0	9			- - -
June to Sept.	<u> </u>	A R AVR COU ADO	· ·	27.0	5.07	1	0 0
	ì	2.00 mon		0.27	45.14	0.01	
	6	0 40 0×4 8 4				•	
		0.40.004.000 0-1. 004.000	5.0	0.24	46.08	0.12	;
		•					-
Cow Winter	9.	MALS	0.01	2			
Weight Loss'	â	A.B. AXB. WWLS	ç	38	8:	•	90
	ົວ	•		5 3.0	47.42	2,35	E 12
	•			-0.22	47.81	2.	
			i.				• • •
Calf ADG J-S.	6. a)	Calf ADG J-S	3.32	76 0	4		
	٩		1.67		0/.87	•	82.15
•		Calf ADG U-S	;		49.26	4.19	15.60
	ີບ	A.B.AxB.CA.S	2.55	0	200	•	
		Calf ADG U-S			22.86	7.26	29.00

non-continuous variable days days Kg Kg/day

Table II.10. Regression of September milk yield on cow and calf variables, Kinsella 1977.

Variable of Interest	Equation	Order of Forced Variables	Regression Coefficient for Variable of Interest b (Kg/day) SEb	ifficient Toterest SEb	Total Variance Explained R' x 100 (%)	Additional Variance Explained by Variable of Interest after Forcing other Variables (%):	i.
I. 1977 Sex of Calf	-	·	Ç				
	<u>.</u>		2.0	0.10	0.42	•	- .01
	a`	A.B. AXB. S	-0.07	90.0 0	46.86	0.18	0.77
	Ĉ	A,B,AxB,CA,S	90.0-	0.08	47.77	0. 15	9.0
Calf Age'	2. 8)	V	-0.02	0.01	1.40	ı	Ž
	<u>a</u>	AC. B. AXB, CA.	-0:01	· •	47.62	186) y
	ິ ບ		-0.01	1	47.77	16.0 16.0	3.92
Calf Birthweight	3. 8		6		6		
•	î q		2 6	5 6	77.77		57.32
,	` ()	A		2 6	80.74	.01	4.37
	•				10.10	69.0	8.8
Cow ADG Calving	4. 0)	COW ADG J-S	-0.81	0.40	1.66	1	•
to September*	<u>م</u> م	A.B. AXB, COW ADG		0.34	46.76	0.10	0 33
	,	٠ - S				•) !
	(ΰ		-0.36	0.35	48.03	Ó.25	11.11
		COW ADG, U-S					
Cow Winter	. a)	MALS	0.01	ı	35	,	, ,
Weight Loss	Q	A.B. AXB, WWLS	-0.01	•	47.60	0.92	64. C
-	(၁	A.B.AxB.CA.S WWLS	-0.01		48.65	0.88	3.84
Calf ADG J-S.	6. a	Calf ADG U-S	5.25	0.48	33,44	ı	9
	Q	A,B,AxB,	3.21	0.53	30.45	7.36	36.38
	•	Calf ADG J-S					
	(ပ	A.B. AxB, CA. S	4.07	0.56	57.62	99 · 90	52.05
					•		

hon-continuous variable days

Ag/day

A. B. AxB explained 46.68% of variance; A. B. AxB, CA, S : 47.77%.

The effects of A, B and A x B accounted for 28.2 and 46.7% for 1977 June and September milk yields respectively. The addition of calf sex and age to the equation accounted for a small amount of additional variance, but both showed little association.

As demonstrated in Table II.4, correlations between dam breed and milk variables were generally higher for September. This may suggest that the effect of breed of me exerts a stronger influence on persistency of milk over the course of the lactation period. The milking potential of the range cows cannot be realized if the calf's capacity for ingestion is limited in the early stages of lactation.

The association of calf birthweight with both June and September yields was similar, although correlations indicated a stronger relation between September yields and birthweight (Table II.4). The simple regression coefficients were highly significant (P<0.001). Every 1 Kg increase in birthweight resulted in an average increase of 0.14 Kg/day in average June or September milk yield. This may suggest that the phenotypic effects of birthweight, confounded by age and breed of dam, have a persistent association with milk variables.

Cow weight changes before and after calving accounted for little of the variation in either June or September milk yields after the main effects were accounted for. Cow ADG from calving to June had a positive association with June milk, but b coefficients were negative when cow ADG from

calving to September was regressed on 1977 September milk yield.

As expected, calf ADG during the preweaning period had a very high association with both June and September yields, accounting for 9.85 and 18.38% of the additional variance over main effects. The b coefficients were also high and significant, but were lower for September yields than for June.

D. CONCLUSIONS

Data from a milking experiment were collected to examine various factors affecting lactation performance of Hereford, Beef-Synthetic, Dairy-Beef and Dairy-Synthetic cows. Regression equations were computed with dependent variables of average, June and September milk yields and constituent yields. Independent cow and calf factors examined were cow age, breed, age x breed interaction, winter weight loss during gestation, weight change during lactation, calf sex, calf age, birthweight and ADG from birth to September.

Age and breed of dam together exerted major effects on all milk variables, accounting for between 34.0 and 44.0% of the total variation. It is not clear which of age or breed of dam accounted for more of the variation. However, breed of dam was more associated with milk yield variables in September than in June, perhaps suggesting that under range suckling conditions, breed of dam effects are more pronounced

later in lactation.

when included in the equation, calf preweaning ADG showed a predictably strong association with all milk yields, accounting for up to 24% of total variation in any one milk or constituent yield variable. The correlation coefficients of calf ADG with any milk yield variable averaged 0.5. Results of the stepwise regressions indicated that every 0.1 Kg/day increase in calf ADG was associated with 0.62 Kg/day increases in average daily milk.

Calf age accounted for a variable proportion of the variation in milk yields depending on the order of variables entered into the equation. In all the equations however, the partial b coefficients were low or negative, contrary to results in other published material.

Calf birthweight did not account for a significant proportion of the variation after accounting for the effects of age, breed, age x breed of dam interaction and calf sex. These effects removed a large proportion of the variation associated with calf birthweight.

Cow weight changes during pregnancy and lactation were not significantly associated with any of the milk yield variables.

A highly significant relationship was found for average milk yields with constituent yields (r=0.9). The similar association between independent cow and calf variables and each of the dependent variables indicated that little additional information is provided by using more than one

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measure of lactation performance.

III. FACTORS INFLUENCING WEANING WEIGHTS OF RANGE BEEF AND

A. INTRODUCTION

Preweaming average daily gain is an important factor affecting the profitability of a cow-calf operation. As weaning weight and preweaming average daily gain are essentially the same measure of growth (Kennedy and Henderson, 1975), and realized response to their selection is favorable (Preston and Willis, 1974), selection for either trait should be incorporated into an efficient beef breeding program.

Among the factors affecting preweaning performance, milk has been shown to exert a major influence on range beef calves. Measurements of milk yield provide a valid prediction of associated response in calf preweaning growth rate (Jeffery and Berg, 1971; Marshall et al., 1976; Spelbring et al., 1977a, 1977b.).

Studies have indicated that the simple correlation between average daily milk yield and weaning weight is only moderately high, suggesting that factors other than milk yield account for additional variation in weaning weight (Gleddie and Berg, 1968; Butson and Berg, 1980). The present study was conducted to examine the influence of milk yield, milk constituent yields and other cow-calf variables on weaning weight.

B. MATERIALS AND METHODS

The breeding plan, general management and method of milk removal were discussed in Chapter I. Data on calf birthweight (BW), calf age (CA), sex (S), weaning weight (WW), cow winter weight loss from October of the preceding year to day of calving (WWLS), cow average daily gain from calving to weaning in October (ADGCO) and September averages of daily yields of milk, butterfat (BF), protein (PROT) and lactose (LACT) were collected from an experiment involving measurements of lactation in range beef cows conducted in 1976 and 1977.

In 1976, 206 cow-calf pairs consisting of 45 Herefords (HE), 102 Beef Synthetics (SY), 26 Dairy-Beef (DB) and 33 Dairy-Synthetics were studied. The following year a total of 242 cow-calf pairs were studied and included 58 HE, 123 SY, 33 DB and 28 DS.

Statistical Analyses

The method of statistical analysis was similar to that outlined in Chapter I for least squares analyses of covariance for unequal subclass numbers. Least squares constants for age of dam (A), breed of dam (B), age x breed of dam interaction (A x B), sex of calf (S) and sex of calf x breed of dam interaction (S x B) were computed and used to calculate least squares means for calf weaning weights.

Stepwise multiple regressions of weaning weight on cow and calf variables for both years were computed using the SPSS REGRESSION as outlined by Nie et al., (1975). The

effects of variables on weaning weight were calculated according to two models of regression: a) ignoring the effects of A, B and A x B and allowing variables of interest to enter the equations sequentially based on partial correlations and c) forcing A, B and A x B to enter the equation first. The second method permitted examination of the percent of additional variance explained over and above the effects of A and B.

Ordered multiple regression models of weaning weight on individual milk yields, component yields and cow and calf variables were computed according to the method outlined by Overall and Klett (1972) and described in Chapter II. These permitted the examination of the influence of each of the variables of interest while other variables were either ignored or accounted for in the equations. As the interaction of S x B was not a significant source of variance in the analyses of covariance, this effect was not included in the regression equations. Regressions were computed by: a) ignoring A, B, A x B, S and CA and forcing in the variable of interest; b) forcing the effects of CA and S to enter the equation first before the variable of interest, and c) forcing all of A, B, A x B, CA and S to enter first before forcing in the variable of interest:

C. RESULTS AND DISCUSSION

Analyses of covariance of weaning weight (Table III.1) indicated that for both years the effects of breed and age of dam, and sex and age of calf on weaning weight were highly significant (P<0.01).

Effect of Breed of Dam

Least squares means and standard errors of weaning weight by breed and breed-age groups of dam are presented in Table III.2 for 1976 and 1977. The overall average weaning weights were similar for both years, averaging 192.0 ± 2.0 Kg in 1976 and 193.4 ± 1.7 Kg in 1977.

In 1976 cows in the DS line weaned calves that were significantly heavier than calves in other breed groups (P<0.01), averaging 47.2, 22.3 and 19.3 Kg more than HE, SY and DB calves respectively. However, in 1977, calf weaning weights of the DB line were significantly heavier than those of the HE and SY breed groups (P<0.01), weighing 41.8 and 16.0 Kg more than HE and SY calves respectively. Preliminary analysis of data showed that DB and DS cows produced more milk and were more persistent than those in the HE and SY groups.

Generally calves from dams with dairy breeding tended to exhibit higher weaning weights than those from either straightbred or crossbred beef breeding (Figures III.1 and III.2). These results are similar to others who compared dairy-cross cows with beef cows (Brown et al., 1972) and Wyatt et al., 1977b). Problems were noted by some workers in

Table III.1. Analyses of variance of weaning weight. With age of calf as a covariate.

	į			
Source	1976	1977	1976	1977
Breed of dam (8)	3	е ,	10377.9**	11745.7**
Age of dam (A)	ო	6	3319.9**	7910.0**
▼ ×	o	o	233.0	350.5
Sex of calf (S)	-	-	4083.5**	(5.0.6) 11 B.0
8 × 8	င	6	88.8	1.525.2 ST
Age of calf (AC)	Ţ	-	40876.6**	4109)
Error	185	221	338.3	312.9
Total	205	241		

P<0.01

á	
(K	
6	
5	
6	
	1977
ں •	976-
407	
•	Ĕ
200	×
ž E	ð
Ě	•
100 E	2
•	8
3	2
.2 Least aquares means and standard errors of calf weaning weights (Kg) by	breed and age groups of dams, Kinsella 1976-1977
Table 111.2	
-	

			1976			1877	
	D	No. of	16.		Mo. of	, C	S
Grand Total	114	206	192.0	2.0	242	183.4	1.7
	Ĭ		167.14	-	2		•
	S	102	192	7	123	192 25	-
	8	*	194.80	0	E	208 20	4
	80	33	214.20	8 .0	28	206. Bc	7.7
Age in years			•				
~	A11	7.8	20.00	2.4	83	175.2v	2.2
	Ĭ	=	151 0		•	141.7	•
	×S	8	173.0	-	É	17.3	
	8	2	-	4	2	9	•
	80	£.	5 86	W.	*		e W
m	~	38	187.0v	4.0	53	194.8x	0 E
	¥	•	162.8	7.0	^		•
	۲	5	184.0	-	8	183.0	
	8	•	197.3	•	•	2.59.5	•
	98	•	203.9	7.0	•	308	a b
•	F 1	8	203.8x	9.0	*	196 . 6x	•
	Ĭ	wî.	175.3	-	•	172 6	-
	S	5	202.2	a	-	184	
	2	•	218.0	5		7	5
	0.5	· 69	220.0	6.7	~	214.2	13.2
Meture	Ē	75	196. fx	¥.	4 2	204 Bx	2.9
	¥	13	178. 1	6 , 0	25	178.8	0
	24	Ę		9.7	;	207.7	4.4
	28	•••	- 14	# #	•	223.4	0
							,

7

 a.b.c - Least aquares means of vesning veights by breed of dam vithin years with different alphabetic latters are significantly different (PcO.01).

w.x - Least squares means of vesning veights by age of das within year with different alphabetic letters are significantly different (P<0.01).

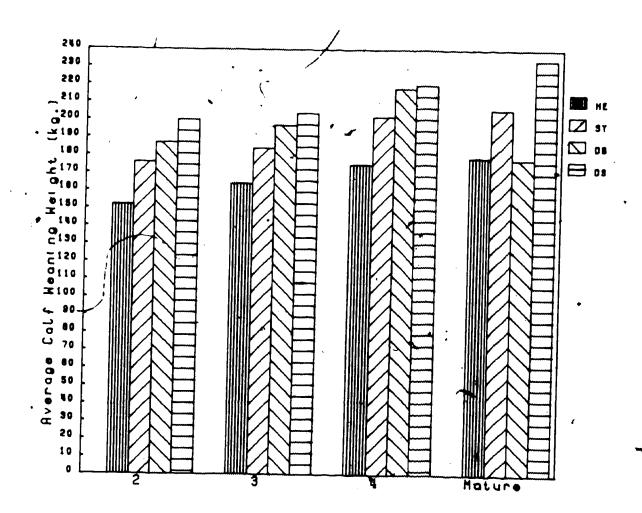


Figure III.1 Average calf weaning weighty by breed and age of dam, Kinsella 1976

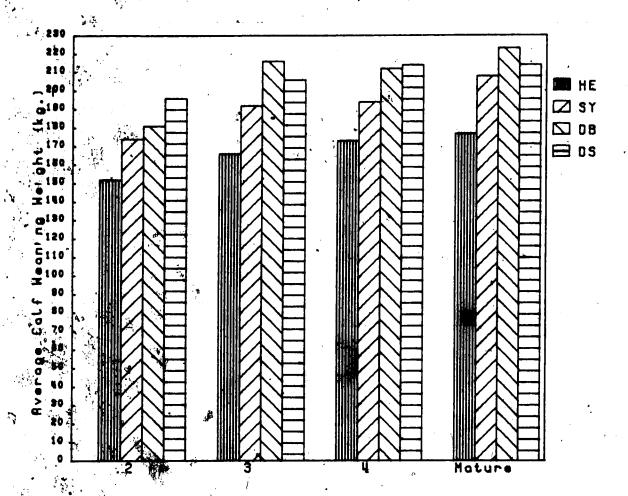


Figure III.2 Average calf weaning weight by breed and age of dam, Kinsella 1977

that high-producing cows tended to have lower reproductive rates (Deutscher and Whiteman, 1971; Bair et al., 1972). However, the selection system in the Kinsella experimental herds emphasized regular reproduction. Although this selection may have eliminated the highest milk-producing cows, overall calf crop percentages averaged 84% for SY dams and approximately 78% for HE and DS dams (Berg, 1978), indicating that dairy-cross cows in the present study were able to maintain comparable reproduction to beef cows.

Effect of Age of Dam

Least squares means for breed-age groups of dams (Table III.2) indicated that weaning weight (156 days of age) in 1976 for calves ranged from 151.0±5.6 Kg for calves of the HE 2-year olds to 233.9±5.8 Kg for calves of mature DS dams. In 1977 the range of calf weaning weights (159 days of age) followed a similar pattern of from 151.7±4.1 Kg for calves of 2-year old HE to 223.4±9.0 Kg for calves of mature DB dams.

Although observed differences in adjusted calf weaning weight between the 4-year old and mature dams were small and not significant (P<0.01), increases were noted from 2- to 4-year old dams, similar to the general trend of other published reports. Fahmy and Lalande (1973) reported that Shorthorn calves with maximum weight at weaning were those born to dams averaging 8.2 years of age. Brown et al., (1970) found that the quadratic effects of age of dam were highly significant for weaning weight, demonstrating a rapid

incline from 3 to 6.5 years, a gradual incline to 8.5 years and a decline to 11 years. Schaeffer and Wilton (1974) observed significant effects (P<0.05) of the age of dam with herd performance interaction on calf preweaning ADG and suggested that in herds of low performance an increase in cow age would produce a more noticeable improvement in calf gains than in high performance herds.

In the present study, the overall effect of dam age accounted for a significant (P<0.01) source of variance in calf weaning weight (Table III.1). Older cows tended to wean heavier calves, as shown in Table III.2, despite the lack of statistically significant differences between all but the 2-year old age groups in 1977.

Effect of Age x Breed of Dam Interaction

The age x breed of dam interaction (Table III.1) was not a significant source of variance indicating that the differences among breeds were similar for the different ages.

Effect of Sex of Calf

Sex of calf accounted for a highly significant (P<0.01) source of variance in weaning weight (Table III.1).

Within each year all males were significantly heavier at weaning (P<0.01) than all females (Table III.3). The least squares means of overall weaning weights for males and females (Table III.3) were 197.4±2.5 Kg and 186.7±2.6 Kg respectively in 1976 (a 5.5% difference) and 199.7±2.1 Kg and 187.1±2.2 Kg respectively in 1977 (a 6.3% difference).

Least squares means and standard errors of weaning weight by sex of calf and breed of dam, Kinsella 1976-1977. Table III.3.

			Males		•	Females		.*
	8 e c	No.	Weaning Weight (kg)	ж. ш.	Š	Weaning Weight (kg)	S.E.*	Male-Female Weaning Weight Difference (kg)
1976	ııĄ	, 106	197.4	2.5	8	186.7	2.6	10 7**
	뿟	25	170.4	œ e	Ç	163	,	(
	λS	48	8 961	, c	4 E		- (6.7
	90	Ť	20.4		9 -	707	7.00	9.7
	2			•	4	188.0	6.7	13.7
	Š	11	220.7	4 .9	1 3	207.7	5.6	13.0
1977	All	125	1.89.7	2.1	117	187.1	2.2	12.5**
	¥	17	169.7	6 0	22	163	•	•
	SΥ	57	196.6	4 0	, K		•	6.7
	90	23	219.9	.) .		4.	10
	SO	28	2,12.5	. 4	- C	201.2	- 0 - 0	23.5**

S.E. * standard arror: ** DVO O4

Results concur with those reported by Bair et al., (1972) and Fahmy and Lalande (1973), but were lower than the 8.8 and 8.3% differences found by Bailey et al., (1975) and Marshall et al., (1976) respectively among calves weaned at 8 months. Preston and Willis (1974) noted that sex differences in weaning weight varied with calf age at weaning and reported that studies conducted on Brahman-cross calves indicated that no sex differences were observed because calves were weaned 90 days before sex hormone influences were manifested.

Effect of Sex of Calf x Breed of Dam Interaction

The sex x breed interaction (Table III.1) did not account for a significant source of variance in calf weaning weight. Smith et al., (1976) reported that less than 1% of the accountable variation in birth weight was explained by the sire breed by sex interaction effect, despite the significant difference (P<0.01) at weaning between males and females.

Effect of Age of Calf

The average ages of calves at weaning in 1976 and 1977 were respectively 156.3 and 158.8 days, with a range of from 122 to 180 days in 1976 and 132 to 189 days in 1977.

As expected, calf age at weaning was a highly significant (P<0.01) source of variance in weaning weight (Table III.1) and was moderately correlated with weaning weight, with r values of 0.55 and 0.48 (Table III.4) for 1976 and 1977 respectively. Schaeffer and Wilton (1974)

Phenotypic correlations of average milk yield and cow and calf variables: 1976 data coefficients above diagonal (N=206); 1977 data coefficients below diagonal (N=242). Table III.4.

Variables*	Ave. M11k	GOW	V AOO	Calf	Ca) f	Calf	Calf
						AUGBU	\$
Milk		.21	07	01	4.	.67	09
COW WWLS	. 16		09:	. 10	. 42	. 17	. 24
COW ADGCD	10	. 67		23	. 26	12	16
Calf Age	03	.00	26		- 14	71.	, 10
Calf BW	. 46	. 38	9.	12	•	.38	4.
Calf ADGBO	.71	60.	20	. 18	<u>16</u>		6
Calf WW	.62	7	22	87	6		

Cow WWLS-winter weight loss: Cow ADGCO-cow ADG from calving to weaning in October: Calf BW-calf birthweight; Calf ADGBO-calf ADG from birth to weaning in October; Calf WW-calf weaning weight. *Variables:

r = 0.14 significant at P<0.05; r = 0.18 significant at P<0.01 (d_f = 200).

cited work that indicated a linear relationship between age at weaning and ADG for 120 to 250 days of age.

Stepwise regressions of weaning weight on cow and calf variables are presented in Table III.5. When weaning weight was regressed sequentially on cow and calf factors (Equation 1a and 1b) calf age accounted for 25.5 and 25.2% of the additional variance over milk variables in 1976 and 1977 respectively, similar to the 21.2% explained by calf age as noted by Lawson (1976). Moreover in both 1976 and 1977 a 1-day increase in calf age was associated with a significant 1.3±0.1 Kg increase (P<0.01) in weaning weight.

In equations 2a and 2b (Table III.5) effects of age of dam (A), breed of dam (B) and the age x breed (A x B) interaction were forced to enter the regression first. The total of these effects accounted for 47.6 and 45.3% of the variance in weaning weight in 1976 and 1977 respectively. Calf age (CA) however still accounted for 20.4 and 19.5% of additional variance over and above these effects. The partial regression of weaning weight on calf age remained highly significant (P<0.01) with b coefficients of 1.3±0.1 and 1.2±0.1 for 1976 and 1977 respectively.

Despite its marked influence on weaning weight, the average age of calves at weaning can only be effectively increased by either prolonging the preweaning period or by reducing the calving interval. Optimum weaning time would be determined by grazing conditions, potential harmful effects on cow conditions, subsequent reproduction and wintering

Table III.5. Stepwise regressions of weaning weight on cow, calf and milk variables, Kinsella 1976-1977

Estation #		Variables entered sequentially into forward regression	% Total variance explained (R' x 100)	% Additional variance explained	Partial b	SED?
3	9.76	Sept. BF Yield Calf Age: Calf Birthweight Ave. Milk	42.2 67.1 73.1 75.4	2. 20 Rb C C - 20 Rb C C C C C C C C C C C C C C C C C C	***	1.011
b) 1977	977	Ave. Milk Calf Age. Calf Birthweight June Milk	38.4 63.6 74.7 75.9	255.2 11.1	* * * * * *	1.011
a) 1976	976	A, B, A × B Calf Age: Sept. BF Yield Calf Birthweight	47.6 68.0 76.1 78.6	4.0.60	* * ! !! ! !	1011
2. b) 1977	7.76	A, B, A x B Calf Age: Ave. Lactose Calf Birthweight	8. 35 8. 35 8. 35 7. 08	- 00 m - 00 m - 00 m	1.2*	10.1

Days Standard error of b

** P<0.01

costs. Bailey et al., (1975) reported that calves weaned at 10 months of age were heavier than those weaned at 8 months (P<0.05) but that liveweight gains would be largely affected by grazing conditions and stocking rates. Furthermore, these workers reported that cows suckling their calves lost more liveweight, particularly during the period from weaning to calving.

Effect of Milk Yields and Component Yields

Earlier analyses indicated that simple correlations of average milk yield with component yields of butterfat, protein and lactose were high, ranging from 0.8 to 0.9 for both years (Chapter I). These figures were in agreement with coefficients reported by Gleddie and Berg (1968) and Jeffery and Berg (1971). All milk yield variables in the present study were moderately correlated with weaning weight and averaged 0.60 as indicated for average milk yield in Table III.4. Koch (1972) calculated r values that ranged from 0.5 to 0.8 for average milk weight with weaning weight, similar to estimates reported by Gleddie and Berg (1968) and Marshall et al., (1976).

These relationships may suggest that a measure of association of one milk yield variable with weaning weight would be expected to reflect a similar relationship as another variable. Jeffery and Berg (1971) reported milk yield alone was as adequate as any other single milk variable as a measurement of associated response in preweaning growth rate, and that the addition of milk

component yields over milk alone had limited value in increasing the efficiency of predicting calf growth.

The 1976 September BF yield accounted for 42.3% of total variance explained in calf weaning weight (Table III.5). Average milk yield entered the equation after calf age and birthweight, accounting for only 2.3% of the additional variance.

As milk component yields were so highly correlated with milk yields, it is likely that the variation in weaning weight explained by September BF removed a large proportion of the variation in weaning weight accounted for by average milk yield. This also occurs in the 1977 results (Table III.5 Equation 1b) as average daily milk yield entered the regression first, explaining 38.4% of the variance in weaning weight uncorrected for age of calf. However, when calf age entered the equation next, milk yield and calf age accounted for 63.6% of total variance in weaning weight.

These age-corrected values are somewhat lower than those reported by others who studied the effects of milk yield on preweaning ADG. Gleddie and Berg (1968) and Jeffery and Berg (1971) found that average milk yield accounted for approximately 71% and 60% of the variation in calf preweaning ADG respectively. Rutledge et al., (1971) noted that on a within herd-year-sex basis, 60% of the variation in 205-day weight was due to the independent effects of the first four months of milk yield. Slen et al., (1963) claimed that as no difference was found in protein or butterfat

content (P<0.05) in five breeds of sheep these constituents had little influence on body weight gain of lambs or its correlation with milk production. All authors reported that the inclusion of milk yields in the later stage of lactation added little to the explained variance in either weaning weight or preweaning ADG.

In equations 2a and 2b (Table III.5) the effects of A, B, A x B were forced to enter the regression first. These effects accounted for 47.6 and 45.3% of the total variation in weaning weight for 1976 and 1977 respectively. Age of calf (CA), September BF and calf birthweight (BW) accounted for 20.4, 8.2 and 2.5% of additional variation in 1976.

Results in 1977 were similar with the exception that average LACT yield replaced September BF as a variable accounting for 10.6% of additional variation.

Ordered multiple regressions of calf weaning weight on individual average milk and component yields were computed for the 1976 and 1977 data and are are presented in Table III.6. Regressions were computed a) without accounting for any main effects; b) after accounting for the influence of S, CA and c) after accounting for the influence of S, CA and A x B.

Prior to the addition of any other variable, a 1 Kg increase in daily milk yield was associated with a 12.4 ± 1.2 Kg and 11.3 ± 1.3 Kg increase in weaning weight each year respectively. However, after adjusting for the effects of S, CA, A, B and A x B, the regression coefficients decreased to

Table III.6. Regrassions of waaning waight (Kg) on individual silk variables, Kinsella 1976-1977.

		10000	Regression Coefficient	floient	Total Variance	Explained by Verige
Variable of Interest	Equation .	Variables.	for Variable of Interest b(Kg) Stb	Interest	Explained R' x 100 (X)	
1976 Ave Milk'		Ave #134				
	â	AC S ACS MAIN	7.0	~	35.0	:
	6		7.0	-	6.7	34.8
	•	S. Ave. M13k	•	0.	. 9 /	o∵ .
Ave. SF vield!	•					
		AVE. 67 T101G	7.7	0.2	28.8	
	ริจิ	Diatrice State	77	7 .0	63.8	28.7
	3	S. Ave. OF	•	0 7	76.4	
Ave and ave						•
	• 7	Ave. PROT YIELD	4.0	₹.0	32.4	, •
	3	7. S. W.C.	₩.	0 .0	3	91.6
	Û	A. 6, A × 6, AC.		•		' !
		S, Ave. PROT.		,	•	6.3
Ave. LACT YIEld!	(*	Ave. LACT Viets	•	•	;	
	â	AC. S. Ave I ACT	· ·	7 (7.7	:
		Yield	-	7.0	3	97.6
	ច	A. 6. A × 6. 16.	£	0.2		(
		S. Ave. LACT		1	•	•
1877						
AVE. MIIR!	(d .)	AVE. MISK	-			
	3	AC. S. Ave. MITH		- 6	7.1	• • •
	ê	A. G. A.X B. S.	W. V.	•		P. C.
		Ave. Miff		;		n in
Ave. 8F Yield!	6. 8	Ave. 87 7161d	2 .	•	,	
	<u>a</u>	AC. S. BF Yield	2.7		7.6	; ;
•	๋	A. B. A X B. S.	₽.₽		* * * * *	D (
		Ave. Hilk		!	?	0.
Ave. PROT YIELD!	7	Ave. PROT YIELD			;	
	a	AC S AVE	,	?	× :	:
		PROT YIELD	,	N.	0.6	37.6
	3	A. B. A x B. AC.		0.3	78.2	•
		S. Ave. PROT				
Ave. LACT YIELD!	•	Ave. LACT YIELD	2.1	,		
	a	AC, 5, Ave.	2.1			; ;
	7	LACT VIETG				,
	3	7		•		

NS - P.D. Kg AC and S accounted for 35.2% and 27.2% of total variance in 1978 and 1977 respectively; A, B, A × B, AC and S accounted for 70.2% and 67.5% of total variance in 1976 and 1977 respectively * AC * Calf Age; S * Calf Sex; A * Age of dem; B * Breed of dem; A × B * Age × Breed of dem interaction ' Kg ' p.01 Kg

7.8 and 7.5 in each year respectively. Milk yield accounted for an additional 8.0 and 9.5% of total variance over and above all other effects in 1976 and 1977 respectively.

Other workers reported a significant association of milk yield with preweaning growth measurement. Jeffery and Berg (1971) calculated increases of between 0.06 and 0.09 Kg/day in preweaning average daily gain for every Kg increase in the averages of August and October milk yields. For traditional beef breeds and crosses, Marshall et al., (1976) reported rather low responses of 1.5±0.4 Kg at weaning for every Kg increase in milk yield over the 5 Kg/day average. Rutledge et al., (1971) found that every 1 Kg increase in daily milk yield averaged over the first four months of lactation resulted in increases of only 2.5 Kg at weaning.

In equations which accounted for the effects of S and CA, percentage of total variation explained was 35.2 and 27.2% for 1976 and 1977. For S, CA, A, B and A x B, percentage of total variation explained was respectively 70.2 and 67.5% in each year. These effects accounted for somewhat less variation than those reported by Rutledge et al., (1971), who found that year, herd, sex, sires, milk yield and other cow and calf variables accounted for 92.4% of the variation in weaning weight.

Percentage of additional variation explained and regression coefficients for average daily milk constituent vields were similar with respect to association with weaning

weight. The average daily BF, PROT and LACT yields explained from 6 to 10% of total variation in weaning weight over and above effects of S, CA, A, B and A x B in both years. A 0.01 Kg increase in each of average component daily yields effected increases in weaning weights of approximately 1.4, 2.1 and 1.3 Kg respectively for both years (Table III.6). Effect of Cow Winter Weight Loss

The least squares means and standard errors of cow winter weight loss (WWLS) for both years are presented in Table III.7. The means of WWLS from October of the preceding year to post-calving for the relatively mild winters of 1975-1976 and 1976-1977 were 54.9±2.9 and 51.9±2.9 Kg respectively for all cows. Although differences between breed groups were not significant there were significantly different results within the age-breed categories. Between age groups the mature cows demonstrated a significantly (P<0.05) higher winter weight loss than younger cows. The noticeable difference between 2-year olds and other age groups in this study reflects the different feeding and management for first-calf heifers as these results are . unlike those reported by Deutscher and Whiteman (1971) and Wyatt et al., (1977b). These authors observed that cows with dairy breeding and 2-and 3-year olds lost more weight than did Herefords and older cows.

The correlations of WWLS with weaning weight and milk yield (Table III.4) were low, averaging approximately 0.20 each year. Results concur with those cited by Morris and

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Streed Number Sear S.E. Maar S.E. Maar S.E. Sear S						WRS	WLS' (kg)		
Total All 206 242 54.9mg 2.9 51.9mg 2.9 51.9		Deer B	1876	. 1977		S.E.	1	S.E.	
HE 45 B8 B1.58 4.7 B6.38 1.58 5.7 B6.38 5.8 5.8 5.8 5.8 5.8 5.8 5.8 5.8 5.8 5.	Grand Total	154	206	242	54.9m	2.8	91.9ms	2.8	,
SY 102 123 85.0e 3.1 95.6e 1		¥	4	3	51.5	4.7	25	4	
Years Year		\$	102	123	82.0	-	27.2		
Years 1	•	2	38	CC	61.28	1.1	51.8	•	
HE 11 15 82 11.9v 3.8 28.3v SY 36 36 11.3 4.8 34.4 SY 15 20 4.8 6.8 34.4 HE 7 7 7 61.8 10.1 81.3 SY 15 30 66.3 7.0 48.7 DS 6 8 6 8 6 7.5 11.5 38.4 A11 20 26 62.6x 7.2 52.5x HE 9 7.5 11.5 38.4 A11 75 82 82.14 7.4 77.7 HE 9 82.8 A11 75 82 12.7 A11 75 82 12.7 HE 9 77.7 HE 9 7	Age in Years	S	33	5	8 . 9	ю	46.7	•	
HE 11 19 16.9 8.1 34.4 SY 36 25 11.3 4.8 34.2 OS 13 8 24.3 7.7 8.0 HE 7 7 66.9 7.0 45.7 SY 15 30 66.3 7.0 45.7 OS 6 8 6.7 8.7 8.6 HE 9 7 7 86.0 65.7 HE 9 7 86.0 65.3 A11 75 82 82.19 7.4 77.7 HE 9 7.7 86.6 SY 10 14 61.8 93.8 OS 9 6.0 15.8 93.8 A11 75 82 82.19 7.4 77.7 OS 95.1 82.4 BY 10 14 61.8 62 87.3 37.1 OS 95.1 82.4 A11 75 82 82.19 7.4 77.7 OS 95.6 OS 95.7 82.4 OS 95.7 83.8 OS 95.7 83.4 OS 95.7 83.8 OS 95.7 83.4 OS 95.7 83.4 OS 95.7 83.4 OS 95.7 83.4 OS 95.7 83.8 OS 95.7 83.4	~	N I I	27	82	2	3.8	28.35	•	
SY 36 25 113 4.8 34.2 OS 13 8 24.3 7.7 8.0 HE 7 7 66.6 82.8 4.8 34.2 SY 15 30 66.3 7.7 81.3 SY 15 30 66.3 7.0 45.7 OS 6 8 67.8 11.5 98.4 A11 20 26 62.6 7.2 95.1 HE 9 7 96.0 (2.2 95.1 SY 10 14 61.4 18.9 93.8 OS 3 2 25 96.0 15.8 4.3 A11 75 82 82.19 7.4 77.79 HE 22 25 68.2 9.7 34.6 OS 31 44 117.8 4.2 76.6		¥	Ξ	Ş	0.91	-	7		
HE 7 7 61.6 77.7 80.7 7.7 80.6 82.8 7.7 80.0 80.0 80.0 80.0 80.0 80.0 80.		λS	ğ	g	=	-	75		
13 8 24.3 7.7 8.0 14 7 7 61.8 70 48.7 2.8 4.8 47.0x 15 30 66.3 7.0 48.7 2.8 13.3 38.4 4.8 7.0 7.0 4.8 7.0 7.0 7.0 7.0 7.0 7.0 7.0 7.0 7.0 7.0		8	.	8	•		20.7	-	•
HE 7 7 61 6 70 45 13 13 15 15 15 15 15 15 15 15 15 15 15 15 15		20	Ç	•	24.3	7.7	0	6	
HE 7 7 61.6 10.1 81.3 57 15.5 57 15.5 57 15.5 57 15.5 57 15.5 57 15.5 57 15.5 57 15.5 57 15.5 58 15.5	~ ,		36	13 13 13 13 13 13 13 13 13 13 13 13 13	62.8x	₩.	47.0x	9 .	
SY 15 30 66.3 7.0 46.7 5 10.0 10.0 10.0 10.0 10.0 10.0 10.0 10	ē.	¥	1	e K	• • • • • • • • • • • • • • • • • • • •	0	61.3	-	
Maria 20 26 82.6x 7.2 82.5x 11.5 38.4 11.5 39.4 11.5 39.		>	<u>.</u>	8	9	7.0	45.7		
A11 20 26 62.6x 7.2 52.5x HE B 7 56.0 (2.2 54.1 SY 10 14 61.4 6.5 6.2 DS 2 3 74.6 16.9 53.6 A11 75 82 82.19 7.4 77.79 HE 22 25 56.0 15.6 65.4 SY 41 44 61.8 42 76.6 DS 15 8 2 19 7.4 77.79		2	•	•	36 .7	•	o. 3	18.0	
HE B 7 86.0 (2.2 54.1 57.7 5 52.5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5	•	8	•	•	67.5	= .s	7.8	6 .3	
HE B 7 59.0 (2.2 54.1 57.7 59.0 (2.2 54.1 57.7 59.0 6.2 59.1 59.0 6.2 59.1 59.0 6.2 59.1 59.0 6.2 59.0	•	T.	8	9	62.6×	1.2	52.5×	7.5	
SY 10 14 61.4 6.5 53.8 53.8 53.8 53.8 53.8 53.8 53.8 53		¥	•	1	0.3	(2.2	-	9	
DS 2 3 74.8 18.9 83.6 DS 3 2 56.0 19.5 48.3 All 75 82 82.1y 7.4 77.7y HE 22 28 69.2 5.7 85.4 SY 41 44 81.9 4.2 76.6 DS 11 9 10 27.3 37.1		Š	2	=	4.19		53.8	7.7	
A11 75 82 82.17 7.4 HE 22 25 69.2 5.7 SY 41 44 81.8 4.2 BS 11 4 117.9 27.3 BS 11 9 50.8		8	ĸ	m	74.8		93.6	11.3	
HE 22 25 69.2 5.7 7.4 5.7 5.4 6.5 6.5 6.5 6.5 6.5 6.5 6.5 6.5 6.5 6.5		\$0	60	~	0.0	18.8	48.3	•	
22 28 69.2 5.7 41 44 91.8 4.2 1 4 117.8 27.3 15 8 88 8	Mature	I V	27	82	82. ty	1.4	¥7.77	. •	
41 44 65 6 4.0 1 4 17.0 (J.) (J.) (J.) (J.) (J.) (J.) (J.) (J.)		¥	22	25	69.2	9.7	7.0		
		51	= .	1 .	0.7	7	7.6		
		8 S	- =	• •	20.00	27.3	27.1 2.1	<u>.</u>	

ns - Grend total year means within a trait not significantly different (P<0.05)

S. cow winter, weight loss from October of the preceding year to callying:

a - Least sides as means of Wils Dy breed of dem within years with different

- Least aguares means of WMLS by breed of dam within years with different alphabetre's leaters are significantly different (P.O. 01)
- Least aguares means of WMLS by breed of dam within years with different alphabetre strategies and electricities and second

Ordered regressions of weaning weight on WWLS indicate little association with calf weaning weight (Table III.8)

After adjusting for S, CA, A, B and A x B (Table III.8

Equations 1c and 4c) a 1 Kg loss in winter weight was associated with no decrease in 1977 and a decrease of 0.1 Kg in weaning weight in 1976. Moreover, before adjustments WWLS accounted for only 5.8 and 2.1% of the total variation in weaning weight in 1976 and 1977. After adjusting for other effects it did not account for any of the additional variation.

Winter weight loss as it occured in these cows treated alike had neither a significant correlation with milk production nor did it account for a significant proportion of either total or additional variation explained in calf weaning weight. Further investigation is required to determine the effects of various levels of energy supplementation on dams' winter weight losses and growth performance of the suckling calf.

Effect of Cow Average Daily Gain from Calving to Weaning

The least squares means of cow ADG from calving to weaning in October (ADGCO) in 1976 and 1977 were respectively 0.6 and 0.4 Kg/day for all age and breed groups examined (Table III.9). Cows gained significantly (P<0.01) less weight per day in 1977, possibly the result of differences in rainfall distribution between years. The SY,

Regression of weaning weight on individual cow and calf variables, Kinsella 1976-1977. Table III.8.

Variable of Interest	Equation	Order of Forced Variables*	Regression Coefficient for Variable of Interest b(Kg) SEb	efficient of Interest SEb	Total Variance Explained Rt x 100 (%)	Additional Variance Explained by Variable of Interest after forcing other Variables (%):
1976 Cow WWLS 1	6 (G)	COW WWLS AC, S, COW WWLS A, B, AB, AC, COW WWLS	0000	000		3.1
COW ADGCD!	(C) (D)	COW ADGCO AC, S, COW ADGCO A, B, AB, AC, S,	2-0-1 0-4-1 3-1-1	+ + 0 - 0 89.	33 2 35 35 36 36 36 36	- 0 0 - 0 0
Calf BW:	(a) (a) (b) (c) (c) (c) (c) (c) (c) (c) (c) (c) (c	Calf BW AC, S, Calf BW A, B, AB, AC, S, Calf BW	44-	400	16.2 55.9 73.3	20.7 3.1
II. 1977 COW WWLS!	4. 8 9 0	COW WWLS AC. S. COW WWLS A. B. AB. AC. S. COW WWLS	000	000	2.1 65.2 2.0	0.2 1.2
COW ADGCO?	9 0 0 0	COW ADGÇO AC, S, COW ADGCO A, B, AB, AC, S,	O	0.0 0.0 7.0	65.7.2 65.5.30	80 Fr
Calf BW.	6. a) b) c)	Calf BW AC, S, Calf BW A, B, AB, AC, S,	9.8 9.9 9.0	e e e	27.7 57.7 71.8	30.8

** AC = Calf Age; S = Calf Sex; A = Age of dam; B = Breed of dam; A × B = Age × Breed of dam interaction.

, Kg

Ac and S accounted for 35.2% and 27.2% of total variance in 1976 and 1977 respectively; A, B, A \times B

Table III.9. Least squeres means and standard errors of cow average daily gain from calving to October. Kinsella 1876-1877.

						.	
	Breed of Dam	M	Number 6 1977	Mean 1976	. S. E.	Mean 1977	S.E.
Grand Total	114	208	242	0.0	0.0	0.4	0.0
	5	;	i				
	: >	P 5	, 0	•	0	0 . Se	0
	8	9 6	3 6		0	0. 1 0	0.0
	S	93	3 2		- c	8	- 0
Age in Years			;	5	9	9.	0.0
~	A11	27.	. 28	0. Bv	0.0	. 3e	0.0
ų	¥	=	ā		6	•	
ę	λ	8	in In	. sr) c	- •	0 0
•	8	<u>.</u>	20	•	0		o •
	S	1 3	-	4.0	0.1	0	-
	A11	36	93	0. *	0.0	0.4x	0.0
	¥	1	1	●.0	1 0		` .
	≿ ક	Ē.	8	0.7	0	•	- 0
	š		•	.	-	•	-
	3		•	•	-	4.0	- 0
-	A1.1	8	38	0.7y	0	0.5x	0.0
	星	s n	_	0.	-	er C	•
	5 8	ō .	=	7.0	-		
	8 2	n (ri (6	- -	0.0	-
	ŝ	,	N	• •	- .	0.0	0.
la ture	A11	7.5		0.7y	0 . t	0.4x	0.0
	¥;	77	52	9 .0	0.0	0	0
	5 8	₹ .	7	•	0.0	•	
	8 2	-:	₹ (<u>-</u> (0.3	4.0	- 0
	3	.,	•	2	•		

Cow AbdCD - cow average daily gain from calving to weaning in October: 8.b.c - Lesst squares sears of Cow AbdCD by breed of daw within years with different siphebatic letters are significantly discessed to the contract of the search of the contract of th

W.x.y " Least squares means of Cow ADOCO by age of dam within years with different alphabetic letters are significantly different (P<0.01)

DB and DS groups demonstrated a tendency to gain less than HE dams. Deutscher and Whiteman (1971) and Wyatt et al., (1977a) noted that as a larger proportion of feed is converted to milk and not to body fat by the high milk-producing crossbreds, these cows did not regain winter weight loss as noticeably as purebred Angus and Hereford cows.

There was a tendency for cows that exhibited more weight loss during pregnancy to gain more per day during lactation between age and breeds, similar to trends observed by Wyatt et al., (1977b).

Phenotypic correlations (Table III.4) of cow ADGCO with milk and weaning weight were low and negative at respectively -0.07 and -0.16 in 1976 and respectively -0.10 and -0.10 and -0.20 in 1977, similar to values reported by Hohenboken et al., (1973) and Koch (1972) (Table III.4). These results suggested that cow weight gain during lactation may be slightly at the expense of milk production. Morris and Wilton (1976) noted that although the magnitude depended on the particular stage of lactation considered, the relationship between milk production and body weight change during lactation was negative as was the relationship between calf weaning weight and weight change of the lactating cow.

Ordered regressions of weaning weight on cow ADGCO (Table III.8 Equations 2a, 5a) indicated that as an individual variable, cow ADGCO accounted for only 2.5 and

4.2% of the total variation explained in weaning weight in 1976 and 1977 respectively, similar to the 1% of variation explained by cow weight and condition as noted by Marshall et al., (1976). After the introduction of the effects of S and CA, and subsequently S, CA, A, B and A x B, (Table III.8, Equations 2b, 2c, 5b, 5c) the percent of additional variation explained over and above these effects remained small for both years.

For every 0.1 Kg increase in cow ADG during lactation, calf weaning weight showed a decrease of 1.5 and 1.0 Kg in 1976 and 1977 respectively (Table III.8 Equations 2c and 5c) after other effects had been forced to enter the regression first. Singh et al., (1970) calculated that a 1% loss in cow weight during suckling was associated with an increase of from 0.14 to 1.09 Kg in weaning weight, suggesting that cows producing more milk had faster gaining calves and lost weight while nursing them.

Morris and Wilton (1976) suggested that heavier weaning weights were derived from cows which lost more body weight during lactation. In the present study however, cow ADGCO accounted for little of either total or additional variation explained in weaning weight. This factor may indicate that cow summer weight changes are confounded with the effects of age and breed of dam which accounted for a large proportion of the variation in weaning weight.

Effect of Calf Birthweight

Least squares means of calf birthweight for 1976 and 1977 were 35.4±0.4 and 38.0±0.4 Kg respectively overall, ranging from 29.9±1.2 Kg for calves of 2-year old HE dams to 44.2±1.2 Kg for calves of mature DS dams. The difference between years was significant (P<0.01) with calves weighing approximately 2 Kg more at birth in 1977 than in 1976. Although calves of older dams and those with dairy breeding tended to be heavier at birth, differences were consistently significant (P<0.05) only between calves of 1-year olds and those born to dams of other age groups. Fahmy and Lalande (1973) reported significantly lighter birth and weaning weights for calves from 2- and 3-year old dams and maximum weights at birth and weaning from cows averaging 7.6 and 8.2 years respectively. Similar trends were noted by Singh et al., (1970), Lawson (1976) and Smith et al., (1976).

Phenotypic correlations of birthweight with weaning weight calculated from the present data were 0.40 and 0.53 in 1976 and 1977 respectively (Table within the range of estimates reported by Rutledge et al., (1971) and Fahmy and Lalande (1973).

Calf birthweight exhibited little association with weaning weight in the regression analysis of adjusted data (Table III.8, Equations 3 and 6) as most of the variation explained was accounted for by the effects of calf sex and age of dam. Calf birthweight accounted for an additional 3.1 and 6.7% of additional variation over the effects of S, CA,

Table III.10. Least squares means and standard Kinsella 1976-1977.

·	Breed of Dam	1976	Number 5 1977	Mean 1876	. S. F.	Mean 1977	. S. F.
Grand Total	A13	206	242	35.40	0.4	36.0••	0
	뿔	4	10	33.8	-	37 80	•
	λS	102	123		, in		
	2	3 8	33	34.28	-	8	
Age in Years	So	33	8 8	8. 8.	• .	40.86	0.
~	114	7.5	83	31.44	8 9.0	34.54	0.1
	¥	Ξ	ā	20 0			•
	λ	36	8	3			۰ ر - ر
٠	2	ā	2	30	0	37.5	; -
-	SO	Ē	••	33.4	<u>-</u>	34.8	. 0.
e.	114	.96	62	34.3×	0.1	8	0.7
	¥	7	1	32.5	10	, B. C.C.	•
	3	5	8	34.1	0.	37.2	0
	8	-	•	33.8	7.		7
	So	•	•	36.0	- 1	43.2	5.6
4	Ę	2	58	37.8y	1.1	39.2×	-:
	¥	10	1	26.2	-		•
	λ	ō	=	8		30	-
	8	7	0	37.4	3.E	8	-
	SO .	0	~	38.3	. 3 . 3	43.2	2.9
Mature		7.8	93	38.27	<u>-</u>	40.4x	1.0
	뿔	33	25	37.1	0.0	37.7	.0
	<u>ک</u>	7	7	8 .	•.0	40.7	
٠,	2	- ;	4	34.8	4 .0	42.0	6.
	02	=	a	. 77	•	•	

'BW = calf birthweight a.b - Least squares means of BW by breed of dam within years with different alphabetic letters are significantly different (P<0.01)

** significantly different at P<0.01

A, B and A x B in each year respectively.

After adjusting for other effects a 1 Kg increase in weight at birth effected an increase of 1.5 Kg at weaning in 1976 and 1.9 Kg in 1977, similar to results found by Lawson (1976), Rutledge et al., (1971) and Singh et al., (1970).

Although higher birthweights may be associated with increased preweaning growth response, this advantage is possibly outweighed by the higher incidence of dystocia and reproductive problems. Berg et al., (1978) noted that a reduction in birth weight would be as undesirable as an excessive increase, because increased mortality is associated with both very small and large calves. Rather, concurrent selection for dams with greater pelvic capacity and for fast-gaining bulls who sire calves with moderate birthweights would be recommended for improving weaning weight.

D. CONCLUSIONS

The effects of dam breed and age were highly significant (P<0.01) and together accounted for between 45 and 48% of the variation in calf weaning weight each year. Cows with dairy background produced more milk and weaned heavier calves than the Hereford and Beef-Synthetic dams. The greatest difference of age of dam on calf weaning weight in this study was between 2-year olds and older dams.

Age and sex of calf were also significant effects. Age of calf accounted for between 20 and 26% of the variation in

calf weaning weight. Regression of weaning weight on calf age was 1.3 Kg/day. There was approximately 6% difference in weight at weaning at approximately 157 days of age between sexes each year.

Measurements of milk yield or associated constituent yields served as good predictors of calf preweaning growth response. Much of the influence of milk was confounded with age and breed of dam. Regression analyses showed that milk yield variables explained a significant 6 to 10% of variation in weaning weight after removing the effects of cow age and breed, and calf age and sex. While measurements of milk constituent yields are useful in determining specific differences between breeds, they accounted for little of the variation in weaning weight over milk yield alone.

The effects of other cow and calf variables were generally small and not significant. Cow winter weight loss during gestation and summer weight gain of the lactating cow had little influence on calf weaning weight. Calf birthweight was moderately correlated with weaning weight. However, as this factor is often associated with high levels of dystocia, selection for increased birthweight is not advisable for increasing weaning weight.

Results indicate that milk and constituent yields are largely influenced by the age and breed of dam. The confounding effect of the breed of dam may be partly due to different inherent growth rate potential between breeds, but

this factor requires further investigation.

As the association between milk yield and weaning weight is so significant, the introduction of dairy breeding into the dam line to increase milk yields and subsequent weaning weights may be a viable consideration for a breeding program in a cow-calf operation.

GENERAL SUMMARY AND CONCLUSIONS

The objectives of the present study were to examine differences in lactation trends among four breed groups including purebred Hereford (beef); a synthetic of three beef breeds; a synthetic of dairy and beef breeds; and dairy x beef crossbred cows; and to determine various cow and calf factors influencing lactation performance of range cows and weaning weights of their calves,

An experiment using exogenous oxytocin was conducted to collect milk samples from a total of 448 cows over two years.

Milk extraction using 20 I.U. of oxytocin and teat tubes appeared adequate as a method of measuring a cow's milk yield. Residual milk collected from cows tested was negligible and generally, yields were similar to those reported in the literature.

Samples were analyzed for constituent percentages and cow and calf weights and ages were recorded.

Few studies comparing lactation trends had been conducted on large samples of range beef cattle, and comparison of results was therefore difficult as methods of milk extraction, sample size, breed differences and statistical analyses varied considerably among research reports reviewed.

A more precise indication of specific lactation trends and effects of age and breed of dam would have been possible

if cows were separated into breed groups according to exact breed compostion. If a larger number of older cows were available, cow age effects could have been partitioned into 9 age groups for cows ranging in age from 2 to 10 years. Had milk measurements been taken more often during the lactation, the lactation curve would provide more information on changing trends.

In the present study however, the author was concerned with general rather than specific trends and associations; conclusions drawn from this study can be practical and applicable in different environments.

Given the wide range in calf age and unequal numbers of breed of dam and age of dam groups, one would expect a high error term. However, error was considerably reduced in the statistical treatment of data by the method of adjustment for main effects including dam age and breed and calf age and sex.

Lactation Trends

Dams with backgrounds which included Holstein and Brown Swiss breeding yielded more milk and exhibited greater persistency than beef crossbreds and purebred Herefords. Because dairy crossbreds in this experiment were not milked regularly for commercial purposes, they produced considerably less milk than dairy cows used in industry. Commercial dairy cows have been selected for high milk production and the introduction of such breeds into the breeding system at Kinsella has served to increase milk

yields among the crossbreds.

However, this factor alone is not the sole reason for increased production. Heterosis may have been a factor that influenced milk production as crossbreds at Kinsella with various beef breed backgrounds yielded more milk than the purebred Hereford dams studied.

Heterosis possibly accounted for the high constituent percentages of butterfat, protein and lactose produced by all the crossbreds. These results paralleled those reported in the literature for crossbred beef cattle, but were higher than for purebred beef cows or traditional dairy cows.

Factors Influencing Milk and Constituent Yields

Regressions on milk and constituent yields indicated that breed and age of dam effects exerted a strong influence on the level of milk trait measured, accounting for up to 44% of the variation in milk yields. Calf preweaning average daily gain (ADG) also exhibited a high association with any one milk yield variable, being associated with 24% of the total variation in yields, ignoring the effects of age and breed of dam.

Factors such as calf birthweight and cow weight changes during pregnancy and lactation did not account for a significant proportion of the variation in any one milk yield variable tested. In the regressions, main effects removed a large proportion of the variation associated with calf birthweight and cow weight changes. Assuming identical management treatment and level of nutrition, these factors

are largely a reflection of calf sex and dam age and breed.

Factors Influencing Weaning Weights

As calf preweating average daily gain is a function of age and weight and was highly associated with milk yields, it was expected that milk and constituent yields would demonstrate a high association with weating weight corrected for calf age.

As milk yields are influenced by the age and breed of dam, regressions on weaning weight indicated that age and breed of dam accounted for up to 48% of the variance in weight of calf weaned.

Mature cows and dairy crossbreds produced the heaviest calves at weaning. However, beef crossbreds which generally yielded more milk than the Herefords, weaned heavier calves than the Hereford dams, again demonstrating the positive effects of crossbreeding, as well as breed composition.

When the effects of age and breed of dam were accounted for, regressions of weaning weight on milk yield showed that milk accounted for up to 10% of the additional variation in weaning weight (P<0.05).

Calf age and sex had moderate effects on weaning weight. Age at weaning however cannot be drastically increased through management without affecting subsequent reproductive performance of the dam.

Calf birthweight and cow weight changes did not significantly influence weaning weights. These effects however, were confounded by age and breed of dam.

Of interest to the producer are factors that can be used to increase herd productivity. Results of this investigation support evidence suggesting that the systematic addition of dairy cattle breeds in a crossbreeding program will positively affect the weight of the weaned calves by increasing milk production of their dams. Combined with selection that emphasizes reproductive performance and high weaning weight, the level of herd productivity can be expected to respond favorably.

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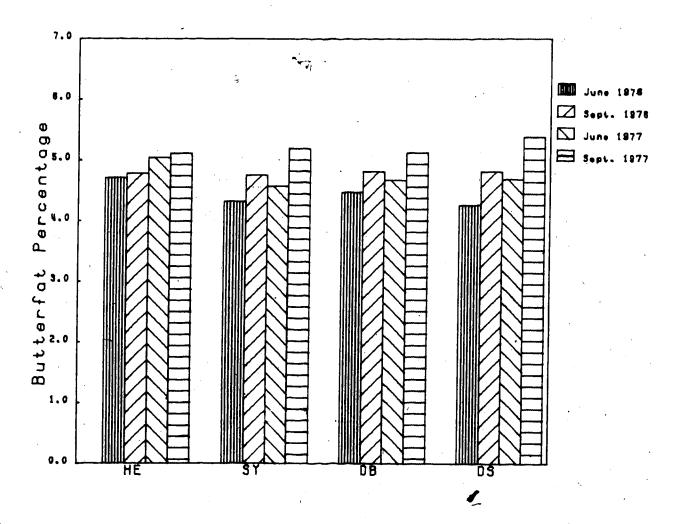
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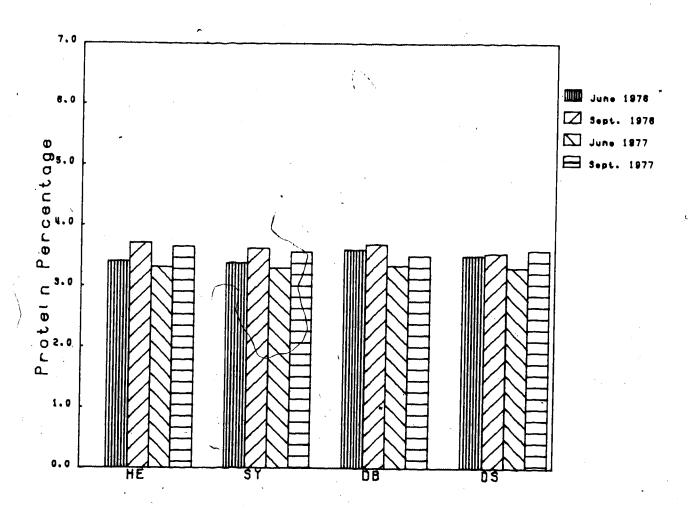
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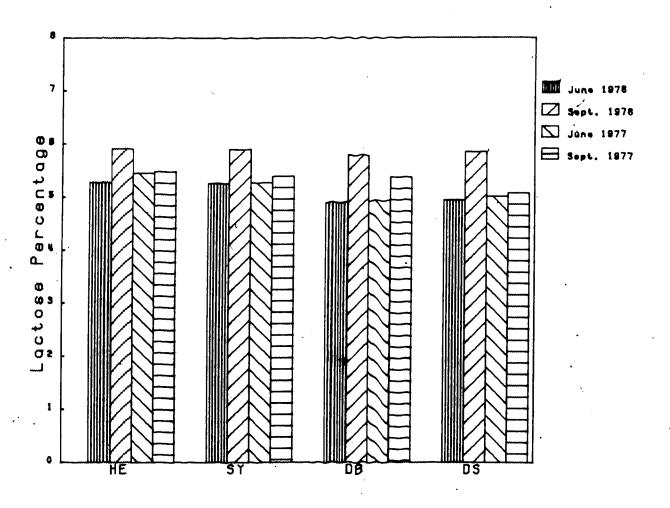
APPENDI CES



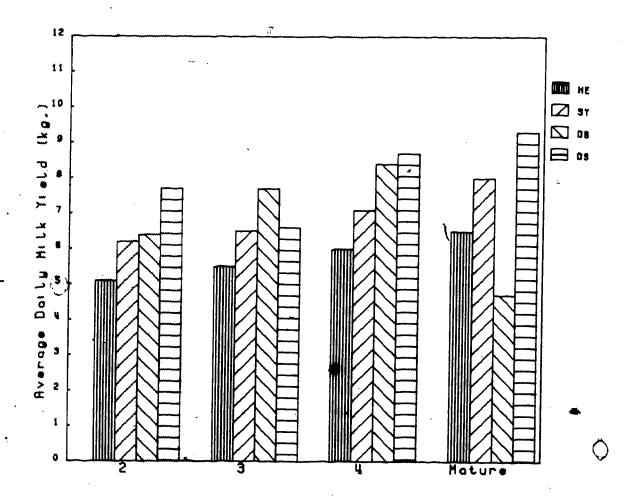
Appendix 1. June and September butterfat percent averages by breed of dam, Kinsella 1976-1977



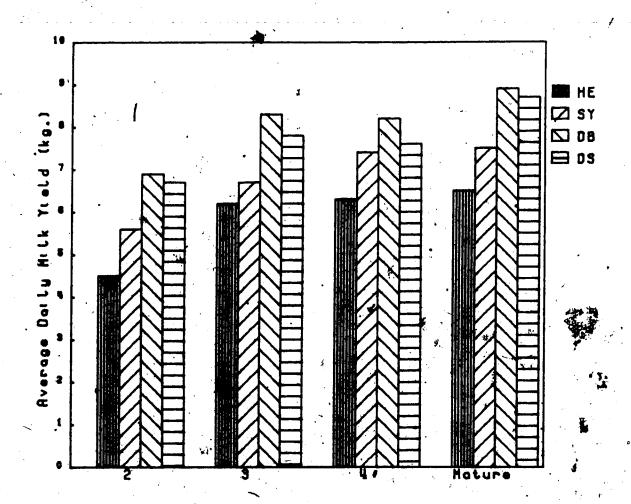
Appendix 2. June and September protein percent averages by breed of dam, Kinsella 1976-1977



Appendix 3. June and September lactose percent averages by breed of dam, Kinsella 1976-1977



Appendix 4. Average milk yield by breed and age of dam, Kinsella 1976



Appendix 5. Average milk yield by breed and age of dam, Kinsella 1977

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